# Stochastic Analysis of Particle Movement over a Dune Bed

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By BAUM K. LEE and HARVEY E. JOBSON

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## METRIC-ENGLISH EQUIVALENTS

Metric unit	English equivalent			
Meter (m)	= 3.28 feet			
Centimeter (cm)	$= 3.28 \times 10^{-2} $ foot			
Tonne per day per meter (t/day · m)	= .336 ton per day per foot			
Dyne per square centimeter (d/cm <sup>2</sup> )	= $2.09 \times 10^{-3}$ pound per square foot			
Dyne per centimeter per second (d/cm·s)	$=6.85 \times 10^{-5}$ pound per foot per second			

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### **SYMBOLS**

A,B	Constants in estimating the conditional mean of the rest periods	$egin{array}{c} p(ullet) \ P[ullet] \end{array}$	Sample probability mass function Probability
C,D	Constants in estimating the conditional variance of the rest periods	$\hat{q}_B'(j)$	Estimate of the mean bed-load discharge per unit width associated with the elevation $y_i$
$C\sqrt{g}$	Dimensionless Chezy discharge coefficient	$\hat{q}_T$	Estimate of the mean total bed-material discharge
$C_T$	Concentration of total bed-material discharge		per unit width
d	Depth of flow	$[\widehat{Q}_{B},\widehat{Q}_{B}]$	Estimates of the mean bed-load discharge
$d_{50}$	Median sieve diameter of bed material	$\hat{Q}_S, \hat{Q}_S'$	Estimates of the mean suspended-load discharge
d <sub>84</sub>	Sieve diameter of bed material for which 84 percent is smaller	$egin{array}{c} \widehat{Q}_{S}^{'}, \widehat{Q}_{S}^{'} \ \widehat{Q}_{T}^{'}, \widehat{Q}_{T}^{'}, \widehat{Q}_{T}^{''} \end{array}$	Estimates of the mean total bed-material discharge Water discharge
d <sub>16</sub>	Sieve diameter of bed material for which 16 percent is smaller	r	Number of class intervals for the realization of $T$ Shape parameters for the conditional distribution of
$oldsymbol{E_{i,j,v}}$	Event that a particle eroded from elevation $y_i$ passes $v$ dune crests before it is deposited at elevation $y_i$	$r_{1,y,y'}, r_{2,y}$	the step lengths and the rest periods, respectively
$E_v$	Event that a particle passes $v$ dune crests before it is	<i>s</i>	Number of class intervals for the realization of X
	deposited	s <sub>y</sub>	Standard deviation of bed elevation
$\mathbf{E}[\cdot],  \hat{\mathbf{E}}[\cdot]$	Mathematical expectation and its estimate, respec-	$S_e$	Slope of energy-grade line
	tively	t	Measure of time Statistic which measures the conditional rest
$\mathbf{F}_r$	Froude number of the flow	$t_{j,oldsymbol{k}}$	periods
$f(\bullet), F(\bullet)$	Probability density and distribution functions,	<b>,</b>	Class mark for the statistic, $t_{j,k}$
(n)	respectively	$egin{array}{c} t_lpha \ T \end{array}$	Random variable describing the rest periods of a
f(•)	<i>n</i> -fold convolution of the probability density, $f(\cdot)$	1	particle
g L	Gravitation acceleration	$T_i$	Random variable describing the duration of ith rest
h	Average depth of the zone in which bed-material movement occurs determined from the $y_t(x)$	-1	period
ь.	record  Scale parameter of the conditional step length dis-	T(n)	Stochastic process describing the sum of n rest periods
$k_{1,y,y'}$	tribution	$\overline{oldsymbol{U}}$	Mean flow velocity
$k_{2,y}$	Scale parameter of the conditional rest period dis-	$\overline{U}_{f *}$	Mean shear velocity
2,3	tribution	$\hat{V}_B, \hat{V}_S$	Estimates of the mean transport speed of a had lead
$L_t$	Length of the $y_x(t)$ record	V <sub>B</sub> , V <sub>S</sub>	Estimates of the mean transport speed of a bed-load particle and a suspended-load particle, respec-
$L_{x}$	Length of the $y_t(x)$ record		tively
$m_{i,j}$	Total number of bed forms contained in the $y_t(x)$	$\hat{V}_{B}(j)$	Estimate of the mean transport speed of a bed-load
	record for which the upstream side intersects the		particle associated with the elevation $y_j$
	elevation $y_i$ and the downstream side intersects the elevation $y_i$	$\hat{\mathbf{v}}_{_{m{T}}},\hat{\mathbf{v}}_{_{m{T}}}$	Estimates of the mean transport speed of a bed-
$m_{i,j,v}$	Total number of possibilities of the event $E_{i,j,v}$ con-	^	material particle
	tained in the $y_t(x)$ record	$\hat{V}_{T^{(j)}}$	Estimate of the mean transport speed of a bed- material particle associated with the elevation $y_i$
$m_i'$	Total number of bed forms contained in the $y_x(t)$	Var[•],Var[•]	Variance and its estimate, respectively
	record and which also contain some erosion in the class interval associated with the elevation $y_i$	W	Width of channel
$m_j$	Total number of bed forms contained in the $y_x(t)$	$\boldsymbol{x}$	Measure of longitudinal distance
,	record and which also contain some deposition in	$\overline{x}$	Average distance traveled by bed material in time $t$
	the class interval associated with the elevation $y_i$	$x_{oldsymbol{eta}}$	Class mark for the conditional step lengths
$m_{j,j}$	Total number of bed forms contained in the $y_x(t)$	$x_{i,j,k}$	Statistic which measures the conditional step
	record and which also contain both an up-crossing		lengths of a bed-load particle
	and a down-crossing at the elevation $y_j$	$x_{i,j,v,k}$	Statistic which measures the conditional step
n	Number of steps taken by a particle or number of		lengths of a bed-material particle
3.7	class intervals for the realization of $Y_D$ and $Y_E$	$\boldsymbol{X}$	Random variable describing the step lengths of a
$N_d$	Total number of particles per unit area deposited in time t		particle
N	Total number of particles per unit area eroded in	$X_i$	Random variable describing the length of ith step of
$N_e$	time t		a particle
$N_d(y_i)$	Total number of particles per unit area deposited	X(n)	Stochastic process describing the longitudinal posi-
u ~ i	within the class interval $(n_i, n_{i+1}]$ in time $t$	~	tion of a particle after n steps
$N_e(y_i)$	Total number of particles per unit area eroded with-	$\widetilde{X}(t)$	Stochastic process describing the longitudinal posi-
	in the class interval $(n_i, n_{i+1}]$ in time $t$		tion of a particle at time t
N(t)	Counting process describing number of steps taken	y, y'	Measure of vertical distance
	by a particle in time $t$	$y_i$ , $y_j$	Class marks for $Y_E$ and $Y_D$

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$y_{\text{max}}$ , $y_{\text{min}}$	Highest and lowest elevations, respectively, at which particles are deposited or eroded	$\Delta y_{j,k}^-$	Vertical fall of the bed in the class interval associated with $y_i$ for the kth erosion period of the $y_x(t)$
$\widehat{m{y}}_{ ext{max}}, \widehat{m{y}}_{ ext{min}}$	Estimates of $y_{\text{max}}$ and $y_{\text{min}}$ , respectively		record
$y_t(x)$	Elevation of the bed, y, as a function of the	$  \gamma_s  $	Specific weight of bed material
	longitudinal coordinates, $x$ , at a given time, $t$	Γ(•)	Gamma function
$y_x(t)$	Elevation of the bed, y, as a function of time, t, at a	$\zeta_j, \xi_j$	Effective volume ratios
	fixed point, $x$	$\eta_j, \eta_{j+1}$	Lower and upper class limits for $y_j$ , respectively
$Y_D, Y_E$	Random variables describing the elevation of parti-	$\theta$	Bulk porosity of the bed material in place
	cle deposition and erosion, respectively	$\lambda_{\beta}$ , $\lambda_{\beta+1}$	Lower and upper class limits for $x_B$ , respectively
$Y_D(n)$	Stochastic process describing the vertical position of	$\hat{\rho}$	Estimate of correlation coefficient
	a particle after $n$ steps	$\sigma_{a}$	Geometric standard deviation of particle size
$\widetilde{\mathbf{y}}(t)$	Stochastic process describing the vertical position of	$\left  \begin{array}{c} \sigma_{\mathbf{g}} \\ \overline{ au}_{b} \end{array} \right $	Mean bed shear stress
	a particle at time t	$\tau_{\alpha}, \tau_{\alpha+1}$	Lower and upper class limits for $t_{\alpha}$ , respectively
$\Delta y_j$	Class width associated with $y_i$	$\chi_c$	Critical value of chi-square statistic
$\Delta y_{j,k}^+$	Vertical rise of the bed in the class interval associ-	Ω	Number of particles per unit volume of the bed
· <b>J</b> 5**	ated with $y_j$ for the $k$ th deposition period of the $y_x(t)$ record	$\Omega_j$	Number of particles per unit volume of the bed associated with $\boldsymbol{y}_j$

# STOCHASTIC ANALYSIS OF PARTICLE MOVEMENT OVER A DUNE BED

By BAUM K. LEE and HARVEY E. JOBSON

#### **ABSTRACT**

Stochastic models are available that can be used to predict the transport and dispersion of bed-material sediment particles in an alluvial channel. These models are based on the proposition that the movement of a single bed-material sediment particle consists of a series of steps of random length separated by rest periods of random duration and, therefore, application of the models requires a knowledge of the probability distributions of the step lengths, the rest periods, the elevation of particle deposition, and the elevation of particle erosion. In the past, it has proven impossible to estimate these distributions except by use of tedious and time-consuming single particle experiments.

By considering a dune bed configuration which is composed of uniformly sized particles, the probability distributions of the rest period, the elevation of particle deposition, and the elevation of particle erosion are obtained from a record of the bed elevation at a fixed point as a continuous function of time. By restricting attention to a coarse sand, where the suspended load is negligible, the probability distribution of the step length is obtained from a series of "instantaneous" longitudinal bed profiles in addition to the above information. Using these probability distributions, three bed-material transport equations and a two-dimensional stochastic model for dispersion of bed-sediment particles are developed.

The procedure was tested by determining these distributions from bed profiles formed in a large laboratory flume with a coarse sand as the bed material. The elevation of particle deposition and the elevation of particle erosion can be considered to be identically distributed, and their distribution can be described by either a "truncated Gaussian" or a "triangular" density function. The conditional probability distribution of the rest period given the elevation of particle deposition. The conditional probability distribution of the step length given the elevation of particle erosion and the elevation of particle deposition also closely followed the two-parameter gamma density function. For a given flow, the scale and shape parameters describing the gamma probability distributions can be expressed as functions of bed elevation.

The bed-material transport equations were tested for three flow conditions. The errors in the predicted mean total bed-material transport rates were -3.0, +3.5, and 80.1 percent for equation 55, and -1.7, +26.9, and +64.1 percent for equation 63. For the run with the large error, the mean total load concentration was small (8.9 milligrams per liter), and flow conditions were somewhat out of equilibrium.

#### INTRODUCTION

The movement of sediment in alluvial streams is so complex a process that it may never be subjected completely to a deterministic solution. It represents, in fact, an extreme degree of unsteady, nonuniform flow, since the streambed as well as the water surface may be continuously changing with time and position.

Numerous formulas and equations have been developed to predict sediment transport rates. Most of these developments ignore the actual nature of sediment movement and have assumed that the sediment transport rate can be described by a deterministic function of certain flow parameters. Unfortunately, after decades of searching, no universally accepted sediment transport equation has been found. The theories of probability, statistics, and stochastic processes have been used to describe the kinematics of a single bed-sediment particle in an alluvial channel flow and to predict the dispersion characteristics of a group of such particles. These theories have clearly demonstrated a great potential for development of stochastic models of sediment transport and dispersion.

Most of the stochastic models (Shen and Todorovic, 1971; Grigg, 1969; Yang, 1968; Sayre and Conover, 1967; Hubbell and Sayre, 1964; Crickmore and Lean, 1962; Einstein, 1937) are based on the proposition that the movement of bed-sediment particles consists of a series of steps separated by rest periods, so that determination of the probability distributions for the step lengths and the rest periods of a bed-sediment particle plays the major role in quantifying the bed-sediment transport. While this movement concept can easily be verified through laboratory observations, Einstein (1937) was the first to use it. He developed a one-dimensional probabilistic model for bedload transport. More recently, Sayre and Conover (1967) derived a two-dimensional stochastic model by introducing the prob-

ability distribution of the elevation at which a bed-sediment particle is deposited.

The probability distributions of the step lengths and the rest periods of a bed-sediment particle have been estimated from single particle experiments (Grigg, 1969) or by using a group of tracer particles (Yang, 1968; Hubbell and Sayre, 1964; Crickmore and Lean, 1962). Because of the considerable effort required to conduct such experiments, it seems clear that some way must be found to estimate the probability distributions from more readily accessible data if significant further progress is to be expected. To apply the Sayre-Conover (1967) two-dimensional stochastic model, the probability distribution of the elevation at which a bed-sediment particle is deposited must be known. A method for estimating this distribution is developed in this report.

The objectives of this study are:

- 1. To present a method of estimating the following probability distributions for dune-bed conditions using only sounding records of the bed elevation (a) probability distributions (note that there are two separate distributions) of the elevation at which a bed-sediment particle is eroded and deposited, and (b) conditional probability distributions of the step lengths of a bed-sediment particle given the elevations at which the particle is eroded and deposited. A method estimating the conditional probability distribution of the rest periods of a bed-sediment particle given the elevation at which the particle is deposited has been presented by Sayre and Conover (1967).
- 2. To develop bed-material transport equations based on the above probability distributions and to compare the results with the experimentally measured values.
- 3. To derive a two-dimensional stochastic model for dispersion of bed-sediment particles as a function of the above probability distributions.

The probability distributions of the elevation at which a bed-sediment particle is eroded and deposited, and the probability distribution of the rest periods, conditioned on the elevation of deposition, will be obtained from a continuous record of the bed elevation at a particular point as a function of time. The probability distribution of the step lengths, conditioned on the elevation of erosion and the elevation of deposition, will be obtained from a series of "instantaneous" longitudinal bed profiles. With these distributions obtained, various related probability distributions of vital interest will be estimated, and a relation between the rest periods and the step lengths of a bed-sediment particle will be investigated.

Three experimental runs are analyzed and the relations between the statistics describing the postulated probability distributions and the hydraulic conditions are investigated. All data were obtained from a tilting

recirculating flume of rectangular cross section 61 m long, 2.4 m wide, and 1.2 m deep. The bed material used in these experiments was screened river sand with a median sieve diameter equal to 1.13 mm and a geometric standard deviation equal to 1.51.

#### **ACKNOWLEDGMENT**

The data contained herein are essentially the same as those contained in a dissertation by Lee (1973). The data were collected under the general supervision of the second author. Special thanks are due E. V. Richardson, D. B. Simons, C. F. Nordin, D. C. Boes, and R. P. Osborne.

#### **BACKGROUND**

#### THEORETICAL MODELS

Einstein (1937) treated the movement of a single sediment particle over an alluvial bed as a stochastic process described by an alternating sequence of two independent random variables, namely, step lengths and rest periods. Considering the particle movement in the distance-time plane on a Galton's board (Parzen, 1960), Einstein derived exponential probability density functions for the step lengths and the rest periods,

$$f_X(x) = k_1 e^{-k_1 x}$$
,  $x > 0$  (1)

and

$$f_T(t) = k_2 e^{-k_2 t}, t > 0$$
 (2)

respectively, where

X,T = random variables describing the step lengths and rest periods of a particle, respectively;

x,t =distance and time, respectively;

 $f_X(x), f_T(t) = \text{common probability density functions}$  of the step lengths and rest periods, respectively; and

 $k_1, k_2 = \text{positive constants}.$ 

For a sediment particle introduced into the stream at distance x = 0 in such a way that it takes its first step at time t = 0, Einstein obtained the probability density function of the total distance traveled by the particle at time t to be

$$f(x;t) = k_1 e^{-k_1 x - k_2 t} \sum_{n=1}^{\infty} \frac{(k_1 x)^{n-1}}{\Gamma(n)} \frac{(k_2 t)^{n-1}}{\Gamma(n)},$$

$$x > 0, t > 0, (3)$$

in which  $\Gamma(\cdot)$  denotes the gamma function. Equation 3 also represents the concentration distribution of a group of identical sediment particles with respect to

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longitudinal position, x, as a function of time, t.

The probability density function for the case when the particle is initially (t = 0) at rest at x = 0 also was obtained by a similar procedure.

$$f(x;t) = k_1 e^{-k_1 x - k_2 t} \sum_{n=1}^{\infty} \frac{(k_1 x)^{n-1}}{\Gamma(n)} \frac{(k_2 t)^n}{\Gamma(n+1)}$$

$$x > 0, t > 0$$
 . (4)

It should be noted that equation 4 applies only to the particle that has taken at least one step.

Einstein (1950) also developed his well-known bedload equation by considering the dynamic lift force as a random variable. The idea is that the probability of a sediment particle being eroded from the bed surface is equal to the probability that the lift force exerted on the particle exceeds its submerged weight. He obtained

$$p = 1 - \frac{1}{\sqrt{\pi}} \int_{-B_{*}\Psi_{*}}^{B_{*}\Psi_{*}} - \frac{1}{\eta_{0}} e^{-z^{2}} dz = \frac{A_{*}\Phi_{*}}{1 + A_{*}\Phi_{*}}, \quad (5)$$

where

p = probability of a sediment particle being eroded;

 $\eta_0, A_*, B_* = \text{constants};$ 

 $\Psi_*$  = intensity of shear for an individual particle

 $\Phi_*$  = intensity of transport for an individual particle size.

Solving equation 5 for  $\Phi_*$ , which is a function of the bedload transport rate, one obtains the bed-load discharge for individual particle sizes from hydraulic parameters and sediment properties.

Hubbell and Sayre (1964) presented a one-dimensional stochastic model for the longitudinal dispersion of bed-material particles in an alluvial channel. The results are identical to Einstein's (eqs. 1-4). The assumptions are: (1) the flow is in equilibrium (Simons and Richardson, 1966); (2) the particle always moves in a downstream direction with a series of alternate steps and rests; (3) the duration of movement is insignificant compared to the rest periods; and (4) the stochastic processes describing the number of steps taken by a particle in a distance interval and a time interval are independent of each other and both are homogeneous Poisson processes (Parzen, 1967). These assumptions are essentially the same as those of Einstein's (1937) although stated in a different way.

Based on the concept of continuity, Hubbell and Sayre (1964) proposed the transport equation for the bed material of a certain characteristic,

$$(Q_T)_c = i_c (\gamma_s)_c (1 - \theta) Wh\left(\frac{\overline{x}}{t}\right)_c , \qquad (6)$$

where

 $Q_T$  = bed-material discharge in weight per unit time;

 $i_c$  = ratio of the volume of particles possessing the characteristic size to the volume of bedmaterial particles in the zone of particle move-

 $\gamma_s$  = specific weight of the bed material;

 $\theta$  = bulk porosity of the bed in place;

W =width of channel;

h =average depth of the zone in which particle movement occurs:

 $\overline{x}$  = average distance traveled by bed material in time t:

t =measure of time; and

c =subscript that denotes terms associated with particles possessing a certain characteristic

Combining equation 6 with the result from the Hubbell-Savre one-dimensional stochastic model gives the total bed-material discharge for all particle sizes,

$$Q_T = \sum_c i_c (\gamma_s)_c (1 - \theta) Wh \left(\frac{k_2}{k_1}\right)_c , \qquad (7)$$

in which  $k_1$  and  $k_2$  are defined in equations 1 and 2, respectively.

Savre and Conover (1967) extended the one-dimensional stochastic model derived by Hubbell and Sayre (1964) to two dimensions by introducing the vertical level at which particles are deposited. Their analysis led to the joint probability density function for the event that a particle has, at time t, traveled a distance equal to x and is located at an elevation equal to y,

f(x, y, t)

$$= f_{Y_D}(y) \sum_{n=1}^{\infty} f_X^{(n)}(x) \int_0^t f_T^{(n)}(t') \int_{t-t'}^{\infty} f_{T \setminus Y_D}(t \setminus y) dt dt', (8)$$

where

 $f_{Y_D}(y)$  = probability density function for the

elevation of particle deposition;  $f_X(x), f_T(t) = n$ -fold convolutions of  $f_X(x)$  and  $f_T(t)$ , respectively;

 $f_{T \setminus Y_D}(t \setminus y) =$ conditional probability density function for the rest periods given the elevation at which the particle is deposited; and

t' = sum of the first n rest periods.

If a group of identical sediment particles are released simultaneously at x = 0,  $y = y_0$ , and t = 0, equation 8 gives the concentration of the particles, which were initially at rest and have moved from their respective initial positions, with respect to longitudinal position, x, and vertical position, v, as a function of time, t.

In order to apply equation 8, the density functions,  $f_{Y_D}(y)$ ,  $f_{T \setminus Y_D}(t \setminus y)$ , and  $f_X(x)$  must be specified. The unconditional density function,  $f_T(t)$  is related to  $f_{T \setminus Y_D}(t \setminus y)$  and  $f_{Y_D}(y)$  by the relation

$$f_T(t) = \int_{y_{\min}}^{y_{\max}} f_{T \setminus Y_D}(t \setminus y) f_{Y_D}(y) dy, \qquad (9)$$

in which  $y_{\text{max}}$  and  $y_{\text{min}}$  are the highest and lowest elevations at which particles can be deposited, respectively. The marginal case of equation 8 is

$$f(x;t) = \sum_{n=1}^{\infty} f_X^{(n)} \int_0^t \left[ f_T^{(n)} - f_T^{(n+1)} \right] dt$$
$$= \sum_{n=1}^{\infty} f_X^{(n)} P[N(t) = n], (10)$$

in which P[N(t) = n] denotes the probability that a particle takes n steps in a time interval t. Equation 10 is a general one-dimensional stochastic model where only longitudinal dispersion is considered. One may note that the substitution of equations 1 and 2 into equation 10 reduces to equation 4.

Yang (1968) assumed the step lengths are gamma distributed with a shape parameter, r, and the common density function,

$$f_X(x) = \frac{k_1}{\Gamma(r)} (k_1 x)^{r-1} e^{-k_1 x}, x > 0$$
, (11)

and the rest periods are exponentially distributed with the common density function given in equation 2. Substituting equations 2 and 11 into equation 10, he obtained

$$f(x;t) = k_1 e^{-k_1 x - k_2 t} \sum_{n=1}^{\infty} \frac{(k_1 x)^{nr-1}}{\Gamma(nr)} \frac{(k_2 t)^n}{\Gamma(n+1)} ,$$

$$x > 0, t > 0$$
 . (12)

Since the gamma distribution reduces to the exponential distribution when r = 1, equation 4 is actually a special case of equation 12.

Shen and Todorovic (1971) generalized the Hubbell-Sayre one-dimensional model given in equations 1, 2, and 4. The essential difference between the two models is that the former was based on the nonhomogeneous Poisson processes (Parzen, 1967), while the latter was based on the homogeneous Poisson processes. In the Shen-Todorovic model, the probability density functions of the step lengths and the rest periods are, respectively,

$$f_{X}(x) = k_{1}(x) e^{-\int_{x_{0}}^{x} k_{1}(s) ds}, x > 0, (13)$$

and

$$f_{T}(t) = k_{2}(t) e^{-\int_{t_{0}}^{t} k_{2}(s) ds} , t > 0 , \qquad (14)$$

where

 $k_1(x)$ ,  $k_2(t)$  = functions of x and t, respectively; and

 $x_0$ ,  $t_0$  = initial position and time, respectively. The probability density function of the total travel distance of a particle, which was initially at rest and has moved from its initial position,  $x_0$ , was found to be

moved from its initial position, 
$$x_0$$
, was found to be
$$-\int_{x_0}^{x} k_1(s) ds - \int_{t_0}^{t} k_2(s) ds$$

$$f(x;t) = k_1(x) e$$
(15)

$$\sum_{n=1}^{\infty} \frac{\left[ \int_{x_0}^{x} k_1(s) \, ds \right]^{n-1}}{\Gamma(n)} \frac{\left[ \int_{t_0}^{t} k_2(s) \, ds \right]^n}{\Gamma(n+1)}, \ x > 0, \ t > 0.$$

It is seen from equations 13 and 14 that the mean number of steps taken by a particle in  $(x_0, x]$  and  $(t_0, t]$  are

$$\int_{x_0}^{x} k_1(s) ds$$
 and  $\int_{t_0}^{t} k_2(s) ds$ ,

respectively, whereas those of Hubbell-Sayre's model are  $k_1(x-x_0)$  and  $k_2(t-t_0)$ , respectively. The Hubbell-Sayre (1964) one-dimensional model is a special case of the Shen-Todorovic model.

#### **EXPERIMENTAL STUDIES**

Hubbell and Sayre (1964) conducted concentration distribution experiments both in the field and laboratory to evaluate the one-dimensional stochastic model given by equation 4. The bed configurations in these experiments were large dunes in the field and ripples in the laboratory flume. Using radioactive tracer particles, a series of longitudinal concentration-distribution curves were obtained at different times for a given flow condition. The longitudinal concentration-distribution function,  $\Phi(x;t)$ , is defined to be the weight of tracer particles per unit volume of bed material as a function of longitudinal distance and time and is related to f(x;t) by

$$\Phi(x;t) = \frac{W_T}{Wh} f(x;t) , \qquad (16)$$

in which  $W_T$  is the total weight of the tracer particles placed in the channel, W is the channel width, h is

average depth of the zone of bed material movement, and f(x;t) is given by equation 4. Based on equation 16, the parameters  $k_1$  and  $k_2$  were estimated. With these estimates, Hubbell and Sayre reported that the theoretical and observed concentration-distribution functions agree reasonably well.

Yang (1968) carried out a set of experiments using radioactive tracer particles to verify the model given by equation 12. Experiments were performed with ripple and dune bed conditions in a laboratory flume 0.6 m wide by 18.3 m long. He reported that the shape of the experimental longitudinal dispersion curves are fairly well represented by equation 12. Yang also made preliminary runs with a single plastic particle in a small flume and found that the step lengths very closely follow a gamma distribution with the parameter r approximately equal to 2 and that the rest periods follow an exponential distribution very closely.

The first intensive experimental study on the movement of single particles was done by Grigg (1969). The experiments were conducted in a laboratory flume with two bed material sizes. The bed configurations were ripples and dunes. Using single radioactive tracer particles, he measured the step lengths and the rest periods directly and found the step lengths to be approximately gamma distributed and the rest periods to be approximately exponentially distributed as proposed by Yang.

Grigg found interesting correlations between: (1) Various properties of the step length distribution, the stream power (product of mean bed shear stress and mean flow velocity), and the distribution of bedform lengths; and (2) various properties of the rest period distribution and statistical properties derived from the variation of bed elevation with respect to time.

Based on an idea suggested by Hubbell and Sayre (1965), Grigg also made some progress toward experimentally testing the Sayre-Conover two-dimensional stochastic model. By analyzing a record of the bed elevation as a function of time, he showed that the conditional probability density function of the rest periods can be approximated by the exponential function,

$$f_{T \setminus Y_D}(t \setminus y) = k_3(y) e^{-k_3(y)t} , \qquad (17)$$

and

$$k_3(y) = \frac{1}{\alpha} e^{\beta y} , \qquad (18)$$

in which  $\alpha$  and  $\beta$  are constants and  $\gamma$  measures bed elevation in terms of the standard deviation about mean bed elevation.

#### REMARKS

Based on the review given in the previous sections, the following remarks are offered.

- 1. The Sayre-Conover model given by equation 10 is the most general one-dimensional model. The rest of the one-dimensional models, which were previously discussed, can be obtained from this model by proper substitutions. Therefore, it may be rated as the best existing one-dimensional model.
- 2. The Sayre-Conover model given by equation 8 is the only existing two-dimensional stochastic model. The derivation of the Sayre-Conover model has been discussed by Lee (1973). To verify equation 8, a method of estimating the probability distribution of the elevation at which particles are deposited must be known. One of the purposes of this report is to present such a method.
- 3. In order for the stochastic model to serve a prediction purpose, the relation between flow conditions and the parameters describing the probability distributions must be known. Without such knowledge the stochastic models cannot contribute much to the prediction problem.
- 4. A great deal of effort is required to perform dispersion and single particle experiments. If another method can be developed to estimate the necessary probability distributions from more readily accessible data, considerable savings would result. The methods developed in this report require only records of bed elevation.

#### DEVELOPMENT OF THEORY

## CHARACTERISTICS OF PARTICLE MOVEMENT OVER A DUNE BED

Dunes are one of the most common bed forms in alluvial channels. Field observations by Simons and Richardson (1966) indicated that dunes may form in any alluvial channel, irrespective of the size of bed material, if the stream power is sufficiently large to cause general transport of the bed material without exceeding a Froude number of unity. The longitudinal profile of a dune is approximately triangular in shape with a gentle upstream slope and steep downstream slope. The upstream slope depends somewhat on flow conditions, whereas the downstream slope is more dependent on the angle of repose of the bed material. The length of a dune ranges from about 0.61 m to several hundred meters, depending on the scale of the flow system. The Chezy discharge coefficient,  $C/\sqrt{g}$ , ranges from 8 to 12, and the total bed-material discharge concentration ranges from 100 to 1,200 milligrams per liter for dune flow conditions. For further information readers may refer to Simons and Richardson (1966).

For dune flow conditions, a record of the bed elevation as a function of time at a particular location reveals an alternating sequence of periods during which either erosion or deposition is occurring. This type of record is commonly obtained from the output of a depth sounder which is at a fixed location and hereafter will be referred to as the  $y_r(t)$  record, that is, the elevation of the bed, y, positive upward, as a function of time, t, at a fixed location, x. When deposition occurs,  $[dy_r(t)/dt] > 0$ , and when erosion occurs,  $[dy_r(t)/dt] < 0$ , provided these derivatives exist. An instantaneous longitudinal bed profile may be characterized by an alternating series of erosion and deposition reaches. An instantaneous longitudinal profile can be obtained by mounting a depth sounder in a boat, provided that the speed of the boat is large relative to the speed of bed forms. These bed profiles will hereafter be referred to as the  $y_t(x)$  records, that is, the elevation of the bed, y, positive upward, as a function of the longitudinal coordinates, x, at a given time, t. The longitudinal coordinate will be assumed to increase in the downstream direction; therefore, the reaches with positive slopes,  $[dy_t(x)/dx] > 0$ , represent the upstream or stoss sides of the dunes, and the reaches with negative slopes,  $[dy_{i}(x)/dx] < 0$ , represent the downstream or slip faces of the dunes. The dune crest is defined by a local maximum in the  $y_t(x)$  record, and the dune trough is defined by a local minimum in the record.

Anyone who has an opportunity to observe closely the movement of sediment is aware that dunes move downstream owing to erosion from their upstream face and deposition on their downstream face. That is, the bed forms migrate downstream because deposition occurs on the downstream face, where  $[dy_t(x)/dx] < 0$ , and erosion occurs on the upstream face, where  $[dy_t(x)/dx] > 0$ . It will be assumed throughout this report that no deposition occurs on the upstream sides of dunes and no erosion occurs on the downstream faces of dunes. This assumption is not strictly true physically but is necessary for the determination of the conditional step length distributions. If the assumption is true, each sediment particle on the stoss side of a dune must make a step in the downstream direction before it is deposited on the slip face of any dune. Once deposited it rests there until the dune has migrated downstream and it becomes reexposed on the stoss side. In other words, sediment particles are transported downstream in an alternating sequence of steps and rests of random length and duration. The frequencies and magnitudes of these steps and rests are of basic interest in understanding the nature of the movement of the sediments.

Because particles must be eroded from and deposited on the surface of the bed, the step length of a particular particle depends only on the elevation from which it is eroded, the elevation at which it is deposited, the number of dune crests which it passes before being deposited, and the scale and shape of the bed surface  $(y_t(x) \text{ record})$  during the time of its movement. Likewise

the rest period of a particular particle depends on the scale and shape of the  $y_x(t)$  record and on the elevation at which the particle is deposited. If the bed material size is not uniform, the elevation of deposition or erosion may also depend on the size of particles because of vertical sorting.

The intimate relationship between the bed-form shape, as measured by the  $y_r(t)$  and  $y_t(x)$  records, and the step lengths and rest periods of a bed-material particle allow the probability distributions of the step lengths and the rest periods to be estimated from the bed-form data. In the following three sections, a method of estimating the probability distributions of the rest periods, step lengths, and elevations at which a particle is deposited or eroded using the  $y_r(t)$  and  $y_t(x)$  records will be presented. In the last two sections the bedmaterial transport equations and a general two-dimensional bed-material dispersion equation will be derived as functions of these probability distributions. In the next chapter the transport equations will be tested using data from three flume runs and the results will be discussed.

#### ESTIMATION OF THE PROBABILITY DISTRIBUTIONS OF THE ELEVATIONS OF DEPOSITION AND EROSION

The probability that particles are deposited between the elevations  $\eta_i$  and  $\eta_{i+1}$  may be written as

$$\begin{split} &P[\eta_{j} < Y_{D} \leq \eta_{j+1}] \\ &= \lim_{t \to -\infty} \left\{ \begin{aligned} & \text{number of particles deposited} \\ & \underbrace{\begin{aligned} & \text{within the interval } (\eta_{j}, \ \eta_{j+1}] \ \text{in time } t} \\ & \text{number of particles deposited} \\ & \text{over all intervals in time } t \end{aligned} \right\}, (19) \end{split}$$

where

 $P[\bullet] = \text{probability};$ 

 $Y_D$  = random variable describing the elevations at which particles are deposited;

 $\eta_{j}, \eta_{j+1} = \text{lower and upper class limits associated with the class mark of the elevation } y_{j}, \text{ respectively; and}$ 

t = time during which the observations were made.

The elevation at which particles are deposited will hereafter simply be referred to as the elevation of deposition,  $Y_D$ .

If the number of particles per unit volume of the bed,  $\Omega$ , is constant, the flow is stationary (statistical sense), and both erosion and deposition cannot occur at the same point at the same time, the numerator and the denominator of equation 19 can be obtained from the  $y_x(t)$  record. The total number of particles deposited per unit area within the class interval  $(\eta_j, \eta_{j+1}]$  in time t, denoted by  $N_d(y_j)$  is given by

$$N_d(y_j) = \Omega \sum_{k=1}^{m_j} \Delta y_{j,k}^+; j = 1, 2, ..., n$$
, (20)

where

 $y_i =$ class mark for the realization of  $Y_D$ ;

n = number of class intervals for the realization of

 $\Delta y_{i,k}^+$  = vertical rise of the bed in the class interval associated with  $y_i$  for the kth deposition period; and

 $m_i = \text{maximum number of bed forms contained in}$ the  $y_r(t)$  record and which also contain some deposition in the class interval associated with  $y_i$ .

Figure 1 illustrates the class marks,  $y_j$ , and the vertical rise of the bed,  $\Delta y_{i,k}^+$ , within the class intervals,  $\Delta y_i$ , for a typical  $y_x(t)$  record. It is clear that  $\Delta y_{j,k}^+ \leqslant \Delta y_j = \eta_{j+1} - \eta_j$  . The total number of particles per unit area deposited over all intervals, the denominator of equation 19, is designated by  $N_d$  and is obtained by summing equation 20 over all class marks:

$$N_d = \sum_{j=1}^n N_d(y_j) = \Omega \sum_{j=1}^n \sum_{k=1}^{m_j} \Delta y_{j,k}^+ \ . \ \ (21)$$
 where 
$$Y_E = \text{random variable describing the elevations at}$$
 which particles are eroded; 
$$y_i = \text{class mark for the realization of } Y_E;$$
 on 19 now becomes 
$$\eta_i, \eta_{i+1} = \text{lower and upper class limits of } y_i, \text{ respectively;}$$

Equation 19 now becomes

$$P[\eta_{j} < Y_{D} \leq \eta_{j+1}] = \lim_{m_{j} \to \infty} \frac{N_{d}(y_{j})}{N_{d}}$$

$$= \lim_{m_{j} \to \infty} \frac{\sum_{k=1}^{m_{j}} \Delta y_{j,k}^{+}}{\sum_{j=1}^{n} \sum_{k=1}^{m_{j}} \Delta y_{j,k}^{+}} . (22)$$

Similarly an analysis of the erosion periods can be used to estimate the probability that particles are eroded between the elevations  $\eta_i$  and  $\eta_{i+1}$ ,

$$P[\eta_{i} < Y_{E} \leq \eta_{i+1}] = \lim_{\substack{m'_{i} \to \infty}} \frac{N_{e}(y_{i})}{N_{e}}$$

$$= \lim_{\substack{m'_{i} \to \infty}} \frac{\sum_{\substack{m'_{i} \in Ay_{i,k}}}^{m'_{i}}}{\sum_{\substack{i=1 \ k=1}}^{m} \sum_{\substack{k=1 \ k=1}}^{m'_{i}}} , (23)$$

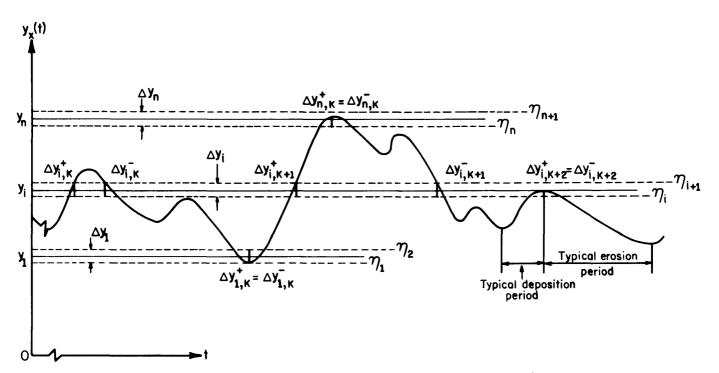


Figure 1. — Typical  $y_{x}(t)$  record illustrating the class marks for deposition and erosion.

 $m'_i$  = maximum number of bed forms contained in the  $y_x(t)$  record and which also contain some erosion in the class interval associated with  $y_i$ ;

 $N_e(y_i)$  = total number of particles per unit area eroded from the interval  $(\eta_i, \eta_{i+1}]$  centered at  $y_i$ ;

 $N_e$  = total number of particles per unit area eroded over all intervals;

n = number of class intervals for  $Y_E$ ; and

 $\Delta y_{i,k}^-$  = amount of erosion which occurred during the kth erosion period in the vertical class interval associated with  $y_i$  (fig. 1).

In the limit as  $m_j$  and  $m_i'$  approach infinity for a stationary record, the distributions of  $P[\eta_j < Y_D \le \eta_{j+1}]$  and  $P[\eta_i < Y_E \le \eta_{i+1}]$  must be identical. The elevation at which particles are eroded will hereafter simply be referred to as the elevation of erosion,  $Y_E$ .

To repeat, the following assumptions are necessary for equations 22 and 23 to be valid: (1) Flow is in equilibrium such that both deposition and erosion processes are stationary with respect to time t; (2) both erosion and deposition do not occur at the same point during the same time period; and (3) the number of particles per unit volume of the bed is constant.

If the measuring equipment were sensitive enough to detect the movement of single particles, the second assumption would not be necessary because it would be physically impossible for one particle to be eroded and another to be deposited at the same point and at the same time. For  $y_x(t)$  records obtained from less sensitive equipment, of course, the assumption may not be strictly true. When the bed material is not uniform in size, equations 20 through 23 are not strictly true because the number of particles per unit volume of the bed,  $\Omega$ , is a function of elevation due to a vertical sorting. However, equations 22 and 23 should serve as first approximations to the true probabilities,  $P[\eta_j < Y_D \le \eta_{j+1}]$  and  $P[\eta_i < Y_E \le \eta_{i+1}]$ , even for the nonuniform bed material.

If the number of particles per unit volume were known as a function of elevation, y, assumption (3) could be dropped. In this case, the counterpart of equation 22 is

$$P[\eta_{j} < Y_{D} \leq \eta_{j+1}]$$

$$= \lim_{m_{j} \to \infty} \frac{\Omega_{j} \sum_{k=1}^{m_{j}} \Delta y_{j,k}^{+}}{\sum_{j=1}^{n} \Omega_{j} \sum_{k=1}^{m_{j}} \Delta y_{j,k}^{+}} , \quad (24)$$

in which  $\Omega_j$  is the number of particles per unit volume of the bed associated with  $y_j$ . The counterpart of equation 23 would be similar. The value of  $\Omega_j$  could be obtained from core-sample segments taken from different elevations within the bed. In the next section,  $\Omega$  will be assumed to be a constant in estimation of  $P[\eta_j < Y_D \leqslant \eta_{j+1}]$  and  $P[\eta_i < Y_E \leqslant \eta_{i+1}]$ . Because the bed material was very uniform in size, however, the assumption should have been very good.

Equations 22 and 23 are estimated by use of the sample probability mass functions which are defined to be

$$p_{Y_D}(y_j) = \frac{N_d(y_j)}{N_d}$$

$$= P[\eta_j < Y_D \le \eta_{j+1}] \text{ for a large } m_j , \quad (25)$$

and 
$$p_{Y_E}(y_i) = \frac{N_e(y_i)}{N_e}$$
 
$$\sum_{i=1}^{n} P[\eta_i < Y_E \le \eta_{i+1}] \text{ for a large } m_i \text{ , (26)}$$

in which  $P_{Y_D}(y_j)$  is used to estimate equation 22 and  $P_{Y_E}(y_i)$  is used to estimate equation 23. Estimates of the mean and variance of the elevation of deposition are respectively

$$\hat{E}[Y_D] = \sum_{j=1}^n y_j p_{Y_D}(y_j) ,$$
and
$$\hat{Var}[Y_D] = \left[\sum_{j=1}^n y_j^2 p_{Y_D}(y_j)\right] - \left[\sum_{j=1}^n y_j p_{Y_D}(y_j)\right]^2$$

in which  $\hat{E}[\cdot]$  and  $\hat{Var}[\cdot]$  denote estimates of the expected value and variance, respectively. Replacing the D's with E's and the j's with i's in equation 27 gives the estimates for the mean and variance of the elevation of erosion.

The probability density functions for the elevations of deposition and erosion  $[f_{Y_D}(y)]$  and  $f_{Y_E}(y)$  may be inferred from the probability mass functions  $[p_{Y_D}(y_i)]$  and  $p_{Y_E}(y_i)$  by means of a statistical fitting procedure. This will be discussed later.

## ESTIMATION OF THE PROBABILITY DISTRIBUTIONS OF THE REST PERIODS

In this study the rest period of a particle is defined as the time lapse between the burial and reexposure of the particle. This definition is consistent with the assumption that erosion and deposition do not occur at the same point during the same time period, and it is also necessary in analyzing single particle measurements because the measurement techniques cannot detect momentary rests by the particle. Using the burial definition, the  $y_x(t)$  record provides a means of estimating the probability density functions of particle rest periods conditioned on the elevations of deposition. The method is illustrated schematically in figure 2, where the statistic  $\{t_{i,k}; j = 1, 2, ..., n; k = 1, 2, ..., n\}$  $m_{i,j}$  measures the conditional rest period, the index k signifies a particular bed form. The term  $m_{j,j}$  designifies a nates the maximum number of bed forms which are contained in the  $y_r(t)$  record and which also contain both an up-crossing and a down-crossing at the elevation  $y_r$ . This use of the  $y_r(t)$  record was first suggested by Hubbell and Sayre (1965) and was partly evaluated by Grigg (1969).

A relative frequency analysis of the statistic  $\{t_{j,k}\}$  leads to a sample conditional probability mass function of the rest periods which is defined to be

$$P_{T \setminus Y_{D}}(t_{\alpha} \setminus y_{j}) = P[\tau_{\alpha} < T \le \tau_{\alpha + 1} \setminus \eta_{j} < Y_{D} \le \eta_{j + 1}] ;$$

$$j = 1, 2, ..., n; \alpha = 1, 2, ..., r , (28)$$

where

T = random variable describing the rest periods;

$$t_{\alpha}$$
,  $y_{j}$  = class marks for  $T$  and  $Y_{D}$ , respectively;  $\tau_{\alpha}$ ,  $\tau_{\alpha+1}$  = lower and upper class limits of  $t_{\alpha}$ , respectively; and

r = number of class intervals for T.

Equation 25 can be used to release the condition on equation 28 and to obtain the marginal sample probability mass function for the rest periods,

$$p_{T}(t_{\alpha}) = P[\tau_{\alpha} < T \le \tau_{\alpha+1}] = \sum_{j=1}^{n} p_{T \setminus Y_{D}}(t_{\alpha} \setminus y_{j}) p_{Y_{D}}(y_{j});$$

$$\alpha = 1, 2, \dots, r . . (29)$$

From equations 28 and 29 the corresponding probability density functions for the conditional rest periods,  $f_{T \setminus Y_D}(t \setminus y)$ , and for the marginal rest periods,  $f_T(t)$ , may be approximated by means of a statistical fitting procedure.

The mean and variance of the conditional rest periods are estimated from the sample moments

$$\hat{E}[T \setminus Y_D = y_j] = \frac{1}{m_{j,j}} \sum_{k=1}^{m_{j,j}} t_{j,k} ,$$
and
$$Var[T \setminus Y_D = y_j]$$

$$= \frac{1}{m_{j,j}} \left[ \sum_{k=1}^{m_{j,j}} (t_{j,k})^2 \right] - \left[ \frac{1}{m_{j,j}} \sum_{k=1}^{m_{j,j}} t_{j,k} \right]^2$$
(30)

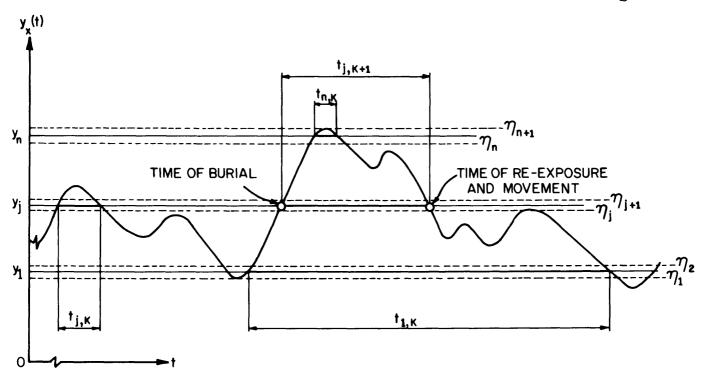


Figure 2. — Typical  $y_x(t)$  record illustrating the conditional rest periods of a particle.

and

Estimates of the mean and variance of the marginal rest periods are respectively

$$\hat{E}[T] = \sum_{j=1}^{n} \hat{E}[T \setminus Y_{D} = y_{j}] p_{Y_{D}}(y_{j}) ,$$

$$\hat{Var}[T] = \hat{E}[T^{2}] - (\hat{E}[T])^{2} ,$$
(31)

in which

$$\begin{split} \hat{\mathbf{E}}[T^2] &= \sum_{j=1}^{n} \hat{\mathbf{E}}[T^2 \setminus Y_D = y_j] p_{Y_D}(y_j) \\ &= \frac{1}{m_{j,j}} \sum_{k=1}^{m_{j,j}} (t_{j,k})^2 p_{Y_D}(y_j) \end{split} .$$

The joint probability density function of T and  $Y_D$ , denoted by  $f_{T,Y_D}(t,y)$  is estimated from the sample joint probability mass function defined to be

$$p_{T, Y_D}(t_{\alpha}, y_j) = P[\tau_{\alpha} < T \le \tau_{\alpha + 1}, \eta_j < Y_D \le \eta_{j + 1}], (32)$$

where  $j=1,\,2,\,\ldots,\,n$  and  $a=1,\,2,\,\ldots,\,r$ . From equations 25 and 28,  $P_{T,\,Y_D}(t_\alpha,\,y_j)$  is completely determined such that

$$p_{T,Y_{D}}(t_{\alpha}, y_{j}) = p_{T \setminus Y_{D}}(t_{\alpha} \setminus y_{j}) p_{Y_{D}}(y_{j}) . \tag{33}$$

Finally the correlation coefficient between T and  $Y_D$  is estimated to be

$$\hat{\rho}_{T,Y_D} = \frac{\hat{\mathbb{E}}[TY_D] - \hat{\mathbb{E}}[T]\hat{\mathbb{E}}[Y_D]}{\sqrt{\hat{\mathbb{A}}r[T]}\sqrt{\hat{\mathbb{A}}r[Y_D]}} , \qquad (34)$$

in which

$$\hat{E}[TY_D] = \sum_{\alpha=1}^r \sum_{j=1}^n t_{\alpha} y_j p_{T, Y_D}(t_{\alpha', Y_j}) ,$$

 $\hat{\mathbf{E}}[Y_D]$  and  $\hat{\mathbf{Var}}[Y_D]$  are given in equation 27, and  $\hat{\mathbf{E}}[T]$  and  $\hat{\mathbf{Var}}[T]$  are given in equation 31. The joint distribution expresses the relation between the rest period and the elevation of deposition and the correlation coefficient measures a degree of linear association between the rest period and the elevation of deposition.

If the shape of the  $y_x(t)$  record is dependent on the flow conditions and bed-material properties, then the rest period statistics as determined by equations 28 through 34 are also functions of flow conditions and bed-material properties.

In summary, the probability distribution for the marginal rest period of a sediment particle, the rest period conditioned on the elevation of deposition, and the elevation of particle deposition and erosion can all be obtained from a continuous record of the bed elevation at a single point as a function of time. The only assumptions that are needed are: (1) Both erosion and deposition do not occur at the same point at the same time; (2) bed elevation is stationary (in the statistical sense); and (3) the number of particles per unit volume of the bed is constant. These assumptions are not severely restrictive, and the results are equally applicable to both field and laboratory analysis.

## ESTIMATION OF THE PROBABILITY DISTRIBUTIONS OF THE STEP LENGTHS

The  $y_x(t)$  record contained the information necessary to estimate the probability distributions of the rest periods. Both the  $y_x(t)$  and the  $y_t(x)$  records are necessary to determine the step length statistics. Unfortunately, more assumptions are also necessary and these assumptions may be considerably more restrictive than the ones made up to this time.

As previously mentioned, it will be assumed that each sediment particle on the stoss side of a dune makes a step in the downstream direction before it is deposited on the slip face of any dune. Once deposited it rests there until it is reexposed on the stoss side. Let  $E_{i,j,v}$  be the event that a particle, eroded from elevation,  $y_i$ , of the stoss side of a dune, passes v dune crests before it is deposited at elevation,  $y_j$ . Then the statistic  $\{x_{i,j,v,k}; i,j=1,2,\ldots,n; v=1,2,\ldots; k=1,2,\ldots,m_{i,j,v}\}$  (fig. 3) is the measure of the conditional step length of the event,  $E_{i,j,v}$ . The term  $m_{i,j,v}$  represents the total number of possibilities of the event  $E_{i,j,v}$  contained in the  $y_t(x)$  record and the index k specifies a particular possibility. In general, the term  $m_{i,j,v}$  will be different for each combination of values i, j, and v.

A frequency analysis of the statistic  $\{x_{i,j,v,k}\}$  gives a sample conditional probability mass function which is defined to be

$$\begin{split} & p_{X \setminus Y_E, Y_D, E_{\mathbf{v}}}(x_{\beta} \setminus y_i, y_j, \mathbf{v}) \\ &= P[\lambda_{\beta} < X \le \lambda_{\beta+1} \setminus \eta_i < Y_E \le \eta_{i+1}, \eta_j < Y_D \le \eta_{j+1}, E_{\mathbf{v}}]; \\ & \beta = 1, 2, \dots, s; i, j, = 1, 2, \dots, n; \mathbf{v} = 1, 2, \dots, (35) \end{split}$$

where

X = random variable describing the step lengths;

 $x_{\beta} =$ class mark for the realizations of X;

 $\lambda_{\beta}$ ,  $\lambda_{\beta+1} = \text{lower}$  and upper class limits of  $x_{\beta}$ , respectively;

s = number of class intervals for the realizations of X; and

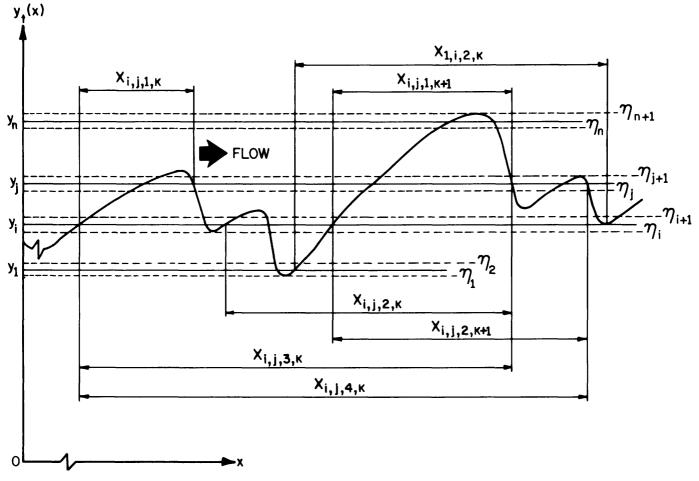


FIGURE 3. — Statistic,  $\{x_{i,j,v,k}; i,j=1,2,\ldots,n; v=1,2,\ldots; k=1,2,\ldots,m_{i,j,v}\}$ , for the step length of a particle.

 $E_v =$ event that a particle passes v dune crests before it is deposited (fig. 3).

The corresponding probability density function,

$$f_{X \smallsetminus Y_E, \ Y_D, E_v}(x \backslash y', \ y, v),$$
 may be determined from

$$p_{X\setminus Y_E, Y_D, E_v}(x_{\beta}\setminus y_i, y_j, v),$$

and its mean and variance are estimated to be

$$\hat{E}[X \setminus Y_E = y_i, Y_D = y_j, E_v] = \frac{1}{m_{i,j,v}} \sum_{k=1}^{m_{i,j,v}} x_{i,j,v,k},$$
and
$$\hat{Var}[X \setminus Y_E = y_i, Y_D = y_j, E_v]$$

$$= \frac{1}{m_{i,j,v}} \sum_{k=1}^{m_{i,j,v}} (x_{i,j,v,k})^2 - \left[\frac{1}{m_{i,j,v}} \sum_{k=1}^{m_{i,j,v}} x_{i,j,v,k}\right]^2$$
(36)

in which

$$\frac{1}{m_{i,j,v}} \sum_{k=1}^{m_{i,j,v}} (x_{i,j,v,k})^2 = \hat{E}[X^2 \setminus Y_E = y_i, y_D = y_j, E_v].$$

If  $Y_E$ ,  $Y_D$ , and  $E_v$  are mutually independent (it seems to be reasonable that after a particle passes the crest of a dune it has probably lost track of where it came from), the density function  $[f_{X \setminus Y_D, E_v}(x \setminus y, v)]$  of the step lengths given that a particle is deposited at elevation y after passing v dune crests is estimated from the sample conditional mass function which is given by

$$p_{X \setminus Y_{D}, E_{v}}(x_{\beta} \setminus y_{j}, v)$$

$$= P[\lambda_{\beta} < X \leq \lambda_{\beta+1} \setminus \eta_{j} < Y_{D} \leq \eta_{j+1}, E_{v}]$$

$$= \sum_{i=1}^{n} p_{X \setminus Y_{E}, Y_{D}, E_{v}}(x_{\beta} \setminus y_{i}, y_{j}, v) p_{Y_{E}}(y_{i}) \quad (37)$$

in which  $p_{Y_E}(y_i)$  is given by equation 26. The mean and

variance of the step lengths of a particle which is deposited at  $y_j$  after passing v dune crests are estimated to be

$$\hat{E}[X \setminus Y_D = y_j, E_v] = \sum_{i=1}^n \hat{E}[X \setminus Y_E = y_i, Y_D = y_j, E_v] p_{Y_E}(y_i),$$
and
$$\hat{Var}[X \setminus Y_D = y_j, E_v] = \hat{E}[X^2 \setminus Y_D = y_j, E_v] - (\hat{E}[X \setminus Y_D = y_j, E_v])^2,$$
(38)

in which

$$\hat{\mathbf{E}}[X^2 \backslash Y_D = y_j, E_v]$$

 $p_{X \setminus Y_D}(x_{\beta} \setminus y_j)$ 

$$= \sum_{i=1}^{n} \hat{E}[X^2 \setminus Y_E = y_i, Y_D = y_j, E_v] p_{Y_E}(y_i) .$$

Likewise, the following sample probability mass functions and corresponding means and variances are obtained:

$$= P[\lambda_{\beta} < X \leq \lambda_{\beta+1} \setminus \eta_{j} < Y_{D} \leq \eta_{j+1}]$$

$$= \sum_{\nu=1}^{\infty} p_{X \setminus Y_{D}, E_{\nu}}(x_{\beta} \setminus y_{j}, \nu) P[E_{\nu}]$$

$$\hat{E}[X \setminus Y_{D} = y_{j}] = \sum_{\nu=1}^{\infty} \hat{E}[X \setminus Y_{D} = y_{j}, E_{\nu}] P[E_{\nu}]$$

$$Var[X \setminus Y_{D} = y_{j}]$$

$$= \left[\sum_{\nu=1}^{\infty} \hat{E}[X^{2} \setminus Y_{D} = y_{j}, E_{\nu}] P[E_{\nu}]\right] - (\hat{E}[X \setminus Y_{D} = y_{j}])^{2};$$

$$p_{X \setminus E_{\nu}}(x_{\beta} \setminus \nu) = P[\lambda_{\beta} < X \leq \lambda_{\beta+1} \setminus E_{\nu}]$$

$$= \sum_{j=1}^{n} p_{X \setminus Y_{D}, E_{\nu}}(x_{\beta} \setminus y_{j}, \nu) p_{Y_{D}}(y_{j})$$

$$\hat{E}[X \setminus E_{\nu}] = \sum_{j=1}^{n} \hat{E}[X \setminus Y_{D} = y_{j}, E_{\nu}] p_{Y_{D}}(y_{j})$$

$$Var[X \setminus E_{\nu}]$$

$$= \left[\sum_{j=1}^{n} \hat{E}[X^{2} \setminus Y_{D} = y_{j}, E_{\nu}] p_{Y_{D}}(y_{j})\right] - (\hat{E}[X \setminus E_{\nu}])^{2};$$

$$(40)$$

$$p_{X}(x_{\beta}) = P[\lambda_{\beta} < X \leq \lambda_{\beta+1}]$$

$$= \sum_{\nu=1}^{\infty} p_{X \setminus E_{\nu}}(x_{\beta} \setminus \nu) P[E_{\nu}]$$

$$\hat{E}[X] = \sum_{\nu=1}^{\infty} \hat{E}[X \setminus E_{\nu}] P[E_{\nu}]$$
(41)

The density functions,  $f_{X \setminus Y_D}(x \setminus y)$ ,  $f_{X \setminus E_v}(x \setminus v)$ , and  $f_X(x)$  are estimated from equation sets 39, 40, and 41, respectively.

 $\widehat{\text{Var}[X]} = \left[\sum_{i=1}^{\infty} \widehat{E}[X^2 \setminus E_{v}] P[E_{v}]\right] - (\widehat{E}[X])^2$ 

The joint probability density function of X and  $Y_D$ , conditioned on the event  $E_v$ ,  $[f_{X,Y_D\setminus E_v}(x,y\setminus v)]$  can be estimated from a sample joint probability mass function,

$$P_{X,Y_D \setminus E_{\nu}}(x_{\beta}, y_j \setminus \nu)$$

$$= P[\lambda_{\beta} < X \le \lambda_{\beta+1}, \eta_j < Y_D \le \eta_{j+1} \setminus E_{\nu}] \quad (42)$$

and

$$\begin{aligned} p_{X,Y_D \setminus E_{\mathbf{v}}}(x_{\beta} \setminus y_j, \mathbf{v}) &= p_{X \setminus Y_D, E_{\mathbf{v}}}(x_{\beta} \setminus y_j, \mathbf{v}) p_{Y_D \setminus E_{\mathbf{v}}}(y_j \setminus \mathbf{v}) \\ &= p_{X \setminus Y_D, E_{\mathbf{v}}}(x_{\beta} \setminus y_j, \mathbf{v}) p_{Y_D}(y_j) \quad . \end{aligned} \tag{43}$$

Note that  $p_{Y_D \setminus E_v} (y_j \setminus v) = p_{Y_D} (y_j)$  because  $Y_D$  and  $E_v$  were assumed to be independent. The correlation coefficient of X and  $Y_D$ , conditioned on the event  $E_v$ , is then estimated to be

$$\hat{\rho}_{X,Y_D \setminus E_V} = \frac{\hat{E}[XY_D \setminus E_V] - \hat{E}[X \setminus E_V] \hat{E}[Y_D]}{\sqrt{\hat{Var}[X \setminus E_V]} \sqrt{\hat{Var}[Y_D]}} , (44)$$

in which

$$\hat{E}[XY_D \setminus E_v]$$

$$= \sum_{\beta=1}^{s} x_{\beta} \left\{ \sum_{j=1}^{n} y_{j} p_{X \setminus Y_{D}, E_{V}}(x_{\beta} \setminus y_{j}, v) p_{Y_{D}}(y_{j}) \right\} , \text{ and}$$

 $\hat{\mathbf{E}}[Y_D]$ ,  $\hat{\mathbf{Var}}[Y_D]$ ,  $\hat{\mathbf{E}}[X \setminus E_v]$ , and  $\hat{\mathbf{Var}}[X \setminus E_v]$  are given by equation sets 27 and 40. Similarly, the joint probability density function of X and  $Y_D$ ,  $[f_{X,Y_D}(x,y)]$  and the

corresponding correlation coefficient,  $\rho_{X,Y_D}$ , are estimated to be

$$p_{X,Y_D}(x_{\beta},y_j) = p_{X \setminus Y_D}(x_{\beta} \setminus y_j) p_{Y_D}(y_j) , \qquad (45)$$

and

$$\hat{\rho}_{X,Y_D} = \frac{\hat{\mathbb{E}}[XY_D] - \hat{\mathbb{E}}[X]\hat{\mathbb{E}}[Y_D]}{\sqrt{\hat{\mathbb{Var}}[X]} \sqrt{\hat{\mathbb{Var}}[Y_D]}},$$
 (46)

in which

$$\hat{\mathbb{E}}[XY_D] = \sum_{\beta=1}^s x_\beta \left\{ \sum_{i=1}^n y_j p_{X \setminus Y_D}(x_\beta \setminus y_j) p_{Y_D}(y_j) \right\} \ .$$

The joint distributions (eqs. 43 and 45) express the relation between the step length and the elevation of deposition, and the correlation coefficients (eqs. 44 and 46) measure the degree of linear association between the step length and the elevation of deposition.

The American Society of Civil Engineers (Task Committee on Preparation of Sedimentation Manual, 1962) defines bedload as that material moving on or near the bed. Accepting this general definition, it would appear consistent to count any sediment particle which was able to skip across a dune trough as suspended load. since it would be extremely unlikely that a particle would be able to pass the trough while moving on or near the bed. A very precise, and admittedly restrictive, definition of bedload is used for the purpose of this report. For the purposes of this report, bed load is defined as that part of bed material which is deposited on the downstream face of the dune from which it is eroded. Then the suspended load must be that material which is not deposited on the downstream face of the dune from which it is eroded, that is, all sediment particles which pass two or more dune crests before being deposited. The same particle could be counted as bedload during one step but as suspended load during the next step. By definition then, it follows that  $P[E_1]$  = probability that a particle is transported as the bedload; and

$$1 - P[E_1] \equiv \sum_{v=2}^{\infty} P[E_v] =$$

probability that a particle is transported as the suspended load during any step. The probability distributions and moments for the step lengths of a bed-load particle may be obtained by putting v=1 in the sets of equations 35 through 40 and 42 through 48.

For a bed material composed of coarse sand it seems reasonable to assume that all particles are transported as bed load, that is, all particles eroded from the stoss side of a dune will be deposited on the downstream side of the same dune; and therefore,  $P[E_1] \equiv 1$ , and  $P[E_v] \equiv 0$  for  $v \ge 2$ . For this case, a frequency analysis of the statistic  $\{x_{i,j,v,k}; i,j=1, 2, \ldots, n; k=1, 2, \ldots, m_{i,j,v}; v=1\}$  gives a sample conditional probability mass function which is defined to be

$$\begin{aligned} & p_{X \setminus Y_E, Y_D}(x_{\beta} \setminus y_i, y_j) \\ &= P[\lambda_{\beta} < X \le \lambda_{\beta+1} \setminus \eta_i < Y_E \le \eta_{i+1}, \, \eta_j < Y_D \le \eta_{j+1}]; \\ & \beta = 1, 2, \dots, s; \quad i, j = 1, 2, \dots, n \end{aligned}$$
(47)

The corresponding probability density function,  $f_{X \setminus Y_E, Y_D}(x \setminus y', y)$ , may be approximated from  $p_{X \setminus Y_E, Y_D}(x_\beta \setminus y_i, y_j)$ . Denoting the statistic  $\{x_{i,j,v,k}; i,j=1,2,\ldots,k=1,2,\ldots,m_{i,j,v}; v=1\}$  simply as  $\{x_{i,j,k}; i,j=1,2,\ldots,n; k=1,2,\ldots,m_{i,j}\}$  (fig. 3), the corresponding mean and variances are estimated to be

$$\hat{E}[X \setminus Y_E = y_i, Y_D = y_j] = \frac{1}{m_{i,j}} \sum_{k=1}^{m_{i,j}} x_{i,j,k},$$
and
$$\hat{Var}[X \setminus Y_E = y_i, Y_D = y_j]$$

$$= \left(\frac{1}{m_{i,j}} \sum_{k=1}^{m_{i,j}} (x_{i,j,k})^2\right) - \left(\frac{1}{m_{i,j}} \sum_{k=1}^{m_{i,j}} x_{i,j,k}\right)^2.$$
(48)

The term  $m_{i,j}$  represents the total number of bed forms in the sample for which the upstream side intersects the elevation  $y_j$  and the downstream side intersects the elevation  $y_i$ . In general, the term  $m_{i,j}$  will be different for each combination of values of i and j (fig. 3).

Based on the statistic  $\{x_{i,j,k}\}$  and assuming that  $Y_E$  and  $Y_D$  are mutually independent, the probability density functions,  $f_{X \setminus Y_D}(x \setminus y)$ ,  $f_X(x)$ , and  $f_{X,Y_D}(x,y)$  as well as the corresponding moments are estimated by setting  $P[E_1] = 1$  in equation sets 37, 38, 39, 40, 42, and 44.

Since the bed-form shape and rate of movement are dependent on the flow condition and bed material properties, it should be clear that equations 35 through 48 are also functions of the flow condition and bed material properties. The statistic  $\{x_{i,j,k}\}$  will be analyzed later to estimate the various probability distributions of the step lengths for a coarse sand for three different flow conditions.

Summarizing this section, the step length distributions can be estimated by combining the information contained in the  $y_x(t)$  and  $y_t(x)$  records. Additional assumptions are required however. These are: (1) No deposition occurs on the upstream sides of dunes, and no erosion occurs on the downstream faces of dunes; and (2) the elevation of particle erosion,  $Y_E$ , the elevations

tion of particle deposition,  $Y_D$ , and the event that a particle passes v dune crests before it is deposited,  $E_v$ , are mutually independent. The first assumption may not be strictly true especially, due to flow separation, in the neighborhood of dune trough where both deposition and erosion may occur at the same point. For dune flow conditions, however, laboratory observation shows that such an area is small enough that the results should be applicable without an appreciable error. The second assumption seems to be reasonable because as a sediment particle passes a dune crest it likely loses the memory of where it came from. Estimation of  $P[E_n]$ would not appear to be a simple task. However, for a bed material composed of coarse sand, the assumption that all particles which are eroded from the stoss side of a dune will be deposited on the downstream side of the same dune seems to be reasonable.

#### **BED-MATERIAL TRANSPORT EQUATIONS**

The mean transport speed of a bed material particle,  $V_T$ , is estimated to be

 $\hat{V}_T = \frac{\text{Total distance traveled by a particle after } m \text{ steps}}{\text{Total time required for a particle to make } m \text{ steps}}$ 

$$= \frac{m\hat{E}[X]}{m\hat{E}[T]} = \frac{\hat{E}[X]}{\hat{E}[T]}$$

$$= \frac{\sum_{j=1}^{n} \hat{E}[X \setminus Y_D = y_j] p_{Y_D}(y_j)}{\sum_{j=1}^{n} \hat{E}[T \setminus Y_D = y_j] p_{Y_D}(y_j)}$$
 for a large  $m$  (49)

in which  $\hat{V}_T$  denotes an estimate of the mean transport speed,  $V_T$ ,

$$\hat{E}[X] = \sum_{v=1}^{\infty} \sum_{j=1}^{n} \sum_{i=1}^{n} \left[ \frac{1}{m_{i,j,v}} \sum_{k=1}^{m_{i,j,v}} x_{i,j,v,k} \right]$$

$$p_{Y_E}(y_i) p_{Y_D}(y_j) P[E_{v}]$$

and

$$\hat{E}[T] = \sum_{j=1}^{n} \left[ \frac{1}{m_{j,j}} \sum_{k=1}^{m_{j,j}} t_{j,k} \right] p_{Y_{D}}(y_{j}) .$$

In equation 49, the duration of particle movement is assumed to be negligible compared to the rest period. This assumption will be used throughout this section.

The mean transport speed could also be estimated to be

$$\hat{V}_{T}^{\prime} = \sum_{j=1}^{n} \frac{\hat{E}[X \setminus Y_{D} = y_{j}]}{\hat{E}[T \setminus Y_{D} = y_{j}]} p_{Y_{D}}(y_{j})$$
 (50)

where  $\hat{V}_T'$  also estimates the mean transport speed of a bed material particle. In general it can be shown that  $\hat{V}_T \neq \hat{V}_T'$  and the results of this study will indicate that  $\hat{V}_T \leqslant \hat{V}_T'$ . Now the question is: Which one will give the better estimate of the mean bed material transport rate? The difference between the two equations is the manner in which the events are averaged or weighted. So, in order to answer the question, one must depend upon physical arguments and reasoning. In equation 50, the average speed of a particle at each elevation is weighted by the number of particles with this speed. Equation 50 gives the best estimate of the arithmetic mean of individual particle speeds. On the other hand, equation 49 computes the estimate of the mean particle speed as the total distance traveled by a number of particles divided by the amount of time required to transport the same number of particles. In other words, the center of mass of a group of particles is translated through a distance,

$$\sum_{j=1}^{n} \widehat{\mathbf{E}}[X \setminus Y_D = y_j] \ p_{Y_D}(y_j)$$

in time,

$$\sum_{j=1}^{n} \widehat{\mathbf{E}} [T \setminus Y_D = y_j] \ p_{Y_D}(y_j).$$

Equation 49 will be used in this section because it is based on a mass flux concept and it gives an unbiased estimate of the mean sediment transport rate. The mean particle speed given by equation 50 will be useful in the study of bed material dispersion because it is based on individual particle speeds.

Defining the bed load and suspended load as given in the previous section, the mean transport speed of a bed-load particle,  $V_B$ , is estimated to be

$$\hat{V}_B = \frac{\hat{E}[X \setminus E_1]}{\hat{E}[T]} , \qquad (51)$$

where  $\hat{V}_B$  is an estimate of  $V_B$  and

$$\hat{E}[X \setminus E_1] = \sum_{j=1}^{n} \sum_{i=1}^{n} \left[ \frac{1}{m_{i,j}} \sum_{k=1}^{m_{i,j}} x_{i,j,k} \right] p_{Y_E}(y_i) p_{Y_D}(y_j)$$

Similarly, the mean transport speed of a suspended load particle,  $V_S$ , is estimated to be

$$\hat{V}_{S} = \frac{\hat{E}\left[X \middle| \bigcup_{v=2}^{\infty} E_{v}\right]}{\hat{E}[T]} = \frac{\sum_{v=2}^{\infty} \hat{E}[X \middle| E_{v}] P[E_{v}]}{\hat{E}[T] \sum_{v=2}^{\infty} P[E_{v}]}, (52)$$

where

and

$$\hat{V}_S= ext{estimate of }V_s;$$
  $\bigcup_{v\ =\ 2}E_v= ext{union of event, }E_v ext{, for }v\geqslant 2.$ 

 $\hat{\mathbf{E}}[X \setminus E_v]$  is given in equation 40. Note that

$$P\left[\bigcup_{v=1}^{\infty} E_{v}\right] = \sum_{v=1}^{\infty} P[E_{v}] = 1$$

$$\hat{E}\left[X \middle| \bigcup_{v=2}^{\infty} E_{v}\right] = \frac{\sum_{v=2}^{\infty} \hat{E}[X \backslash E_{v}] P[E_{v}]}{\sum_{v=2}^{\infty} P[E_{v}]}$$
(53)

Using equations 49, 51, 52, and 53,

$$\hat{V}_{T} = \frac{\hat{E}[X \setminus E_{1}]}{\hat{E}[T]} P[E_{1}] + \frac{\sum_{v=2}^{\infty} \hat{E}[X \setminus E_{v}] P[E_{v}]}{\hat{E}[T]}$$

$$= \hat{V}_{B} P[E_{1}] + \hat{V}_{S} \sum_{v=2}^{\infty} P[E_{v}]$$

$$= \hat{V}_{B} P[E_{1}] + \hat{V}_{S} (1 - P[E_{1}]) . \quad (54)$$

If all bed material particles have identical transport characteristics, which is reasonable for uniformly sized bed material, the mean total bed material discharge is obtained by use of the continuity concept,

$$\hat{Q}_T = \gamma_s (1 - \theta) W h \hat{V}_T \quad , \tag{55}$$

where

 $\hat{Q}_T$  = estimate of the mean total bed material discharge in weight per unit time;

 $\gamma_s$  = specific weight of the bed material;

 $\Theta$  = porosity of the bed;

W =width of the channel;

h = average depth of the zone in which bed material movement occurs; and  $\hat{V}_T$  is given in equation 49.

Similarly, estimates of the mean bed-load discharge and suspended load discharge are, respectively,

$$\hat{Q}_B = \gamma_S (1 - \theta) W h \hat{V}_B P[E_1] \quad , \tag{56}$$

and

$$\hat{Q}_{S} = \gamma_{S} (1 - \theta) W h \hat{V}_{S} (1 - P[E_{1}]) \quad , \tag{57}$$

where  $\hat{V}_B$  and  $\hat{V}_S$  are given in equations 51 and 52, respectively. From equations 54 through 57,

$$\hat{Q}_T = \hat{Q}_B + \hat{Q}_S \quad . \tag{58}$$

Although equations 55, 56, and 57 have the form of a continuity equation, the concept of continuity applies only in a statistical sense, because particles move only when they are exposed on the stoss side of a dune or when they are in suspension. Hubbell and Sayre (1964) proposed that the average depth of the zone of bed material movement, h, be estimated from the  $y_t(x)$ record. For this method, the length of the reach for which h is to be determined is divided into sections. Starting from the upstream end, each section of length li extends from the dune trough at which the section begins to the first trough downstream that is deeper relative to a line parallel to the plane of the mean bed surface. After sectioning, a mean depth of sand above the projected line for each section,  $h_i$  is determined, and the h for the total reach,  $L_x$ , is computed as the weighted average of the  $h_i$ 's for each section. Expressed mathematically,

$$h = \frac{1}{L_x} \sum_{i=1}^{m} \ell_i h_i . agen{59}$$

The reasoning behind the procedure is based upon the assumption that although the individual dunes may change shape as they progress downstream, a statistical constancy of form exists over a long reach. Hence, quantitatively the particles subject to movement are those that would move if the entire profile were to progress downstream without changing form, and the depth of bed material movement is defined by lines that are parallel to the mean bed surface and extend downstream from the deepest trough.

If all bed material particles are assumed to be transported as the bedload,

$$\hat{Q}_T = \hat{Q}_B = \gamma_s (1 - \theta) W h \hat{V}_B , \qquad (60)$$

where  $\hat{V}_B$  is determined from equation 51. For coarse sand  $P[E_1]$  is expected to be very close to unity because the suspended load is negligible compared to the bed load. For a fine sand for which  $P[E_1] \neq 1$ , equation 60 would give only an approximation to the total load.

Equations 55, 56, and 57 can be used with measured  $y_x(t)$  and  $y_t(x)$  records to compute the various transport rates. However, in order to apply the equations to the prediction of the bed material transport rate where the  $y_x(t)$  and  $y_t(x)$  records are not available, the relations,  $E[X \setminus Y_E = y_b, Y_D = y_b, E_v]$ ,

 $\hat{E}[X \setminus Y_E = y_i, Y_D = y_j, E_v],$  $\hat{E}[T \setminus Y_D = y_j], P[E_v], p_{Y_D}(y_j), p_{Y_E}(y_i),$  and h, to pertinent hydraulic and sediment parameters must be established.

The mean transport speed of bed material particles deposited at elevation  $y_j$ , which is denoted by  $V_T(j)$ , (more precisely, deposited between the elevations  $\eta_j$  and  $\eta_{j+1}$ , centered at  $y_i$ ) may be estimated

$$\hat{V}_T(j) = \frac{\hat{E}[X \setminus Y_D = y_j]}{\hat{E}[T \setminus Y_D = y_j]} , \qquad (61)$$

where  $\hat{V}_{T}(j)$  estimates  $V_{T}(j)$ ,

$$\hat{\mathbf{E}}[X \setminus Y_D = y_i]$$

$$= \sum_{v=1}^{\infty} \sum_{i=1}^{n} \left[ \frac{1}{m_{i,j,v}} \sum_{k=1}^{m_{i,j,v}} x_{i,j,v,k} \right] p_{Y_{E}}(y_{i}) P[E_{v}] ,$$

and

$$\hat{E}[T \setminus Y_D = y_j] = \frac{1}{m_{j,j}} \sum_{k=1}^{m_{j,j}} t_{j,k} .$$

Based on equation 61, another transport equation can be developed as follows. Let  $\xi_j$  denote the percentage of volume between elevations  $\eta_j$  and  $\eta_{j+1}$  occupied by

dunes over a given reach; then,  $\xi_j$  can be estimated from the  $y_i(x)$  record (fig. 4),

$$\xi_j = \frac{1}{L_x} \sum_k \lambda_{j,k} , \qquad (62)$$

where  $L_x$  is the total length of  $y_t(x)$  record, and  $\lambda_{j,k}$  is defined in figure 4. Applying equations 61 and 62, the mean total bed material discharge can be expressed as

$$Q_{T}^{\perp} = \gamma_{s} (1 - \Theta) w \sum_{j=1}^{n} \hat{V}_{T}(j) \Delta y_{j} \xi_{j}$$

$$= \gamma_s (1 - \theta) w \sum_{j=1}^{n} \frac{\hat{\mathbb{E}}[X \setminus Y_D = y_j]}{\hat{\mathbb{E}}[T \setminus Y_D = y_j]} \Delta y_j \xi_j \quad , \quad (63)$$

where

 $\hat{Q}_T'$  = estimate of mean total bed material discharge; and

 $\Delta y_j$  = nonstandardized class width associated with elevation  $y_j(\Delta y_j = \eta_{j+1} - \eta_j)$ .

Equation 63 takes into account the local variation of the depth of the zone of bed material movement with respect to the elevation of deposition, and demonstrates to what extent each elevation contributes to the total transport rate. Similarly, the mean transport speed of a bed-load particle deposited at elevation  $y_j$ , which is denoted by  $V_B(j)$ , can be estimated from equation 61 by considering  $P[E_1] = 1$  and  $P[E_v] \equiv 0$  for  $v \ge 2$ . It can be shown, of course, that equations 54, 56, and 57 also apply at each elevation j as well as to depth-averaged values.

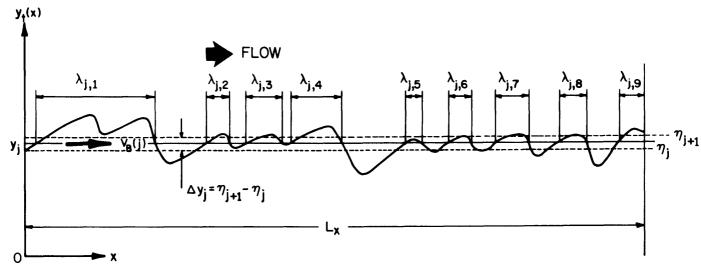


FIGURE 4. — Method for estimating the percentage of volume occupied by dunes between elevations  $\eta_j$  and  $\eta_{j+1}$ ;  $\xi_j = \frac{1}{L_x} \sum_{k=1}^{n} \lambda_{j,k}$ 

The third method to compute the mean total bed material discharge is based on the following reasoning:

$$Q_T = \begin{bmatrix} \text{Number of particles} \\ \text{deposited per} \\ \text{unit time and area} \end{bmatrix} \times \begin{bmatrix} \text{Weight} \\ \text{per} \\ \text{particle} \end{bmatrix}$$

where  $Q_T$  is the mean total bed material discharge in weight per unit time. Restricting the attention to elevation  $y_j$ , the terms in equation 64 are estimated as follows:

Number of particles deposited per unit time and area at 
$$y_j$$
 
$$= \frac{\Omega_j \sum_{k=1}^{m_j} \Delta y_j, k}{L_t}$$
 
$$= \frac{\Gamma_j \sum_{k=1}^{m_j} \Delta y_j, k}{L_t}$$
 (65) Weight per particle at  $y_j$  and 
$$= \frac{\gamma_s (1-\theta)}{\Omega_j}$$
, and which is deposited at elevation  $y_j$  
$$= \hat{\mathbf{E}}[X \setminus Y_D = y_j]$$
,

where  $L_t$  is the total length of  $y_x(t)$  record, and all other symbols have been defined previously. Summing the product of the terms in equation 65 over all elevations,

$$\hat{Q}_{T}^{"} = \frac{\gamma_{s}(1-\theta)W}{L_{t}} \sum_{j=1}^{n} \left\{ \hat{E}[X \setminus Y_{D} = y_{j}] \sum_{k=1}^{m_{j}} \Delta y_{j,k}^{+} \right\}, (66)$$

where

 $\hat{Q}_T^{"}$  = estimate of mean total bed-material discharge;

W =width of channel.

Equation 66 also illustrates the contribution of each elevation to the total transport, but its primary distinction is that the transport rate is computed from the sounding records with a minimum number of computations.

The relationship between the three transport equations, 55, 63, and 66, will now be demonstrated. First, the comparison of equations 55 and 66 is demonstrated. Combining equations 22, 25, and 66,

$$\begin{split} \hat{Q}_{T}^{"} &= \frac{\gamma_{s}(1-\theta)W}{L_{t}} \sum_{j=1}^{n} \left\{ \hat{\mathbb{E}}[X \backslash Y_{D} = y_{j}] p_{Y_{D}}(y_{j}) \frac{N_{d}}{\Omega} \right\} \\ &= \frac{\gamma_{s}(1-\theta)WN_{d}}{\Omega L_{t}} \hat{\mathbb{E}}[X] , (67) \end{split}$$

Multiplying and dividing by the marginal rest period and utilizing equations 20, 21, 25, 30, and 31,

$$\hat{Q}_{T}^{"} = \frac{\gamma_{s}^{(1-\theta)WN}_{d}}{\Omega L_{t}} \quad \frac{\hat{\mathbf{E}}[X]}{\hat{\mathbf{E}}[T]}$$

$$\sum_{j=1}^{n} \left\{ \left( \frac{1}{m_{jj}} \sum_{k=1}^{m_{jj}} t_{j,k} \right) \left( \frac{\Omega}{N_{d}} \sum_{k=1}^{m_{j}} \Delta_{y_{j,k}}^{+} \right) \right\}. \quad (68)$$

As the value of  $\Delta y$  decreases the value of the last term can be approximated without appreciable error, as

$$\sum_{k=1}^{m_j} \Delta y_{j,k}^{\dagger} \stackrel{\text{\tiny $\alpha$}}{=} m_j \Delta y_j \quad . \quad . \tag{69}$$

Strictly speaking  $m_j \Delta y_j$  is equal to or slightly greater than the term

$$\sum_{k=1}^{m_j} \Delta y_{j,k}^+$$

(fig. 1). Assuming a long record with small vertical class intervals such that equation 69 is valid and such that  $m_i = m_{ii}$ , equation 68 reduces to

$$\hat{Q}_{T}^{"} = \gamma_{s} (1 - \Theta) W \frac{\hat{E}[X]}{\hat{E}[T]} \sum_{j=1}^{n} \frac{\Delta y_{j}}{L_{t}} \sum_{k=1}^{m_{j}} t_{j,k} \dots, (70)$$

which would be equivalent to equation 55 provided that the average depth of the zone in which bed-material movement occurs, h' is defined by

$$h' = \frac{1}{L_t} \sum_{j=1}^{n} \left( \Delta y_j \sum_{k=1}^{m_j} t_{j,k} \right) \dots$$
 (71)

Equation 71, is similar to equation 59 except that it is based on the time record of depth,  $y_x(t)$ , while equation 59 is based on the longitudinal profile,  $y_t(x)$ .

To investigate the relation between equations 63 and 66, we proceed as follows. The last term of equation 66 can be approximated without an appreciable error, as  $m_j \Delta y_j$  using equation 69. Replacing the last term by its approximate value, and multiplying equation 66 by the

right-hand side of equation 30 while dividing by the left-hand side,

$$\hat{Q}_T^{"} = \gamma_s (1 - \Theta) W$$

$$\sum_{j=1}^{n} \left\{ \frac{\hat{\mathbb{E}}[X \setminus Y_D = y_j]}{\hat{\mathbb{E}}[T \setminus Y_D = y_j]} \frac{m_j}{m_{jj}} \frac{\Delta y_j}{L_t} \right\} \sum_{k=1}^{m_{jj}} t_{j,k} \quad . \quad . \quad (72)$$

The number of bed forms contained in the  $y_x(t)$  record and which also contains some deposition in the class interval,  $m_j$ , should be almost equal to the total number of bed forms with both an upcrossing and a downcrossing in the interval,  $m_{jj}$ . Assuming  $m_j = m_{jj}$ , equation 72 is identical to equation 63 except that the percentage of the volume in the class interval occupied by dunes is computed from the  $y_x(t)$  record instead of from equation 62 which is based on the  $y_t(x)$  record.

In summary, three transport equations have been presented, equations 55, 63, and 66. Although the equations appear quite different in form, they are all based on similar assumptions. As the record length becomes long and the class intervals reduce to zero in the limit, the three equations would become identical provided that either the  $y_x(t)$  record or the  $y_t(x)$  record could be used to determine the active depth. (This should be true for equilibrium flow.) In the following section, the total load will be computed for three flow conditions using all three equations, and the results will be compared.

# GENERAL TWO-DIMENSIONAL STOCHASTIC MODEL FOR DISPERSION OF BED-MATERIAL SEDIMENT PARTICLES

Let us define the following stochastic processes:

$$\widetilde{X}(t) = \sum_{i=0}^{N(t)} X_i$$
 = longitudinal position of a bed-material sediment particle at time  $t$  in which  $\widetilde{X}(0) = X_0 = 0$ .

N(t) = counting process describing number of steps taken by a bed-material sediment particle in time t.

 $X_i$  = length of *i*th step of a bed-material sediment particle.

$$X(n) = \sum_{i=1}^{n} X_i$$
 = longitudinal position of a bed-  
material sediment particle after  $n$  steps.

 $\widetilde{Y}(t)$  = vertical position of a bed-material sediment particle at time t.

 $Y_D(n)$  = elevation at which a bed-material sediment particle is deposited after n steps.

The probability that the particle has, at time t, traveled a distance equal to or less than x and that it is located at an elevation equal to or less than y may now be expressed as the joint distribution function

$$F(x,y;t) = P[\tilde{X}(t) \leq x, \tilde{Y}(t) \leq y]$$

$$= \sum_{n=0}^{\infty} P[\tilde{X}(t) \leq x, \tilde{Y}(t) \leq y, N(t) = n] . (73)$$

Using the definition of conditional probability and assuming that the duration of the particle movement is negligible, equation 73 can be restated as

$$F(x, y; t) = \sum_{n=0}^{\infty} P[X(n) \le x, Y_D(n) \le y, N(t) = n]$$

$$= \sum_{n=0}^{\infty} P[X(n) \le x, N(t) = n \setminus Y_D(n) \le y] P[Y_D(n) \le y]$$

$$= \int_{y_{\min}}^{y} \sum_{n=0}^{\infty} P[X(n) \le x, N(t)]$$

$$= n \setminus Y_D(n) = y^{\dagger} f_{Y_D(n)}(y^{\dagger}) dy^{\dagger}$$
(74)

where

 $y_{\min}$  = lowest elevation of deposition; and  $f_{Y_D(n)}(y)$  = probability density function of  $Y_D(n)$ .

The event,  $\{N(t) = n\}$ , can be expressed in terms of the rest period of a bed-material sediment particle

$$\{N(t) = n\} = \{T(n) < t\} \bigcap \{T(n+1) > t\}$$
 (75)

where

 $\{ \cdot \} = \text{events};$ 

 $\bigcap$  = intersection of events;

$$T(n) = \sum_{i=1}^{n} T_i$$
; and

 $T_i$  = random variable describing the duration of ith rest period of a bed-material sediment particle.

By virtue of equation 75, it follows that

$$P[N(t) = n] = P[T(n) \le t, T(n+1) > t]$$

$$= P[T(n) \le t, T_{n+1} > t - T(n)] . (76)$$

For further simplification of equation 74, the following assumptions are made: (1) X(n) and N(t) are mutually independent for every n. (2)  $X_i$  for  $i \ge 1$  are independently and identically distributed according to the probability density function  $f_X(x)$ , where

 $0 \le x < \infty$ . Outside this range,  $f_X(x) = 0$ . (3)  $X_i$  is independent of  $Y_D(j)$  for  $i \neq j$ . (4)  $Y_D(i)$  for  $i \geq 1$  are independently and identically distributed according to the probability density function  $f_{Y_D}(y)$ , where  $y_{\min} \leq y \leq y_{\max}$ . Outside this range,  $f_{Y_D}(y) = 0$ . (5)  $T_i$  for  $i \geq 1$  are independently and identically distributed according to the probability density function  $f_T(t)$ , where  $0 \le t < \infty$ . Outside this range,  $f_T(t) = 0$ . (6)  $T_i$  is independent of  $Y_D(j-1)$  for  $i \neq j$ . In other words, assumption 1 states that the total distance X(n) traveled by a sediment particle after n steps should not depend on which time interval within [0,t] that these n steps occurred. The step lengths are always positive so that the particle always moves in downstream direction (part of assumption 2). Each step length depends on the elevation at which the particle is deposited at the end of that step (assumption 3). The elevation at which the particle is deposited at the end of any step does not depend on the elevation at which it was deposited at the end of any previous step (assumption 4). Finally, the duration of each rest period depends on the elevation at which the particle was deposited at the end of the previous step (assumption

Utilizing assumption 1, equation 74 becomes

$$F(x, y; t) = \int_{y_{\min}}^{y} \sum_{n=0}^{\infty} \{P[X(n) \leq x \setminus Y_{D}(n) = y^{t}] \}$$

$$P[N(t) = n \setminus Y_{D}(n) = y^{t}] f_{Y_{D}(n)}(y^{t}) \} dy^{t}$$

$$= \int_{y_{\min}}^{y} P[X(0) \leq x \setminus Y_{D}(0) = y^{t}]$$

$$P[N(t) = 0 \setminus Y_{D}(0) = y^{t}] f_{Y_{D}(0)}(y^{t}) dy^{t}$$

$$+ \int_{y_{\min}}^{y} \sum_{n=1}^{\infty} \{P[X(n) \leq x \setminus Y_{D}(n) = y^{t}] \}$$

$$P[N(t) = n \setminus Y_{D}(n) = y^{t}] f_{Y_{D}(n)}(y^{t}) \} dy^{t} . (77)$$

Under assumptions 2, 3, and 4, and using the concepts of joint and conditional probability,

$$\begin{split} P[X(n) &\leq x \backslash Y_D(n) = y^i] \\ &= P[X(n-1) + X_n \leq x \backslash Y_D(n) = y^i] \\ &= \int_0^x f_{X(n-1)} + X_n \backslash Y_D(n) \xrightarrow{(x^i \backslash y^i)} dx^i \\ &= \int_0^x dx^i \int_0^{x^i} f_{X(n-1)} , X_n \backslash Y_D(n) \xrightarrow{(\zeta, x^i - \zeta \backslash y^i)} d\zeta \end{split}$$

and using assumptions 3 and 4,

$$P[X(n) \leq x \backslash Y_{D}(n) = y^{\dagger}]$$

$$= \int_{0}^{x} dx^{\dagger} \int_{0}^{x^{\dagger}} f_{X(n-1)}(\zeta) f_{X \backslash Y_{D}}(x^{\dagger} - \zeta \backslash y^{\dagger}) d\zeta$$

$$= \int_{0}^{x} dx^{\dagger} \int_{0}^{x^{\dagger}} f_{X}(\zeta) f_{X \backslash Y_{D}}(x^{\dagger} - \zeta \backslash y^{\dagger}) d\zeta \qquad (78)$$

in which

$$X(n-1) = \sum_{i=0}^{n-1} X_i$$
,  $X_0 = 0$ ,

$$f_{X(n-1)}(\zeta) = f_{X}^{(n-1)}(\zeta)$$

$$= \int_{0}^{\zeta} f_{X}(\theta) f_{X}(\zeta - \theta) d\theta ; \quad n = 3, 4, 5, ...$$

and

$$f_X(\zeta) = f_X(x)$$
;  $n = 2$  (79)

In equations 78 and 79,  $f_{X(n-1), X_n \setminus Y_D(n)}(\zeta, x' - \zeta \setminus y')$  denotes the joint probability density function of X(n-1) and  $X_n$ , conditioned on  $Y_D(n)$ ,  $f_X(\zeta)$  is the (n-1)-fold convolution of the probability density function for the length of a single step, and it is equal to the probability density function for the distance traveled by the particle in (n-1) steps, and  $f_{X \setminus Y_D}(x \setminus y)$  is the conditional probability density function for a single step length given that the particle is deposited at elevation y.

Turning now to the other part of equation 77 and using equation 76 and assumptions 4, 5, and 6,

$$P[N(t) = n \setminus Y_{D}(n) = y']$$

$$= P[T(n) \leq t, T_{n+1} > t - T(n) \setminus Y_{D}(n) = y']$$

$$= \int_{0}^{t} \int_{t-t'}^{\infty} f_{T(n)}, T_{n+1} \setminus Y_{D}(n)^{-(t', \tau \setminus y')} d\tau dt'$$

$$= \int_{0}^{t} \int_{t-t'}^{\infty} f_{T(n)}^{-(t')} f_{T_{n+1}} \setminus Y_{D}^{-(n)}^{-(\tau \setminus y')} d\tau dt'$$

$$= \int_{0}^{t} f_{T}^{-(t')} dt' \int_{t-t'}^{\infty} f_{T \setminus Y_{D}}^{-(\tau \setminus y')} d\tau$$
(80)

in which

$$T(n) = \sum_{i=1}^{n} T_{i}$$

$$f_{T(n)}(t') = f_{T}(t') = \int_{0}^{t'} f_{T}(\theta) f_{T}(t' - \theta) d\theta ;$$

$$n = 2, 3, 4, \dots ,$$
and
$$(n)$$

$$(81)$$

In the above,  $T_{n+1}$  is the random variable describing the duration of the (n+1)th rest period of a particle,  $f_{T(n),T_{n+1}\setminus Y_D(n)}(t',\tau\setminus y')$  is the joint probability density function of T(n) and  $T_{n+1}$ , conditioned on  $Y_D(n)$ ,  $f_{T\setminus Y_D}(t\setminus y)$  is the conditional probability density function for the duration of a rest period given that the par-

ticle was deposited at elevation y, and  $f_T(t)$  is the n-fold convolution of the probability density function,  $f_T(t)$ , for the duration of a single rest period and is equal to the probability density function for the duration of n successive rest periods.

Similarly, the terms for n = 0 in equation 77 become:

$$P[X(0) \le x \setminus Y_D(0) = y'] = 1$$
 (82)

because  $X(0) = X_0 = 0$  and  $0 \le x < \infty$ , and

$$P[N(t) = 0 \setminus Y_D(0) = y^*] = P[T_1 > t \setminus Y_D(0) = y^*]$$

$$= \int_{t}^{\infty} f_{T\backslash Y_{D}}(t'\backslash y') dt' \quad , \quad (83)$$

where  $T_1$  is the random variable describing the duration of the first rest period in time t. It is important to note that the initial condition,  $X(0) = X_0 = 0$ , implies that the particle starts its first rest period at t = 0.

Introducing equations 78, 80, 82, and 83 into equation 77,

$$F(x,y;t) = \int_{y_{\min}}^{y} f_{Y_D}(y^i) dy^i \int_{t}^{\infty} f_{T \setminus Y_D}(t^i \setminus y^i) dt^i$$

$$+\int_{y_{\min}}^{y} f_{Y_{D}}(y') dy' \sum_{n=1}^{\infty} \left( \int_{0}^{x} dx' \int_{0}^{x'} f_{X}(\zeta)' \right)$$

$$f_{X \setminus Y_{D}}(x' - \zeta \setminus y') d\zeta \cdot \int_{0}^{t} f_{T}(t') dt' \int_{t-t'}^{\infty} f_{T \setminus Y_{D}}(\tau \setminus y') d\tau \bigg). (84)$$

The first term of equation 84 represents the joint probability that the particle has not moved from its initial position and that its initial elevation is equal to or less than y, and it is not a function of x. Hence,

$$\frac{\partial^{2}}{\partial x \partial y} \left[ F(x, y; t) \right]_{n = 0} = \frac{\partial^{2}}{\partial x \partial y} \int_{y_{\min}}^{y} f_{Y_{D}}(y') dy'$$

$$\int_{t}^{\infty} f_{T \setminus Y_{D}}(t' \setminus y') dt' = 0 .$$

The corresponding density function is therefore

$$f(x,y;t) = \frac{\partial^{2}}{\partial x \partial y} F(x,y;t)$$

$$= f_{Y_{D}}(y) \sum_{n=1}^{\infty} \left( \int_{0}^{x} f_{X}(\zeta)^{n-1} f_{X \setminus Y_{D}}(x - \zeta \setminus y) d\zeta \right)$$

$$\cdot \int_{0}^{t} f_{T}(t^{1}) dt^{1} \int_{t-t^{1}}^{\infty} f_{T \setminus Y_{D}}(\tau \setminus y) d\tau \right) . (85)$$

If a large number of identical particles are initially at rest at x = 0,  $y = y_0$ , equation 85 expresses the longitudinal and vertical distribution at time t of the particles which have moved from their respective initial positions. It should be noted here that f(x, y; t) is not a true probability density function because

$$\int_{0}^{\infty} dx \int_{y_{\min}}^{y_{\max}} f(x,y;t) dy = 1 - P[N(t) = 0] < 1 . (86)$$

That is to say, equation 85 applies only after the particle has moved from its initial position. The expression f(x, y; t) does not exist for x = 0.

If we assume that  $X_i$  is independent of  $Y_D(j)$  for all i and j and drop assumption 3, equation 85 reduces to equation 8,

f(x, y; t)

$$= f_{Y_{D}}(y) \sum_{n=1}^{\infty} f_{X}(x) \int_{0}^{t} f_{T}(t') dt' \int_{t-t'}^{\infty} f_{T \setminus Y_{D}}(\tau \setminus y) d\tau, \quad (8)$$

which is the Sayre-Conover (1967) two-dimensional stochastic model. The difference between equations 85 and 8 is that equation 85 takes some of the dependence between X and  $Y_D$  into account whereas equation 8 is based on the independence of X and  $Y_D$ . Hence, the Sayre-Conover model is a special case of equation 85.

The marginal case of equation 85 gives the longitudinal distribution at time t of the particles which have moved from their initial positions. Integrating equation 85 over y,

$$f(x;t) = \int_{y_{\min}}^{y_{\max}} f(x,y;t) \, dy = \sum_{n=1}^{\infty} \int_{0}^{x} \frac{(n-1)}{f_X(\zeta)} \, d\zeta \int_{0}^{t} f_T(t') \, dt'$$

$$\int_{t-t'}^{\infty} d\tau \left[ \int_{y_{\min}}^{y_{\max}} (x - \zeta \backslash y') f_{T \backslash Y_D}(\tau \backslash y') f_{Y_D}(y') dy' \right]. (87)$$

By virtue of assumption 1,

$$\int_{y_{\min}}^{y_{\max}} f_{X \setminus Y_D}(x - \zeta \setminus y') f_{T \setminus Y_D}(\tau \setminus y') f_{Y_D}(y') dy'$$

$$= \int_{y_{\min}}^{y_{\max}} f_{X, T \setminus Y_D}(x - \zeta, \tau \setminus y') f_{Y_D}(y') dy'$$

$$= f_{X, T}(x - \zeta, \tau) = f_X(x - \zeta) f_T(\tau) . \tag{88}$$

Substituting equation 88 into equation 87 and rearranging terms,

$$f(x;t) = \sum_{n=1}^{\infty} \int_{0}^{x} \frac{f_{X}(n-1)}{f_{X}(\zeta)} f_{X}(x-\zeta) d\zeta$$

$$\int_{0}^{t} f_{T}(t^{i}) dt^{i} \int_{t-t^{i}}^{\infty} f_{T}(\tau) d\tau .$$

Because

$$\int_{0}^{x} f_{X}(\zeta) f_{X}(x-\zeta) d\zeta = f_{X}(x) ,$$

and from equation 76,

$$P[N(t) = n] = \int_{0}^{t} f_{T}(t') dt' \int_{t-t'}^{\infty} f_{T}(\tau) d\tau$$
 (89)

Meanwhile, equation 75 can be restated as

$$\{N(t) = n\} = \{T(n) \le t\} \bigcap \{T(n+1) > t\}$$

$$= \{T(n) \le t\} - \{T(n+1) > t\}^{C}$$

$$= \{T(n) \le t\} - \{T(n+1) \le t\}$$

where  $\{T(n+1) > t\}^c$  denotes the complement of the

event  $\{T(n+1) > t\}$ . Because  $\{T(n+1) \le t\}$  is a subevent of  $\{T(n) \le t\}$ , it follows that

$$P[N(t) = n] = P[T(n) \le t] - P[T(n+1) \le t]$$

$$= \int_{0}^{t} f_{T}^{(n)}(t') dt' - \int_{0}^{t} f_{T}^{(n+1)} dt' . (90)$$

From equations 89 and 90, we have the marginal probability density function,

$$f(x;t) = \sum_{n=1}^{\infty} f_X^{(n)} P[N(t) = n]$$

$$\frac{\infty}{2} (n) \int_{0}^{\infty} f(n) (n+1)$$

$$= \sum_{n=1}^{\infty} f_X(x) \int_0^t \left[ f_T(t') - f_T(t') \right] dt' . (91)$$

Equation 91 is identical to the Sayre-Conover (1967) one-dimensional stochastic model which is given in equation 10. As with equation 85, here also f(x;t) is not a true probability density function because

$$\int_{0}^{\infty} f(x;t) dx = 1 - P[N(t) = 0] < 1 ,$$

where P[N(t) = 0] is the probability that the particle has not moved from its initial position.

In order to apply equations 84 or 85, the probability density functions  $f_{Y_D}(y)$ ,  $f_{T \setminus Y_D}(t \setminus y)$ , and  $f_{X \setminus Y_D}(x \setminus y)$  must be known. These density functions are estimated from equations 25, 28, and 39. The probability density functions  $f_T(t)$  and  $f_X(x)$  are determined by the relations

$$f_{T}(t) = \int_{y_{\min}}^{y_{\max}} f_{T \setminus Y_{D}}(t \setminus y) f_{Y_{D}}(y) dy$$
 (92)

and

$$f_X(x) = \int_{y_{\min}}^{y_{\max}} f_{X \setminus Y_D}(x \setminus y) f_{Y_D}(y) dy$$
 (93)

where  $y_{\min}$  and  $y_{\max}$  are estimated from the  $y_x(t)$  record. Equations 92 and 93 are the continuous forms corresponding to equations 29 and 40 (or 41), respectively. With  $f_T(t)$  and  $f_X(x)$  determined, the corresponding convolutions,  $f_T^{(n)}(t)$  and  $f_X(x)$  are determined from equations 81 and 79, respectively.

#### ANALYSIS AND DISCUSSION OF RESULTS

#### EXPERIMENT AND BASIC DATA

Three dune runs were made in a tilting recirculating flume of rectangular cross section, 61 m long, 2.4 m wide, and 1.2 m deep. The flume has been described in detail by Williams (1971).

The bed material used in these experiments was a screened river sand (Cherry Creek sand), with a median sieve diameter,  $d_{50}=1.13$  mm, and a geometric standard deviation,  $\sigma_g=1.51$ . The size distribution, shown in figure 5, was obtained by a sieve analysis of 3,000 grams of bed material.

After an equilibrium flow, as defined by Simons and Richardson (1966), was established, the  $y_x(t)$  and  $y_t(x)$  records, the total bed-material discharge, and the hydraulic properties of interest were measured. The methods and procedures of the measurements have been described in detail by Lee (1969). The summary of measured and derived data is given in table 1. The values of the water discharge, depth, energy slope, bed shape, and total bed-material concentration presented in table 1 are the average of several individual measurements. The sampled load was measured by a DH-48 sampler. The number of measurements was the same as the number of  $y_t(x)$  charts.

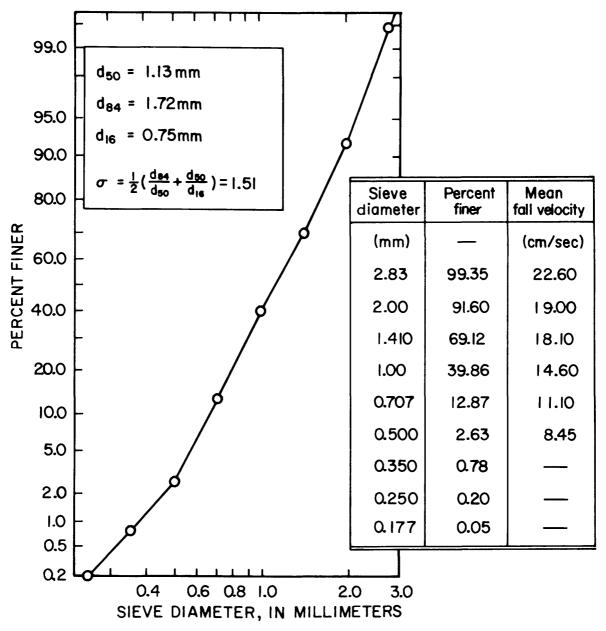


FIGURE 5. — Size distribution curve of bed material.

TABLE 1 -	Rasic data	and computed	narameters

Run $Water discharge \ Q_W, m^3/s \ Mean Standard deviation$			Flow depth $d$ , cm			Energy slope $S_{m{e}}$		Water	Water temperature $T$ , °C	
				Standard		an <u>fl</u> ow velocity $\overline{U}$ , cm/s	Mean	Standard deviation	Mean	Standard deviation
4A	0.464	.003	31.1	0.6		61.3	0.00167	0.00009	20.0	0.3
16	1.24	.006	90.8	0.9		55.8	0.00029	0.00021	22.8	0.2
17	1.53	.008	89.3	0.6		70.4	0.00047	0.00005	22.0	0.2
Total Bed-material discharge										
Run		itration mg/L	Mean total load q <sub>T</sub> , t/day•m	conce	oled load entration ng/L		Mean bed shear stres $\overline{\tau}_b$ , d/cm <sup>2</sup>	s sh <u>ea</u> r vel	ocity	Stream_power $\tau_b U$ , d/cm·s
	Mean	Standard deviation	41, .,	Mean	Standar deviation	rd Cul=	<i>g</i> , a	0.,		
4A	168.6	53.6	2.77	1.5	6.1	8.6	50.8	7.13	3	3110
16	8.9	3.1	0.39	0.0	0.0	12.1	25.9	4.91	_	1450
17	29.7	10.8	1.61	6.3	10.2	11.0	41.2	6.43	}	2900
$y_r(t)$ Record					$y_t$	(x) Record				
Run	Mean Froude nui F <sub>r</sub>		n of record , hours	Lag inte min	rval	Length of record $L_x$ , m	Number of charts	Time inter measurer hour	nents	Range of lag interval cm
4A	0.35		312	2.4		1235	31	6		3.7 - 13.0
16	0.19		80	1.2		1006	33	1		8.6 - 13.6
17	0.24		109	0.6		983	30	1		7.2 - 11.8

The  $y_t(x)$  charts were obtained by mounting a sonic depth sounder on the instrument carriage and traversing it along the centerline of the flume in the upstream direction. The sonic depth sounder has been described by Karaki, Gray, and Collins (1961). Although the duration of traverse was approximately 5 minutes, the  $y_t(x)$  record was considered to be instantaneous. The  $y_{x}(t)$  record was obtained by locating a sonic depth sounder at the centerline of the flume 42.1 m downstream of the headbox. Both the  $y_t(x)$  and  $y_r(t)$ records were digitized with an analog-to-digital converter at the lag intervals shown in table 1. The lag interval on the  $y_t(x)$  charts were not constant because the speed of the carriage was somewhat different for each chart. The output of the converter was to computer cards so that all statistics could be processed on the digital computer.

## PROBABILITY DISTRIBUTIONS OF THE ELEVATIONS OF DEPOSITION AND EROSION

The sample probability mass functions for the elevation of deposition and erosion were computed using equations 25 and 26 and the  $y_x(t)$  records. The results of calculations for the three flume runs are presented in table 2.

The  $y_x(t)$  record of each run was standardized so that the class mark,  $y_i$ , measures the elevation of deposition or erosion in terms of the standard deviation about the mean bed elevation. The class width of 0.4 was used for all class marks. The frequency histograms for the elevation of deposition and erosion are plotted in figure 6.

The truncated Gaussian probability density function, defined by

$$f_{Y_{D}}(y) = f_{Y_{E}}(y)$$

$$= \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^{2}}}{\frac{1}{\sqrt{2\pi}} \int_{-2.4}^{2.4} e^{-\frac{1}{2}y^{2}} dy} \tilde{z} = 1.017 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^{2}}$$
for  $-2.4 \le y \le 2.4$ 

$$f_{Y_{D}}(y) = f_{Y_{E}}(y) = 0 \text{ otherwise } ,$$

appears to fit the data reasonably well. A symmetric triangular density function defined by

$$f_{Y_{D}}(y) = f_{Y_{E}}(y) = -\frac{1}{2.4^{2}}y + \frac{1}{2.4} \text{ for } 0 \le y \le 2.4$$

$$f_{Y_{D}}(y) = f_{Y_{E}}(y) = \frac{1}{2.4^{2}}y + \frac{1}{2.4} \text{ for } -2.4 \le y \le 0$$

$$f_{Y_{D}}(y) = f_{Y_{E}}(y) = 0 \text{ otherwise } ,$$

$$(95)$$

also appears to fit the data reasonably well. Equations 94 and 95 are both plotted in figure 6. In equations 94 and 95,  $f_{Y_D}(y)$  and  $f_{Y_E}(y)$  are the probability density functions of the elevation of deposition and erosion, respectively, and y is the standardized elevation.

Both distributions assume nonzero values only for  $-2.4 \le y \le 2.4$  and the two models postulate that  $Y_D$  and  $Y_E$  are identically distributed. The truncation limits on these distributions are rather arbitrary.

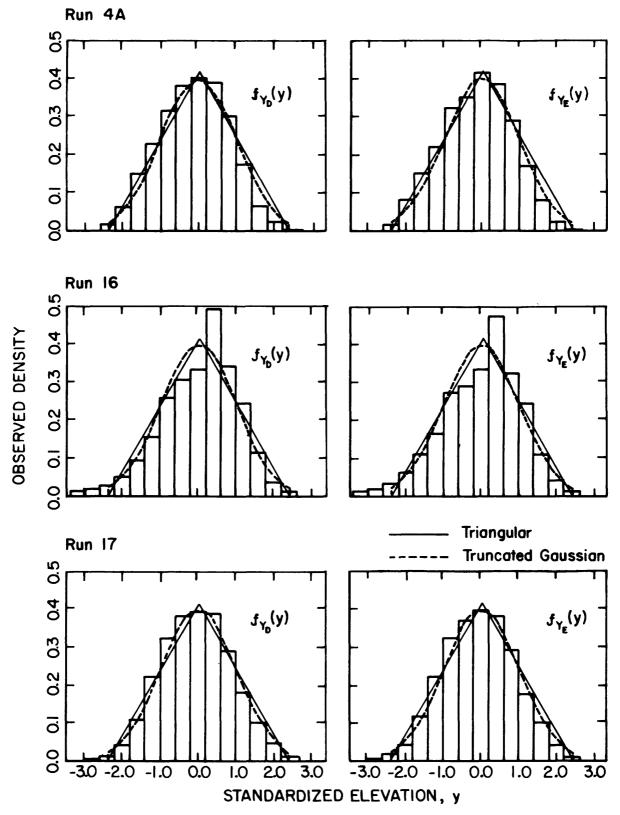


FIGURE 6. — Frequency histograms, triangular density function, and truncated Gaussian density function for the elevation of deposition and erosion.

The mean and variance for the truncated Gaussian density are

$$E[Y_D] = E[Y_E] = 0 ,$$
and
$$Var[Y_D] = Var[Y_E] = E[Y_D^2]$$

$$= E[Y_E^2] = 1.017 \int_{-2.4}^{2.4} y^2 g(y) dy = 0.891$$
(96)

where  $g(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2}$ . For the triangular density

$$E[Y_D] = E[Y_E] = 0$$
and
$$Var[Y_D] = Var[Y_E] = E[Y_D^2]$$

$$= E[Y_E^2] = \frac{2.4^2}{6} = 0.960$$

The variances of these distributions are quite sensitive to the assumed truncation limits.

A goodness of fit test using the chi-square statistic indicated that neither model would be rejected for runs 16 and 17 at a significance level of 0.05. For run 4A, however, both models were rejected at the same level of significance. As seen in figure 6, the truncated Gaussian density appears to give a slightly better approximation to  $f_{Y_D}(y)$  and  $f_{Y_E}(y)$ ; but, the triangular density is much easier to handle in analytical treatments. The variance of the triangular distribution is even more sensitive to the assumed limits than is the variance of the truncated Gaussian distribution. Therefore, the triangular distribution probably should not be used in predicting the variance.

For stationary processes, continuity requires that the probability of erosion equal the probability of deposition for all elevations. Therefore, the density functions for the elevations of deposition and erosion must be identical. The mean and variance of sample histograms as well as the total number of points available for analysis,  $\Sigma m_i$ , are shown in table 2. Little data were available for run 16, only 134 crossings compared to 2,167 for run 4A and 708 for run 17. Although run 16 was continued for 33 hours, the very low transport rate (table 1) and slow movement of the bed forms limited the number of crossings available for analysis. It should also be pointed out that equilibrium flow was never attained for this flow which was barely above the initiation of motion stage.

#### REST PERIOD DISTRIBUTIONS

The sample conditional probability mass function of the rest periods were computed by determining the difference betweeen the time of reexposure and movement and the time of burial of the center of each class mark for each crossing event,  $m_{j,j}$ , that occurred in the  $y_x(t)$  record (fig. 2). The results of the measurements are presented in tables 3 through 5, and examples of the mass functions are presented in figures 7, 8, and 9. The standardized  $y_x(t)$  record was used and the class width of the elevation was taken to be the same as that used in determining the probability distribution for the elevation of deposition, 0.4.

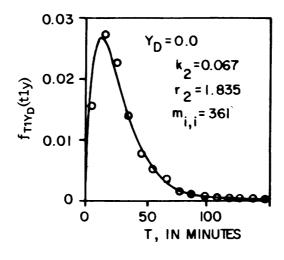
The mean and variance of the conditional rest periods were computed using equation 30, and the results are presented in table 6. These results are also plotted as a function of bed elevation in figure 10. As can be seen from figure 10, both the conditional mean and variance of the rest periods decrease with increasing elevation of deposition. Inspection of figure 2 indicates that the conditional mean should decrease with increasing elevation of deposition. However, the decrease of the variance is not so obvious. Because the mean value is decreasing with increasing elevation, the decrease in the variance is not too meaningful. The coefficient of variation (standard deviation/mean) is probably a better measure of the variability of the rest periods. Restricting our attention to runs 4A and 17, for reasons to be discussed later, the coefficient of variation remained roughly constant in the range of 0.6-0.75 for elevations above the mean bed elevation, and it increased with decreasing elevation to a value of about 1.5 at 2.4 standard deviations below the mean bed elevation. Thus the variability of the rest period, as measured relative to its mean, also decreases with increasing elevation at least up to the mean bed elevation.

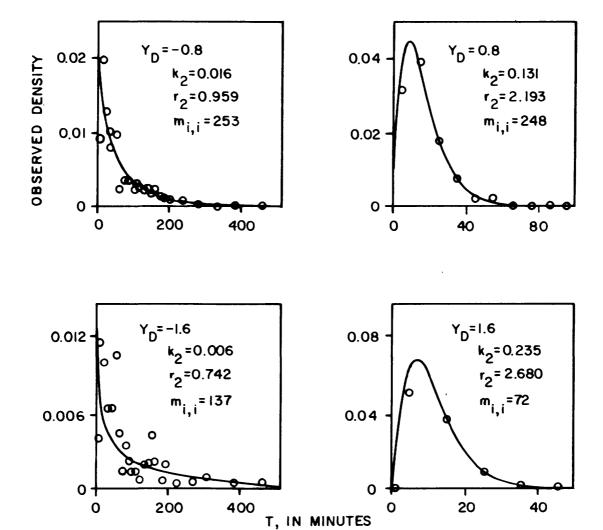
As seen from figure 10, both the mean and variance of the conditional rest periods may be approximated by an expression of the form,

$$\hat{E}[T \setminus Y_D = y] = Ae^{-By}$$
and
$$\hat{Var}[T \setminus Y_D = y] = Ce^{-Dy}$$
(98)

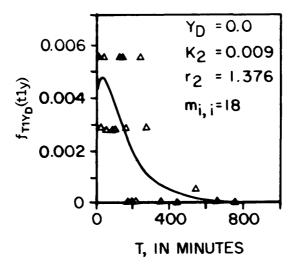
The constants A, B, C, and D in equation 98 were determined by a regression analysis of the data plotted in figure 10, and the resulting values are presented in the figure. The values A and C represent measures of the mean and variance of the rest period, respectively, for the mean bed elevation. The values of B and D are measures of the rate of change of the mean and variance of the rest period with bed elevation, respectively.

The distributions of the conditional rest periods were approximated by the two-parameter gamma probability





 $FIGURE 7. -- Sample \ probability \ mass \ functions \ of the \ conditional \ rest \ periods \ with \ fitted \ two-parameter \ gamma \ functions \ (run \ 4A).$ 



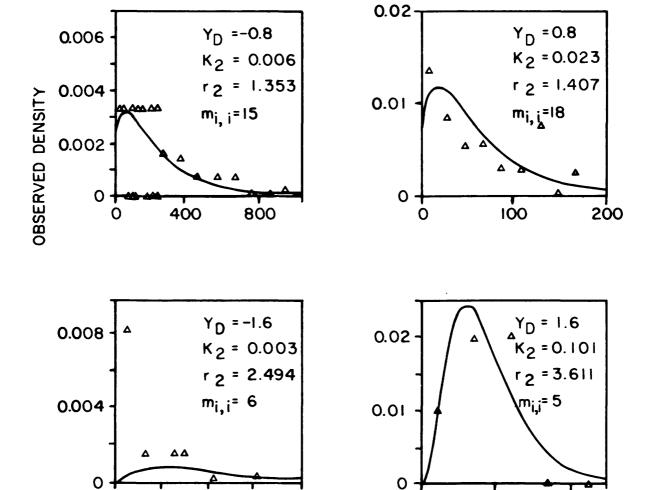


Figure 8. — Sample probability mass functions of the conditional rest periods with fitted two-parameter gamma functions (run 16).

T, IN MINUTES

0.02

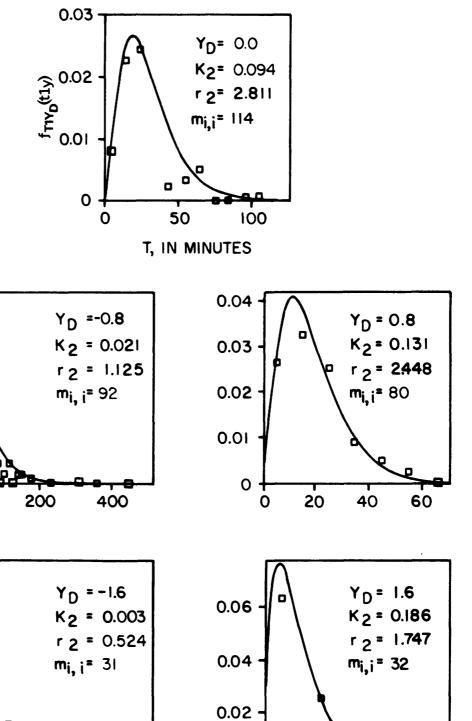
0.01

0.015

0.010

0.005

OBSERVED DENSITY



0

T, IN MINUTES

0

20

40

FIGURE 9. — Sample probability mass functions of the conditional rest periods with fitted two-parameter gamma functions (run 17).

1000

500

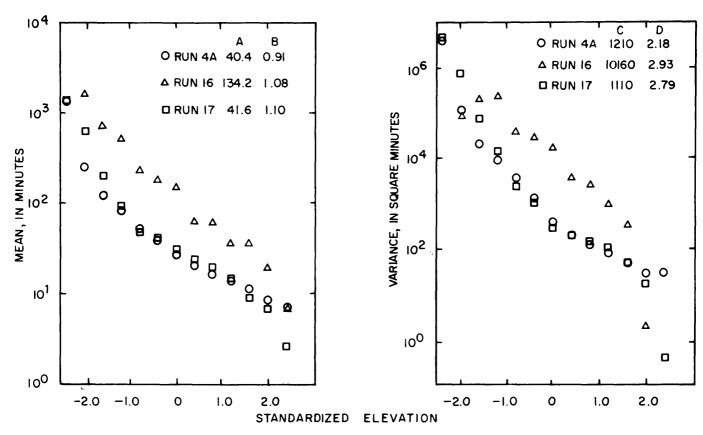


FIGURE 10. — Variation of the conditional mean and variance of rest periods with bed elevation.

density function which has the form,

$$f_{T \setminus Y_{D}}(t \setminus y) = \frac{k_{2,y}}{\Gamma(r_{2,y})} (k_{2,y}t)^{r_{2,y}-1} e^{-(k_{2,y})t}$$
(99)

where

 $\Gamma(\bullet)$  = gamma function; and  $k_{2.v}$ ,  $r_{2.v}$  = scale and shape parameters, respectively.

The scale and shape parameters were estimated by using the method of moments,

$$k_{2,y} = \frac{\hat{\mathbb{E}}[T \setminus Y_D = y]}{\hat{\mathbb{Var}}[T \setminus Y_D = y]}$$
 and 
$$r_{2,y} = \frac{(\hat{\mathbb{E}}[T \setminus Y_D = y])^2}{\hat{\mathbb{Var}}[T \setminus Y_D = y]} = k_{2,y}\hat{\mathbb{E}}[T \setminus Y_D = y]$$
 (100)

and the data contained in table 6. The variation of  $k_{2,\,y}$  and  $r_{2,\,y}$  with bed elevation are presented in table 7 along with the results of a chi-square goodness of fit test. The ability of the two-parameter gamma distribution to fit the measured mass functions is illustrated in figures 7, 8, and 9.

From table 7, as well as from figures 7, 8, and 9, both the scale and shape parameters increase with increasing bed elevation, with a few exceptions for the shape parameter. The shape of the conditional density of the rest periods (figs. 7, 8, 9) approaches a J-shape and becomes more peaked as bed elevation decreases. Therefore, the exponential density might fit better than the two-parameter gamma density below the mean bed elevation (y < 0). The exponential density function is a special case of the gamma density with  $r_{2,y} = 1$ . The better fit of the exponential density seems to be consistent with the fact that all rejections of the chi-square test (6 rejections out of 22 at a significant level of 0.05) occurred below the mean bed elevation (table 7). It would appear that the exponential form for the conditional rest period as proposed by Grigg (1969) is only valid for elevations below the mean bed elevation.

A major factor in determining the degree of fit between the measured density functions and the fitted curves in figures 7, 8, and 9 appears to be the number of points available from which the distribution was constructed. In general, if more than 100 points were available,  $m_{i,i}$ , the fit is pretty good. The weakness of the data for run 16 is very apparent. Even at the mean bed elevation, only 18 crossing events were observed.

Combining equations 98 and 100, the scale and shape parameters can be estimated using only the constants A, B, C, and D.

and 
$$k_{2,y} = \frac{Ae^{Dy}}{Ce^{By}}$$

$$r_{2,y} = \frac{A^{2}e^{Dy}}{Ce^{2By}} = k_{2,y}Ae^{-By}$$

The sample joint probability mass functions of the rest period and the elevation of deposition were computed from equation 33 using the results presented in tables 2 through 5. The results of these computations are presented in tables 8, 9, and 10. The correlation coefficients were computed by using equation 34, along with the data contained in tables 2, 6, 8, 9, and 10. The values of the correlation coefficients were -0.27, -0.53and -0.26 for runs 4A, 16, and 17, respectively. The rest period and the elevation of deposition are negatively correlated, but the degree of their linear association is not strong.

The sample marginal probability mass functions,  $p_T(t_a)$ , were computed by use of equation 29 and the data contained in tables 2, 3, 4, and 5. The results of these computations are also presented in tables 8, 9, and 10. The sample frequency histograms for the marginal rest periods are plotted in figure 11. The mean and variance of the marginal rest periods were computed by use of equation 31. These results are also presented in figure 11. The variance values appear to be extremely large. For example, the standard deviation for run 4A is almost four times the mean value. The computed variance values are extremely dependent on the long rest periods, the extreme events generally occur at low bed elevations. For example, by ignoring rest periods of greater than 2,000 minutes, which have a probability of occurrence of only 0.0015, the variance is reduced from 42,000 to 12,000.

Also shown in figure 11 are exponential density functions with a mean equal to the computed marginal mean. The exponential density function fits the data reasonably well; however, there would appear to be room for improvement. A gamma density fitted by the method of moments would be an extremely poor fit of the data. A gamma distribution, estimated by the maximum likelihood method may fit the data reasonably well.

The marginal distribution of the rest periods could also be estimated by

$$f_{T}(t) = \int_{-2.4}^{2.4} f_{T \setminus Y_{D}}(t \setminus y) f_{Y_{D}}(y) dy \qquad (102)$$

where  $f_{T \setminus Y_D}(t \setminus y)$  is the two-parameter gamma density (eq. 101) with parameters given by equation 101, and  $f_{Y_p}(y)$  is given by equation 94, or it could be obtained by fitting the frequency histograms contained in figure 11 with some assumed distribution.

#### STEP LENGTH DISTRIBUTIONS

The  $y_t(x)$  record was standardized after removing a straight line trend. The trend determined by the method of least squares accounted for the possibility that the sand bed in the flume was not parallel to the instrument carriage rails supporting the sonic sounder. In standardizing the  $y_t(x)$  record, the standard deviation obtained from the  $y_r(t)$  record was used. With these standardized data, the statistic  $\{x_{i,j,k}\}$  was analyzed (fig. 3) to estimate various probability distributions of the step lengths.

The sample probability mass functions given the elevations of deposition and erosion were computed by using equation 47, and the results are presented in tables 11 through 56. Examples of these mass functions are shown in figures 12, 13, and 14. The corresponding means and variances were estimated by equation 48 and summarized in tables 57, 58, and 59.

It can be seen from tables 57, 58, and 59, as well as in figures 12, 13, and 14, that the conditional mean of the step length decreases with an increase in either the elevation of deposition or of erosion. This result could be expected simply from the typical shape of the dunes. Likewise the conditional variance of the step length tends to decrease with an increase in the elevation of either deposition or erosion. The above statements essentially imply that longer step lengths are associated with lower elevations at which a sediment particle is eroded and deposited, and vice versa.

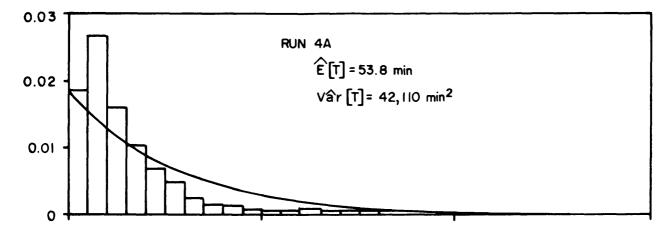
The distribution of conditional step lengths were approximated by the two parameter gamma probability density functions,

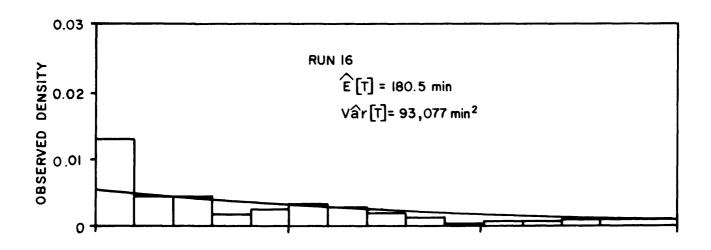
$$f_{X \setminus Y_{E}, Y_{D}}(x \setminus y, y') = \frac{k_{1}, y, y'}{\Gamma(r_{1}, y, y')} (x k_{1}, y, y')^{\left(-1 + r_{1}, y, y'\right)} e^{-(k_{1}, y, y') x}$$
(103)

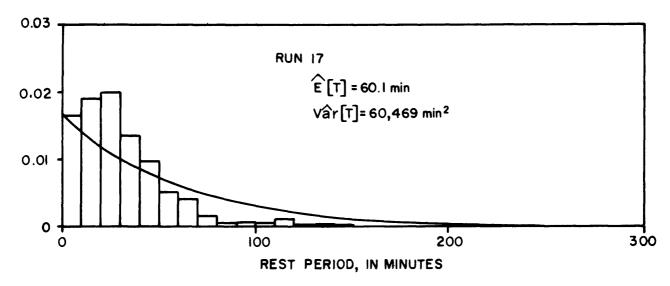
where

y and y' = arguments of  $Y_E$  and  $Y_D$ , respectively; and  $k_{1, y, y'}$  and  $r_{1, y, y'} =$ scale and shape parameters,

The parameters  $k_{1,\;\mathbf{y},\;\mathbf{y}'}$  and  $r_{1,\;\mathbf{y},\;\mathbf{y}'}$  were estimated by the







 $\textbf{Figure 11.} \label{eq:Figure 11.} \ \textbf{Frequency histograms for the marginal rest period and exponential fits}.$ 

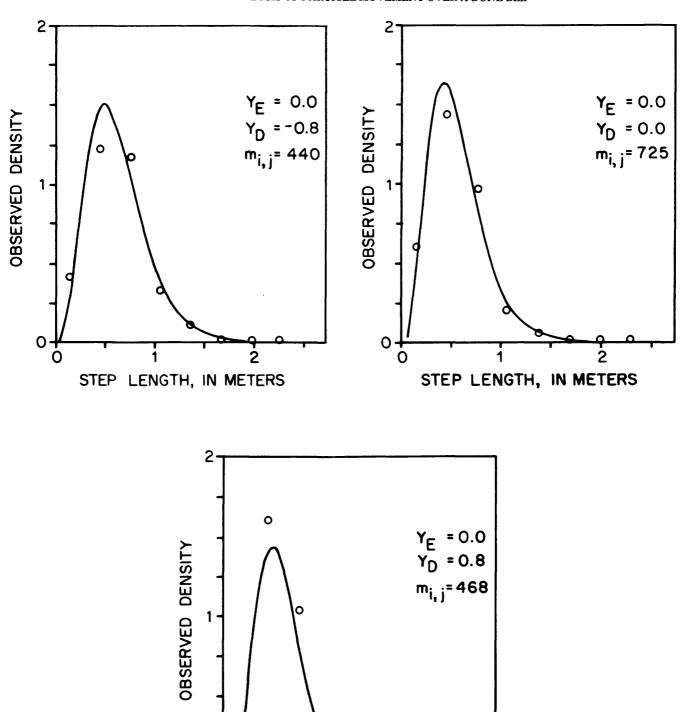
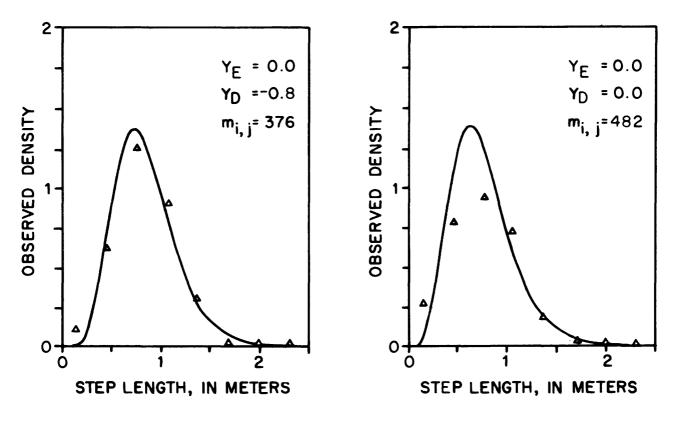


FIGURE 12. — Sample probability mass functions of the conditional step lengths given the elevation of erosion is 0.0 with Gamma fits (run 4A).

STEP LENGTH, IN METERS

0



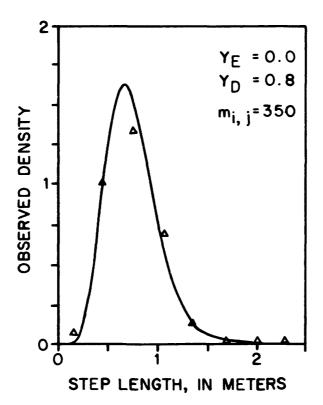


FIGURE 13. — Sample probability mass functions of the conditional step lengths given the elevation of erosion is 0.0 with Gamma fits (run 16).

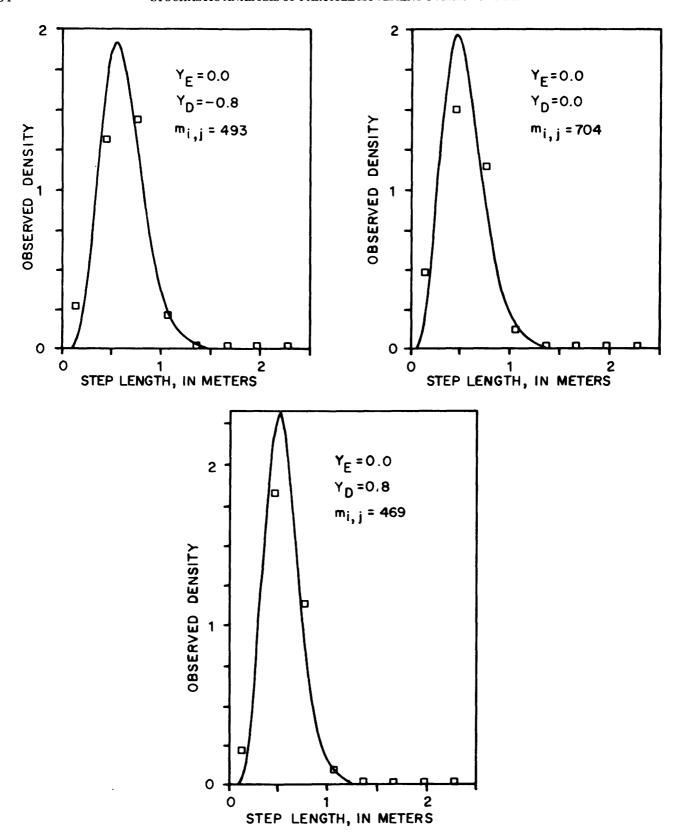


FIGURE 14. — Sample probability mass functions of the conditional step lengths given the elevation of erosion is 0.0 with Gamma fits (run 17).

method of moments, using data contained in tables 57, 58, and 59 and the expressions

$$k_{1,y,y'} = \frac{\hat{\mathbb{E}}[X \setminus Y_E = y, Y_D = y']}{\hat{\mathbb{Var}}[X \setminus Y_E = y, Y_D = y']}$$
and
$$r_{1,y,y'} = \frac{\left(\hat{\mathbb{E}}[X \setminus Y_E = y, Y_D = y']\right)^2}{\hat{\mathbb{Var}}[X \setminus Y_E = y, Y_D = y']}$$

$$= \hat{\mathbb{E}}[X \setminus Y_E = y, Y_D = y']k_{1,y,y'}$$

The variation of  $k_{1, y, y'}$  and  $r_{1, y, y'}$  with the elevations of erosion and deposition are shown in tables 60, 61, and 62. These approximations are also shown in figures 12, 13, and 14.

The chi-square test for goodness of fit was used to test these gamma approximations. The results of these tests are summarized in tables 63, 64, and 65. None of the 81 distributions tested could be rejected at the 0.05 level of significance. In other words, there is no good statistical reason to reject the hypothesis that the probability density functions for the step lengths, given the elevation of deposition and erosion, are distributed according to the two-parameter gamma distribution. The

fitted gamma distributions are also plotted and the example mass functions presented in figures 12, 13, and 14. These figures also help to illustrate the ability of the two-parameter gamma distributions to fit the measured conditional step length distributions.

The sample conditional mass functions, given the elevation of deposition, were computed based on equation 37 and the data contained in tables 2 and 11-56. These mass functions are presented in tables 66, 67, and 68. The corresponding conditional means and variances were computed using equation 38 and are presented in table 69 as well as being plotted in figure 15. Again, the general decrease in the expected value of the step length with an increase in the elevation of deposition is apparent.

The sample joint probability mass function of the step length and the elevation of deposition was computed by equation 43, and the results are shown in tables 70, 71, and 72. The correlation coefficients were computed by equation 44, and their values were -0.15, -0.15, and -0.20 for runs 4A, 16, and 17, respectively, indicating that the step length and the elevation of deposition are negatively correlated, but the degree of their linear associations is not strong. The sample marginal probability mass functions,  $p_X(x_\beta)$ , computed using equation 40, are also shown in tables 70, 71, and 72. The

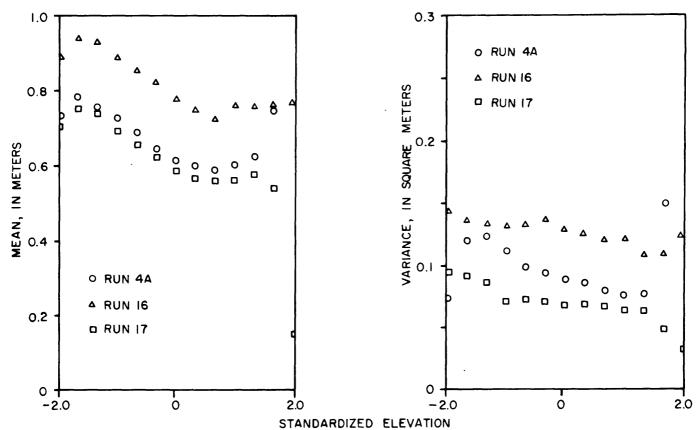


FIGURE 15. — Variation of the conditional mean and variance of step lengths with bed elevation;  $\hat{E}[X \setminus Y_D = y]$ .

sample frequency histograms for the marginal rest periods are plotted in figure 16. The mean and variance of the marginal rest periods were computed by use of equation 41. These results are also presented in figure 16. The range of the means is fairly small, only 0.610 to 0.799 m. The mean dune lengths, as measured by the distance between trough points, for the three runs were 1.19, 1.66, and 1.23 m respectively for runs 4A, 16, and 17. The mean step lengths were, therefore, 54, 48, and 49 percent of the mean dune lengths. Grigg (1969) found the mean step lengths of single tagged particles to be about 60 percent of the mean dune length. Of course, Grigg was working with a much finer sand, .33 to .45 mm, as compared to 1.15 mm for this study. Also shown in figure 16 are gamma density functions for which the parameters k and r were determined from the mean and variance shown in the figure. The gamma functions appear to fit the data very well for all three runs. The value of the parameter r ranged from 4.05 for run 4A to 4.59 for run 17. This is slightly more than twice the value estimated by Yang (1968) from the step length distribution of a single plastic particle.

#### **BED-MATERIAL TRANSPORT**

The following assumptions and conditions were used to estimate the mean total bed-material transport rate by equations 55, 63, and 66: (1) Because the bed material was coarse sand (fig. 5), all sediment particles are assumed to be transported as bed load. Expressed mathematically,  $P[E_1] = 1$ . (2)  $\gamma_s(1-\theta) = 1602 \text{ kg/m}^3$ . (3)  $\Delta y_j = 0.4s_y$  everywhere. By virtue of item 1, it follows that  $\hat{V}_T = \hat{V}_B$ ,  $\hat{V}_T(j) = \hat{V}_B(j)$ , and  $\hat{Q}_T = \hat{Q}_B$ . In item 3,  $s_y$  is the standard deviation of the bed elevation computed from the  $y_r(t)$  record.

All parameters and statistics which are required by equations 55, 63, and 66 are summarized in tables 73 and 74. The average depth of the zone of bed material movement, h, was determined by equation 59. It was found that one chart of the  $y_t(x)$  record (about 34 m) is sufficient to obtain a reliable value of h, although over 30 charts of the  $y_t(x)$  record were used in this study. Each chart contained about ten dunes. Using equation 62,  $\xi_i$ , the percentage of volume between elevations  $\eta_i$  and  $\eta_{i+1}$  occupied by dunes (hereafter will be referred to as the effective volume ratio) was obtained from the  $y_t(x)$  record. The results are presented in table 74 and plotted in figure 17. As shown in figure 17  $\xi_i$  is nearly independent of flow condition. As long as the bed forms are dunes,  $\xi_i$  does not change appreciably. It is also shown in figure 17 that  $\xi_j$ is nearly unity and zero at  $y_i = -2.4$  and  $y_i = +2.4$ , respectively. This is partial justification for the upper and lower limits of the elevations of erosion and deposition used in equations 94 and 95.

Another effective volume ratio can be obtained from the  $y_x(t)$  record. Denoting this ratio by  $\zeta_i$ ,

$$\zeta_{j} = \frac{1}{L_{t}} \sum_{k=1}^{m_{j}} t_{j,k}$$
 (105)

where

 $L_t = \text{total length of } y_x(t) \text{ record};$ 

 $m_j$  = maximum number of bed forms contained in the  $y_x(t)$  record which also contains some deposition at elevation  $y_i$ ; and

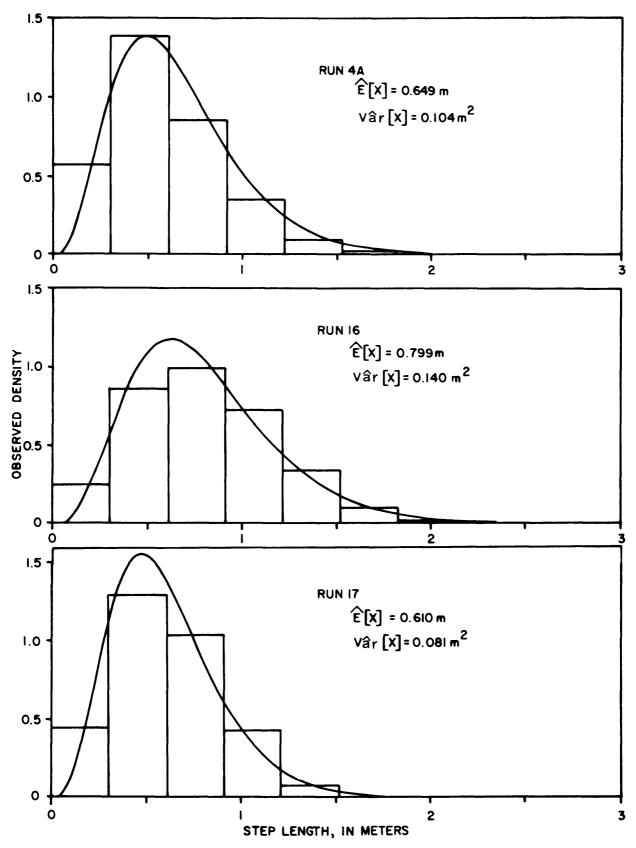
 $t_{i,k}$  = measurement of the conditional rest period.

There is no significant difference between  $\xi_j$  and  $\zeta_j$  (table 74) except for depths greater or less than 2.0 standard deviations from the mean. The longitudinal profiles  $(y_t(x))$  records appear to contain a larger number of extreme events than the time record at a given point (the  $y_x(t)$ ). The explanation for this is probably that the flow was fairly stationary but that it was not longitudinally uniform.

A comparison of measured and computed total bedmaterial transport rates is shown in table 75. It is seen that:

- 1. For run 4A, all three equations provide excellent estimates to the observed mean total bed-material discharges.
- 2. Equation 55 provided an excellent estimate to the mean total bed-material discharge for run 17. However, the other two equations overestimated the discharge by more than 25 percent. The reason for the differences in the equations is not understood.
- 3. None of the equations gave good estimates of the mean total bed-material discharge for run 16. The consistently overestimated discharge ranged from 64 percent for equation 63 to 80 percent for equation 55. It should be remembered, however, that the mean total bed-material discharge was less than 9 mg/L during this run, that the flow was not in equilibrium as illustrated by the large variation of energy slope (table 1), and that very few rest period statistics were available for analysis (fig. 8).

Taken as a whole, the results are very encouraging. Although equation 55 gave the most accurate results for run 17, it should be noted that equations 63 and 66 gave very consistent results for all runs when they are compared one with the other. The discharge predicted by equation 66 was 8.3, 8.4, and 8.5 percent larger than that predicted by equation 63 for runs 4A, 16, and 17, respectively. Although equation 66 is probably simpler to evaluate than equation 63, it appears that some accuracy has been sacrificed. The main difference between equations 63 and 66 is the way in which the effective depth or effective volume ratio (eq. 62, 71) is



 $\textbf{Figure 16.} \begin{tabular}{l}\textbf{Figure 16.} \end{tabular} \begin{tabular}{l}\textbf{Frequency histograms for the marginal step length with $Gamma$ fits.} \end{tabular}$ 

Table 73. — Variation of various statistics with stream power

Run	$\overline{\tau}_b \overline{U}$ (d/cm·s)	Ê[X] (m)	Vâr[X] (m <sup>2</sup> )	Ê[T] (mın)	Vâr[T] (min <sup>2</sup> )	$\hat{V}_T = \frac{\hat{E}[X]}{\hat{E}[T]}$ (cm/s)	$\hat{V}_{T}^{\dagger}$ (cm/s)	h (cm)	s (cm)
<b>4</b> A	3,110	0.649	0.104	53.8	42,110	0.0201	0.0378	9.66	4.26
16	1,450	0.799	0.140	180.5	93,077	0.0073	0.0166	6.89	3.01
17	2,900	0.610	0.081	60.1	60,469	0.0169	0.0372	7.13	3.61

Table 74. — Comparison of the effective volume ratios at elevation  $y_j$ ;  $\xi_j$ , from  $y_t(x)$  record and  $\zeta_j$ , from  $y_x(t)$  record

	Run 4A		Run	16	Run	17
<sup>Y</sup> D <sup>=y</sup> j	ξj	$^{\zeta}j$	$^{\xi}j$	ζj	ξj	$^{\zeta}j$
-2.8	0.968		1.000		1.000	
-2.4	.963		.995		.998	
-2.0	.949	0.903	.975		.989	
-1.6	.925	.883	.935	0.890	.964	0.943
-1.2	.877	.865	. 865	.840	.900	.867
-0.8	.791	.787	.778	.719	. 801	.756
-0.4	.668	.675	.670	.656	.669	.645
0.0	.512	.529	.535	.567	.522	.522
0.4	.344	.379	.385	.388	.364	.380
0.8	.189	.222	.230	.230	.202	.234
1.2	.078	.112	.103	.097	.079	.120
1.6	.015	.044	.036	.037	.020	.046
2.0	.001	.009	.012	.008	.002	.013
2.4	.000		.004	.001	.000	.001
2.8			.001		.000	

computed, and these functions were similar (table 74); therefore, the consistency of their final result was expected. Equation 55 had the lowest average absolute error for all three runs; however, equation 63 gave the most accurate result on two out of three runs. Because

of the similarity of equations 55 and 63, it would be difficult to say one was more accurate than the other. Their relative accuracy probably depends on chance occurrence of extreme events in one or the other records of bed elevation.

In table 75,  $\hat{q}_T$  is the mean total bed-material discharge in weight per width and time, and it was obtained by dividing equations 55, 63, and 66 by the width of the channel, W.

If we define  $q'_B(j)$  as the mean bed-load discharge associated with elevation  $y_i$ , then based on equation 63,

$$\hat{q}_{B}^{+}(j) = \gamma_{s} (1 - \Theta) \hat{V}_{B}(j) \xi_{j}^{\Delta} y_{j}$$
 (106)

where  $\hat{q}_B(j)$  estimates  $q_B'(j)$  and  $\hat{V}_B(j)$  is an estimate of the mean transport speed of a bed-load particle at elevation  $y_j$ . The mean transport speed,  $\hat{V}_B(j)$ , is given by equation 61 provided that the suspended load is negligible. With equation 106, the variation of bed-load discharge with bed elevation may be investigated. This variation is shown in figure 18 for all three runs. It is seen that the maximum bed-load discharge is associated with the mean bed elevation and that an insignificant portion of the bed-load movement appears to occur for  $y_j \leqslant -2.4$  and  $y_j \geqslant +2.4$ .

Table 75. — Comparison of measured and computed total bed-material transport rates

D		Measured to	tal bed-mater	nal discharge (t/c	lay·m)			
Run	Number of measurements	Ma	xımum	Minimum	Standard deviation			
4A	54	5	.30	0.91	0.88	2.77		
16	32	0	.72	0.18	0.14	0.40		
17	32	3	. 04	0.59	0.59	1.61		
	Computed m	ean, $\hat{q}_T$ (t/c	lay·m)	Compute	ed mean/measure	d mean		
Run	Eq. 55	Eq. 63	Eq. 66	Eq. 55	Eq. 63	Eq. 66		
4A	2.68	2.72	2.94	0.970	0.983	1.066		
16	0.70	0.64	0.67	1.801	1.641	1.725		
17	1,67	2.05	2.18	1.035	1.269	1.354		

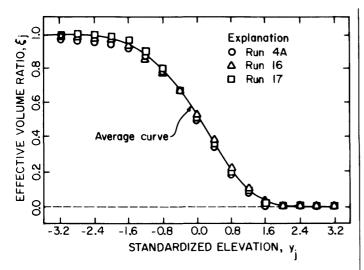


Figure 17. — Effective volume ratio as a function of bed elevation,  $y_j$ ;

$$\xi_j = \left(\sum_{k} \lambda_{j,k}/L_x\right).$$

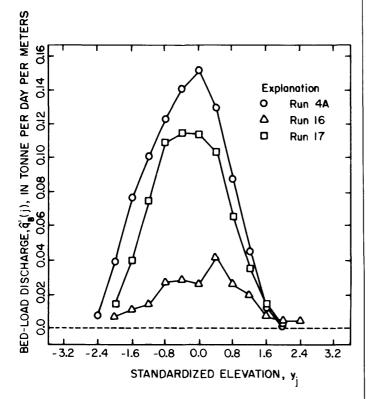


FIGURE 18. — Variation of bed-load discharge with bed elevation.

#### VARIATION OF VARIOUS STATISTICS WITH FLOW CONDITIONS AND A RELATION BETWEEN THE STEP LENGTH AND THE REST PERIODS

Although three flume runs are not sufficient to establish a reliable relation between the various statistics and flow conditions, some qualitative trends can be determined from table 73. The stream power (product of mean bed shear stress and mean flow velocity) was used as a measure of the flow conditions. From table 73, it is seen that:

- 1. The mean transport speed of a bed-material particle  $(\hat{V}_T, \hat{V}_T')$ , the average depth of the zone in which bed material movement occurs (h), and the standard deviation of the bed  $(s_y)$ , appear to increase with increasing stream power  $(\overline{T}_h \overline{U})$ .
- 2. The marginal mean of the step lengths (E[X]), the marginal variance of the step lengths (Var[X]), the marginal mean of the rest periods (E[T]), and the marginal variance of the rest periods (Var[T]), appear to decrease with increasing stream power within the range of stream power investigated here.

The variation of the ratios of the conditional mean step length to the conditional mean rest period

 $(\widehat{V}_B(y) = (\widehat{E}[X \setminus Y_D = y]) \ / \ (\widehat{E}[T \setminus Y_D = y]))$  and of the conditional variance of the step length to the conditional variance of the rest period

$$(\hat{Var} [X \setminus Y_D = y] / \hat{Var} [T \setminus Y_D = y])$$

with bed elevation, y, is shown in figures 19 and 20. From these figures it is seen that both ratios increase with increasing bed elevation.

#### TWO-DIMENSIONAL STOCHASTIC MODEL FOR DISPERSION OF BED-MATERIAL SEDIMENT PARTICLES

A two-dimensional stochastic model for dispersion of bed-material sediment particles was derived earlier and was given by equation 85,

$$f(x,y;t) = f_{Y_{D}}(y) \sum_{n=1}^{\infty} \left( \int_{0}^{x} f_{X}^{(n-1)} f_{X \setminus Y_{D}}(x - \zeta \setminus y) d\zeta \right) \cdot \int_{0}^{t} f_{T}^{(n)} dt' \int_{t-t'}^{\infty} f_{T \setminus Y_{D}}(\tau \setminus y) d\tau \right) . \tag{85}$$

The one-dimensional model as a marginal case of equation 85 was

$$f(x;t) = \sum_{n=1}^{\infty} f_X^{(n)} \int_0^t \left[ f_T^{(n)}(t') - f_T^{(t')} \right] dt' \quad . \tag{91}$$

Note that y is the standardized elevation. In order to apply equations 85 and 91, the probability density functions

from 
$$f_{Y_D}(y)$$
,  $f_{T\setminus Y_D}(t\setminus y)$ ,  $f_T(t)$ ,  $f_T(t)$ ,  $f_T(t)$ ,  $f_T(t)$ ,  $f_{X\setminus Y_D}(x\setminus y)$ ,  $f_X(x)$ ,  $f_X(x)$ , and  $f_X(x)$  must be specified.

Although probability density functions for all these distributions have not been determined in this report, the measured probability mass functions have been presented in tables 2, 3-5, 8-10, and 66-68, respec-

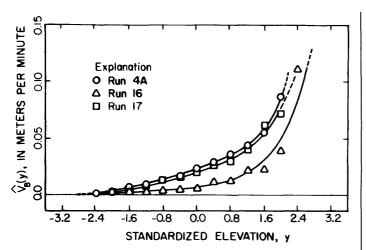


FIGURE 19. — Mean transport speed of a bed-load particle as a function of bed elevation, y.

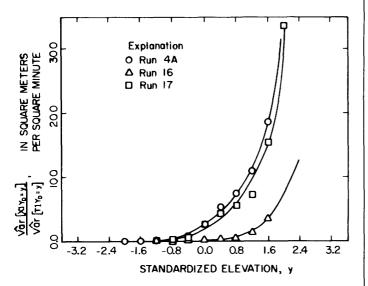


FIGURE 20. — Ratio of the conditional variance of step lengths to the conditional variance of rest periods as a function of bed elevation.

tively. Equations for determining the n-fold convolutions of  $p_T(t)$  and  $p_X(x)$  can be obtained from equations 82 and 79 with proper substitutions (Parzen, 1967). Further progress in the solution of either equation 85 or 91 could proceed along either of two lines. First, all probability density functions could be replaced with the corresponding sample probability mass functions, the integrals approximated by summations, and the solutions obtained numerically. Alternately, the mass functions could be fitted by density functions of some assumed form and an analytical solution attempted. Lee (1973) used various fitting procedures to obtain all the probability density functions required to solve equation 85, but the integration of the equation appears quite formidable.

#### **SUMMARY AND CONCLUSIONS**

Stochastic models were developed which can be used to predict the transport and dispersion of bed-material sediment particles in an alluvial channel. These models are based on the proposition that the movement of bed-material sediment particles consists of a series of steps separated by rest periods and, therefore, their application requires a knowledge of the probability distributions of the step lengths, the rest periods, and the elevation of particle deposition and erosion.

The probability distribution of the rest periods, conditioned on the elevation of particle deposition and the probability distributions of the elevation of particle erosion and deposition, were obtained from a record of the bed elevation at a fixed point as a continuous function of time  $[y_x(t)]$  record. The necessary assumptions were: (1) Equilibrium flow; (2) both erosion and deposition do not occur at the same point during the same time period; and (3) the number of particles per unit volume of the bed is constant.

The probability distribution of the step lengths, conditioned on the elevation of particle erosion and the elevation of particle deposition, was obtained from a series of instantaneous longitudinal bed profiles  $[y_t(x)]$  record. The required assumptions were: (1) All bed-material sediment particles which are eroded from the upstream face of a dune will be deposited on the downstream side of the same dune; and (2) no deposition occurs on the upstream sides of dunes, and no erosion occurs on the downstream faces of dunes. These assumptions appeared to be reasonable at least for a dune-covered bed composed of a coarse sand.

Introducing an additional assumption that the elevation of particle erosion and the elevation of particle deposition are mutually independent, various related probability distributions were obtained. These distributions included: (1) The marginal distributions of the rest periods and the step lengths; (2) the joint distribution of the rest periods and the elevation of particle deposition; and (3) the joint distribution of the step lengths and the elevation of particle deposition.

A two-dimensional stochastic model for dispersion of bed-sediment particles was then derived (eq. 85). In order to apply the model, the probability distributions of (1) the step lengths given the elevation of particle deposition; (2) the rest periods given the elevation of particle deposition; and (3) the elevation of particle deposition, must be known. The mass functions of these distributions were estimated; however, the integrations required by the model remained unsolved.

Applying the concept of continuity, three bedmaterial transport models were presented. Application of these models requires the estimation of: (1) The conditional means of the rest periods and the step lengths; (2) the probability distribution of the elevation of deposition; (3) the average depth of the zone of bed-material movement; and (4) the effective volume ratio. These were all obtained from the  $y_x(t)$  and  $y_t(x)$  records. In the derivation of the models, the bed load was defined as that part of bed material which is deposited on the downstream face of the dune from which it is eroded, and the suspended load was defined as that part of bed material which passes two or more dune crests before being deposited. These definitions are very precise compared to the definitions prepared by the Task Committee on Preparation of Sedimentation Manual (1962).

Based on flume experiments with a coarse sand, the following conclusions were drawn:

- 1. The elevation of particle erosion and the elevation of particle deposition can be considered to be identically distributed, and their distribution can be approximated by either a truncated Gaussian density function or a symmetric triangular density function. In general, the truncated Gaussian density provides slightly better results; although the triangular density is much easier to handle analytically.
- 2. The conditional probability distribution of the rest periods, given the elevation of deposition, can be well described by the two-parameter gamma density function. The shape of the conditional density approaches a J-shape and becomes more peaked as bed elevation decreases.
  - A. Both the conditional mean and variance of the rest periods increase with decreasing bed elevation. These relations can be expressed by exponential functions.
  - B. Both the scale and shape parameters for the conditional distribution of the rest periods increase with increasing bed elevation, and they can be described by exponential functions of bed elevation.
  - C. The correlation coefficient between the rest periods and the elevation of deposition indicated that the rest periods and the elevation of deposition are negatively correlated, but the degree of their linear association is not strong.
- 3. The conditional probability distribution of the step lengths, given the elevation of deposition and the elevation of erosion, can be approximated by the two-parameter gamma distribution. The shape of the conditional density is strongly dependent on the elevation of deposition and erosion.
  - A. For a fixed elevation of deposition, both the double conditional mean and variance of the step lengths increase with decreasing elevation of ero-

- sion. In other words, longer step lengths are associated with lower elevation at which a sediment particle is eroded or deposited and vice versa.
- B. The correlation coefficient between the step lengths and the elevation of deposition indicates that they are negatively correlated, but the degree of their linear association is not strong.
- 4. All three bed-material transport models are found to be quite satisfactory except for run 16.
  - A. The effective volume ratio can be obtained from either the  $y_t(x)$  record or the  $y_x(t)$  record, and it appears to be nearly independent of flow condition.
  - B. The maximum bed-load movement is associated with mean bed elevation, and little movement occurs for  $y \le -2.4$  and  $y \ge +2.4$ .
- 5. The mean transport speed of a bed-material particle, the average depth of the zone of bed material movement, and the standard deviation of bed elevation increased with increasing stream power, whereas the marginal means and variances of the rest periods and the step lengths decreased with increasing stream power.

Figures 10 and 15 suggest that the step lengths and the rest periods are positively correlated in an average sense, but the degree of linear association was not strong.

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# STOCHASTIC ANALYSIS OF PARTICLE MOVEMENT OVER A DUNE BED

Table 2. Sample probability mass functions of elevations of deposition and erosion

	Rur	4A	Rur	16	Rur	17		
Elevation y <sub>i</sub>	$p_{Y_D}(y_i)$	$p_{\underline{Y}_{\underline{E}}}(y_i)$	$p_{\underline{Y}_{\underline{D}}}(y_i)$	$p_{Y_{E}}(y_{i})$	$p_{\underline{Y}_{\overline{D}}}(y_i)$	$p_{Y_E}(y_i)$	Triangular Density	Truncated Gaussian
-3.6	0.000	0.000	0.000	0.000	0.000	0.000		
-3.2	.000	.000	.006	.006	.000	.000		
-2.8	.000	.000	.007	.007	.001	.002		
-2.4	.006	.006	.011	.013	.005	.006	0.004	0.006
-2.0	.025	.032	.019	.025	.017	.017	.028	.022
-1.6	.060	.060	.038	.044	.044	.047	. 0,56	.046
-1.2	.091	.088	.062	.065	.089	.089	.083	.079
-0.8	.125	.129	.104	.109	.129	.129	.111	.118
-0.4	.152	.140	.122	.116	.153	.148	.139	.148
0.0	.160	.166	.133	.134	.154	.157	.158	.162
0.4	.156	.154	.197	.189	.154	.153	.139	.148
0.8	.120	.116	.137	.130	.117	.118	.111	.118
1.2	.070	.068	.097	.097	.073	.072	.083	.079
1.6	.025	.032	.046	.044	.041	.041	.056	.046
2.0	.009	.009	.016	.016	.019	.017	.028	.022
2.4	.001	.001	.005	.005	.004	.004	.004	.006
2.8	.000	.000	.000	.000	.000	.000		
∑ <sub>i</sub> m <sub>i</sub> or ∑ <sub>i</sub> m' <sub>i</sub>	2,167	2,167	134	134	708	708		
$\hat{\mathbf{E}}[\mathbf{Y}_D]$ or $\hat{\mathbf{E}}[\mathbf{Y}_E]$	130	133	.055	.006	039	043		
$\widehat{\text{Var}}[Y_{\overline{D}}]$ or $\widehat{\text{Var}}[Y_{\overline{E}}]$	.813	. 850	.999	1.046	.863	. 870		

Note:  $\mathbf{m}_t$  is the total number of bed forms contained in the  $y_x(t)$  record and which also contain some deposition in the class interval associated with the elevation  $y_t$ .

 $m_{\cdot}^{l}$  is the total number of bed forms contained in the  $y_{x}(t)$  record and which also contain some erosion in the class interval associated with the elevation  $y_{i}$ .

## SUPPLEMENTAL DATA TABLES

Table 3. Sample conditional probability mass function of rest periods,  $p_{T|Y_{D}}(t_{\alpha}|y_{i})$  (Run 4A)

												$T I_{D}$				
1	Γα	0	10	20	30	40	50	60	70	80	90	100	110	120	130	140
	 Γα+1	10	20	30	40	50	60	70	80	90	100	110	120	130	140	150
	ta · I	5	15	25	35	45	55	65	75	85	95	105	115	125	135	145
	-2.8	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
	-2.4	0	0	0	0	0	0	0	0	0	0	0	0	0	.0909	0
	-2.0	.0448	.0746	.1045	.0896	.0448	.0298	.0149	.0298	.0298	.0149	0	.0298	.0149	.0149	.0149
<i>بع</i> 2.	-1.6	.0411	.1241	.1022	.0657	.0657	.1095	.0438	.0146	.0365	.0219	.0146	.0146	.0073	.0219	.0219
	-1.2	.1031	.1443	.1186	.0723	.0979	.0670	.0412	.0258	.0361	.0103	.0154	.0052	.0309	.0206	.0361
ion	-0.8	.0909	.2016	.1265	.0988	.0830	.0988	.0237	.0316	.0316	.0237	.0277	.0237	.0198	.0158	.0158
vat	-0.4	.1019	. 2229	.1337	.1752	.1178	.0637	.0478	.0350	.0223	.0159	.0127	.0096	.0159	.0064	.0032
elevation,	0.0	.1579	.2742	.2271	.1385	.0803	.0499	.0360	.0111	.0083	.0054	0	.0028	.0083	0	0
	0.4	.2189	. 36 39	.2041	.1036	.0710	.0148	.0177	.0059	0	0	0	0	0	0	0
Standardized	0.8	.3145	. 3952	.1694	.0766	.0202	.0242	0	0	0	0	0	0	0	0	0
dar	1.2	.4067	.3533	.1867	.0467	.0067	0	0	0	0	0	0	0	0	0	0
tan	1.6	.5139	. 3750	.0833	.0278	0	0	0	0	0	0	0	0	0	0	0
S	2.0	.7368	.2105	.0526	0	0	0	0	0	0	0	0	0	0	0	0
	2.4	.6667	. 3333	0	0	0	0	0	0	0	0	0	0	0	0	0
	2.8	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
-							·				·					
																i
•	ťα	150	160	170	180	190	200	250	300	350	400	500	600	1,000	2,000	
	τ <sub>α</sub> τα+1	160	160 170	170 180	190	200	250	250 300	300 350	350 400	500	600	1,000	2,000	8,000	$m_{i,i}$
•		1												-		<sup>m</sup> i,i
•	τα+1 tα	160 155	170 165	180 175	190 185	200 195	250 225	300 275	350 325	400 375	500 450	600 550	1,000	2,000	8,000 5,000	ļ
•	τα+1 tα	160 155	170 165	180 175	190 185	200 195 0	250 225 0	300 275 0	350 325 0	400 375	500 450 0	600 550 0	1,000	2,000 1,500	8,000 5,000	0
•	-2.8 -2.4	160 155 0 0	170 165 0 0	180 175 0 0	190 185 0 0	200 195 0 0	250 225 0 .2727	300 275 0 .0909	350 325 0 0	400 375 0 0	500 450 0 .0909	600 550 0	1,000 800 0 .0909	2,000 1,500 0 .0909	8,000 5,000 0 .2727	0 11
	-2.8 -2.4 -2.0	160 155 0 0	170 165 0 0	180 175 0 0 .0149	190 185 0 0 .0448	200 195 0 0	250 225 0 .2727 .0746	300 275 0 .0909 .0597	350 325 0 0 .0149	400 375 0 0 .0298	500 450 0 .0909 .0298	600 550 0 0 .0448	0 .0909 .0896	2,000 1,500 0 .0909 .0448	8,000 5,000 0 .2727 0	0 11 67
4,	-2.8 -2.4 -2.0 -1.6	160 155 0 0 0 .0438	170 165 0 0 0 0	180 175 0 0 .0149 .0146	190 185 0 0 .0448 .0073	200 195 0 0 0 0	250 225 0 .2727 .0746 .0219	300 275 0 .0909 .0597 .0219	350 325 0 0 0 .0149 .0438	400 375 0 0 0 .0298 .0219	500 450 0 .0909 .0298	0 0 0 .0448 .0 19	0 .0909 .0896	2,000 1,500 0 .0909 .0448	8,000 5,000 0 .2727 0	0 11 67 137
4,	-2.8 -2.4 -2.0 -1.6	160 155 0 0 0 .0438 .0258	170 165 0 0 0 .0219 .0052	180 175 0 0 .0149 .0146	190 185 0 0 .0448 .0073 .0103	200 195 0 0 0 0 .0219 .0155	250 225 0 .2727 .0746 .0219 .0309	300 275 0 .0909 .0597 .0219	350 325 0 0 .0149 .0438	400 375 0 0 .0298 .0219 .0052	500 450 0 .0909 .0298 .0292 .0154	0 0 0 .0448 .0 19	0 .0909 .0896 .0146	2,000 1,500 0 .0909 .0448 0	8,000 5,000 0 .2727 0 0	0 11 67 137 194
4,	-2.8 -2.4 -2.0 -1.6 -1.2	160 155 0 0 0 .0438 .0258	170 165 0 0 0 .0219 .0052 .0198	180 175 0 0 .0149 .0146 .0052	190 185 0 0 .0448 .0073 .0103	200 195 0 0 0 .0219 .0155 .0039	250 225 0 .2727 .0746 .0219 .0309	300 275 0 .0909 .0597 .0219 .0464	350 325 0 0 .0149 .0438 .0155	400 375 0 0 .0298 .0219 .0052	0 .0909 .0298 .0292 .0154	0 0 0 .0448 .0 19 0	0 .0909 .0896 .0146 0	2,000 1,500 0 .0909 .0448 0 0	8,000 5,000 0 .2727 0 0 0	0 11 67 137 194 253
	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4	160 155 0 0 0 .0438 .0258 .0158	170 165 0 0 0 .0219 .0052 .0198 .0032	180 175 0 0 .0149 .0146 .0052 .0079	0 0 0.0448 .0073 .0103	200 195 0 0 0 .0219 .0155 .0039	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032	300 275 0 .0909 .0597 .0219 .0464 .0079	350 325 0 0 .0149 .0438 .0155 0	400 375 0 0 .0298 .0219 .0052 0	0 .0909 .0298 .0292 .0154 .0039	0 0 0 .0448 .0 19 0 0	0 .0909 .0896 .0146 0	2,000 1,500 0 .0909 .0448 0 0	8,000 5,000 0 .2727 0 0 0	0 11 67 137 194 253 314
elevation, $y_i$	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4	0 0 0 .0438 .0258 .0158	170 165 0 0 0 .0219 .0052 .0198 .0032	180 175 0 0 .0149 .0146 .0052 .0079 .0032	190 185 0 0 .0448 .0073 .0103 .0079 0	200 195 0 0 0 .0219 .0155 .0039 0	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032	0 .0909 .0597 .0219 .0464 .0079 .0032	350 325 0 0 .0149 .0438 .0155 0	400 375 0 0 .0298 .0219 .0052 0 0	0 .0909 .0298 .0292 .0154 .0039 0	0 0 0 .0448 .0 19 0 0	0 .0909 .0896 .0146 0	2,000 1,500 0 .0909 .0448 0 0	8,000 5,000 0 .2727 0 0 0 0	0 11 67 137 194 253 314 361
elevation, $y_i$	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0	0 0 0 .0438 .0258 .0158 .0032	0 0 0 .0219 .0052 .0198 .0032 0	180 175 0 0 .0149 .0146 .0052 .0079 .0032 0	190 185 0 0 .0448 .0073 .0103 .0079 0 0	200 195 0 0 0 .0219 .0155 .0039 0 0	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032 0	0 .0909 .0597 .0219 .0464 .0079 .0032 0	350 325 0 0 .0149 .0438 .0155 0 0	0 0 0 .0298 .0219 .0052 0 0	0 .0909 .0298 .0292 .0154 .0039 0	0 0 .0448 .0 19 0 0	0 .0909 .0896 .0146 0 0	2,000 1,500 0 .0909 .0448 0 0 0	8,000 5,000 0 .2727 0 0 0 0	0 11 67 137 194 253 314 361 338
elevation, $y_i$	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4	0 0 0 .0438 .0258 .0158 .0032 0	170 165 0 0 0 .0219 .0052 .0198 .0032 0	180 175 0 0 .0149 .0146 .0052 .0079 .0032 0	190 185 0 0 .0448 .0073 .0103 .0079 0 0	200 195 0 0 0 .0219 .0155 .0039 0 0	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032 0	300 275 0 .0909 .0597 .0219 .0464 .0079 .0032 0	350 325 0 0 .0149 .0438 .0155 0 0	0 0 0.0298 .0219 .0052 0 0	0 .0909 .0298 .0292 .0154 .0039 0 0	0 0 0 .0448 .0 19 0 0 0	0 .0909 .0896 .0146 0 0	2,000 1,500 0 .0909 .0448 0 0 0	8,000 5,000 0 .2727 0 0 0 0	0 11 67 137 194 253 314 361 338 248
elevation, $y_i$	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2	0 0 0 .0438 .0258 .0158 .0032 0 0	170 165 0 0 0 .0219 .0052 .0198 .0032 0 0	180 175 0 0 .0149 .0146 .0052 .0079 .0032 0 0	190 185 0 0 .0448 .0073 .0103 .0079 0 0 0	200 195 0 0 0 .0219 .0155 .0039 0 0	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032 0 0	300 275 0 .0909 .0597 .0219 .0464 .0079 .0032 0 0	350 325 0 0 .0149 .0438 .0155 0 0 0	0 0 0.0298 .0219 .0052 0 0 0	0 .0909 .0298 .0292 .0154 .0039 0 0	0 0 0 .0448 .0 19 0 0 0	0 .0909 .0896 .0146 0 0 0 0 0 0 0	2,000 1,500 0 .0909 .0448 0 0 0 0	8,000 5,000 0 .2727 0 0 0 0 0	0 11 67 137 194 253 314 361 338 248 150
4,	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2 1.6	0 0 0 .0438 .0258 .0158 .0032 0 0 0	170 165 0 0 0.0219 .0052 .0198 .0032 0 0	180 175 0 0 .0149 .0146 .0052 .0079 .0032 0 0 0	190 185 0 0 .0448 .0073 .0103 .0079 0 0 0	200 195 0 0 0.0219 .0155 .0039 0 0 0	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032 0 0	300 275 0 .0909 .0597 .0219 .0464 .0079 .0032 0 0	350 325 0 0 .0149 .0438 .0155 0 0 0 0	0 0 0.0298 .0219 .0052 0 0 0 0	0 .0909 .0298 .0292 .0154 .0039 0 0 0 0 0 0	0 0 0 .0448 .0 19 0 0 0 0	0 .0909 .0896 .0146 0 0 0 0 0 0 0 0 0	2,000 1,500 0 .0909 .0448 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	8,000 5,000 0 .2727 0 0 0 0 0 0	0 11 67 137 194 253 314 361 338 248 150 72
elevation, $y_i$	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2 1.6 2.0	0 0 0 .0458 .0258 .0158 .0032 0 0 0	0 0 0 .0219 .0052 .0198 .0032 0 0 0	180 175 0 0 .0149 .0146 .0052 .0079 .0032 0 0 0	190 185 0 0 .0448 .0073 .0103 .0079 0 0 0 0	200 195 0 0 0 .0219 .0155 .0039 0 0 0	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032 0 0 0 0 0 0 0	300 275 0 .0909 .0597 .0219 .0464 .0079 .0032 0 0	350 325 0 0 .0149 .0438 .0155 0 0 0 0	400 375 0 0.0298 .0219 .0052 0 0 0 0	0 .0909 .0298 .0292 .0154 .0039 0 0 0 0 0 0 0 0 0	0 0 0 .0448 .0 19 0 0 0 0	1,000 800 0.0909 .0896 .0146 0 0 0 0	2,000 1,500 0.0909 .0448 0 0 0 0 0	8,000 5,000 0 .2727 0 0 0 0 0 0 0	0 11 67 137 194 253 314 361 338 248 150 72 19
elevation, $y_i$	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2 1.6	0 0 0 .0438 .0258 .0158 .0032 0 0 0	170 165 0 0 0.0219 .0052 .0198 .0032 0 0	180 175 0 0 .0149 .0146 .0052 .0079 .0032 0 0 0	190 185 0 0 .0448 .0073 .0103 .0079 0 0 0	200 195 0 0 0.0219 .0155 .0039 0 0 0	250 225 0 .2727 .0746 .0219 .0309 .0198 .0032 0 0	300 275 0 .0909 .0597 .0219 .0464 .0079 .0032 0 0	350 325 0 0 .0149 .0438 .0155 0 0 0	0 0 0.0298 .0219 .0052 0 0 0 0	0 .0909 .0298 .0292 .0154 .0039 0 0 0 0 0 0	0 0 0 .0448 .0 19 0 0 0 0	0 .0909 .0896 .0146 0 0 0 0 0 0 0 0 0	2,000 1,500 0 .0909 .0448 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	8,000 5,000 0 .2727 0 0 0 0 0 0	0 11 67 137 194 253 314 361 338 248 150 72

Note:  $\tau_{\alpha}$ ,  $\tau_{\alpha+1}$ , and  $t_{\alpha}$  are in minutes.

 $m_{i,i}$  is the total number of bed forms contained in the  $y_x(t)$  record and which also contain both an up-crossing and a down-crossing at the elevation  $y_i$ .

# STOCHASTIC ANALYSIS OF PARTICLE MOVEMENT OVER A DUNE BED

Table 4. Sample conditional probability mass function of rest periods,  $p_{T\mid Y_{D}}(t_{\alpha}|y_{i})$  (Run 16)

										$^{1}$ $^{1}$ $^{2}$ $^{D}$		
1	a	0	20	40	60	80	100	120	140	160	180	200
7	~+1	20	40	60	80	100	120	140	160	180	200	220
t	α	10	30	50	70	90	110	130	150	170	190	210
	-2.8	0	0	0	0	0	0	0	0	0	0	0
	-2.4	0	0	0	0	0	0	0	0	0	0	0
	-2.0	0	0	0	0	0	0	0	0	0	0	0
<i>₩</i>	-1.6	0	0	0	0	0	0	0	.1667	0	0	0
Ę	-1.2	.2500	0	0	0	0	0	0	.1250	0	0	0
tic	-0.8	.1333	.0667	.0667	0	0	.0667	.0667	.0667	0	0	.0666
eva	-0.4	.1765	.0588	0	.0588	.0588	.0588	.0588	.0588	.0588	0	.0588
<u></u>	0.0	.1111	.0556	.1111	.0556	.0556	.0556	.1111	.1111	.0555	0	0
Standardized elevation, $ extit{y}_{ec{t}}$	0.4	.4138	.1034	.0690	.0345	.0345	.2069	.0345	0	.0345	.0345	0
rdi	0.8	.2778	.1667	.1111	.1111	.0556	.0556	.1666	0	.0556	0	0
nda	1.2	.5385	0	.2308	0	.2308	0	0	0	0	0	0
Sta	1.6	.2000	.4000	.4000	0	0	0	0	0	0	0	0
	2.0	.5000	.5000	0	0	0	0	0	0	0	0	0
	2.4	1.0000	0	0	0	0	0	0	0	0	0	0
	2.8	0	0	0	0	0	0	0	0	0	0	0
	***				<del></del>							
			240				FAA	600		900	1 700	
	α	220	240	260	300	400	500	600	700	800	1,300	
т	α+1	240	260	300	400	500	600	700	800	1,300	1,800	m;,i
т		1										<sup>m</sup> i,i
т	α+1 α	240 230 0	260 250 0	300 280 0	400 350 0	500 450 0	600	700 650 0	800 750 <b>0</b>	1,300 1,050	1,800 1,550	0
т	α+1 α	240 230	260 250	300 280	400 350	500 450	600 550	700 650	800 750	1,300 1,050	1,800 1,550	ļ
т	α+1 α	240 230 0	260 250 0	300 280 0	400 350 0	500 450 0	600 550 0	700 650 0	800 750 <b>0</b>	1,300 1,050	1,800 1,550	0
t	α+1 α -2.8 -2.4	240 230 0 0	260 250 0 0	300 280 0 0	400 350 0 0	500 450 0 0	600 550 0	700 650 0	800 750 <b>0</b> 0	1,300 1,050 0	1,800 1,550 0 0	0 0
t	α+1 α -2.8 -2.4 -2.0	240 230 0 0	260 250 0 0 0 0	300 280 0 0	400 350 0 0	500 450 0 0	600 550 0 0	700 650 0 0	800 750 0 0	1,300 1,050 0 0	1,800 1,550 0 0 1.0000	0 0 2
t	α+1 α -2.8 -2.4 -2.0 -1.6	240 230 0 0 0	260 250 0 0 0	300 280 0 0 0	400 350 0 0 0 .1667	500 450 0 0 0	600 550 0 0 0	700 650 0 0 0 0	800 750 0 0 0 0	1,300 1,050 0 0 0 .1666	1,800 1,550 0 0 1.0000 .1666	0 0 2 6
t	α+1 α -2.8 -2.4 -2.0 -1.6 -1.2	240 230 0 0 0 0	260 250 0 0 0 0	300 280 0 0 0 0 0	400 350 0 0 0 .1667	500 450 0 0 0 0 0	0 0 0 0 0	700 650 0 0 0 .1667	800 750 0 0 0 .1667 .1250	1,300 1,050 0 0 0 .1666 .1250	1,800 1,550 0 0 1.0000 .1666	0 0 2 6 8
t	α+1 α -2.8 -2.4 -2.0 -1.6 -1.2	240 230 0 0 0 0 0 0	260 250 0 0 0 0 0 0	300 280 0 0 0 0 0 .1250	400 350 0 0 0 .1667 0 .1333	500 450 0 0 0 0 .1250 .0667	0 0 0 0 0 0 0	700 650 0 0 0 .1667 0	800 750 0 0 0 .1667 .1250	1,300 1,050 0 0 0 .1666 .1250	1,800 1,550 0 0 1.0000 .1666 .1250	0 0 2 6 8 15
t	α+1 α -2.8 -2.4 -2.0 -1.6 -1.2 -0.8	240 230 0 0 0 0 0 0 0	260 250 0 0 0 0 0 0 0 0	0 0 0 0 0 .1250 .0666	400 350 0 0 0 .1667 0 .1333	0 0 0 0 0 .1250 .0667	0 0 0 0 0 0 0 0	700 650 0 0 0 .1667 0 .0667	800 750 0 0 0 .1667 .1250 0	0 0 0 0 .1666 .1250	1,800 1,550 0 0 1.0000 .1666 .1250 0	0 0 2 6 8 15
t	α+1 α  -2.8  -2.4  -2.0  -1.6  -1.2  -0.8  -0.4  0.0	240 230 0 0 0 0 0 0 .0666	260 250 0 0 0 0 0 0 0 0 0	300 280 0 0 0 0 .1250 .0666 .1176	400 350 0 0 0 .1667 0 .1333 .0588	500 450 0 0 0 0 .1250 .0667 .0588	0 0 0 0 0 0 0 0 .0667 0	700 650 0 0 0 .1667 0 .0667 .0588	800 750 0 0 0 .1667 .1250 0	0 0 0 .1666 .1250 0	1,800 1,550 0 0 1.0000 .1666 .1250 0 0	0 0 2 6 8 15 17 18
t	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0	240 230 0 0 0 0 0 0 .0666 0 0	260 250 0 0 0 0 0 0 0 0 0	0 0 0 0 0 .1250 .0666 .1176	400 350 0 0 0 .1667 0 .1333 .0588 0	0 0 0 0 0 .1250 .0667 .0588 0	0 0 0 0 0 0 0 .0667 0 .0555	700 650 0 0 0 .1667 0 .0667 .0588	800 750 0 0 0 .1667 .1250 0 0	0 0 0 .1666 .1250 0	1,800 1,550 0 0 1.0000 .1666 .1250 0 0	0 0 2 6 8 15 17 18 29
t	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0	240 230 0 0 0 0 0 0 .0666 0 0 .0345	260 250 0 0 0 0 0 0 0 0 0 .0588 .1111	0 0 0 0 0 .1250 .0666 .1176 .1111	0 0 0 0 .1667 0 .1333 .0588 0	0 0 0 0 0 .1250 .0667 .0588 0 0	0 0 0 0 0 0 0 .0667 0 .0555	700 650 0 0 0 .1667 0 .0667 .0588 0	800 750 0 0 0 .1667 .1250 0 0	1,300 1,050 0 0 0.1666 .1250 0 0	1,800 1,550 0 0 1.0000 .1666 .1250 0 0	0 0 2 6 8 15 17 18 29
т	α+1 α  -2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2	240 230 0 0 0 0 0 0 .0666 0 0 .0345	260 250 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	300 280 0 0 0 0 .1250 .0666 .1176 .1111 0	0 0 0 0 .1667 0 .1333 .0588 0 0	0 0 0 0 0 .1250 .0667 .0588 0 0	0 0 0 0 0 0 0 .0667 0 .0555 0	700 650 0 0 0 .1667 0 .0667 .0588 0 0	800 750 0 0 0 .1667 .1250 0 0 0	1,300 1,050 0 0 0.1666 .1250 0 0	1,800 1,550 0 0 1.0000 .1666 .1250 0 0 0	0 0 2 6 8 15 17 18 29 18
t	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2 1.6	240 230 0 0 0 0 0 0 .0666 0 0 .0345 0	260 250 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	300 280 0 0 0 0 .1250 .0666 .1176 .1111 0 0	400 350 0 0 0 .1667 0 .1333 .0588 0 0 0	0 0 0 0 0 1250 .0667 .0588 0 0 0	0 0 0 0 0 0 0 .0667 0 .0555 0 0	700 650 0 0 0 .1667 0 .0667 .0588 0 0	800 750 0 0 0 .1667 .1250 0 0 0 0	1,300 1,050 0 0 0 .1666 .1250 0 0 0	1,800 1,550 0 0 1.0000 .1666 .1250 0 0 0	0 0 2 6 8 15 17 18 29 18 13 5
t	-2.8 -2.4 -2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2 1.6 2.0	240 230 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	260 250 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	300 280 0 0 0 0 .1250 .0666 .1176 .1111 0 0	400 350 0 0 0 .1667 0 .1333 .0588 0 0 0	500 450 0 0 0 0 0 .1250 .0667 .0588 0 0 0	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	700 650 0 0 0 .1667 0 .0667 .0588 0 0 0	800 750 0 0 0 .1667 .1250 0 0 0 0	1,300 1,050 0 0 0 .1666 .1250 0 0 0	1,800 1,550 0 0 1.0000 .1666 .1250 0 0 0 0	0 0 2 6 8 15 17 18 29 18 13 5

Note:  $\tau_{\alpha}$ ,  $\tau_{\alpha+1}$ , and  $t_{\alpha}$  are in minutes.

 $m_{i}$ , i is the total number of bed forms contained in the  $y_x(t)$  record and which also contain both an up-crossing and a down-crossing at the elevation  $y_i$ .

# SUPPLEMENTAL DATA TABLES

Table 5. Sample conditional probability mass function of rest periods,  $p_{T\mid Y_{n}}(t_{\alpha}|y_{i})$  (Run 17)

											$T_{1}^{1}D$	a ··· ı		
τ,	α	0	10	20	30	40	50	60	70	80	90	100	110	120
τ	α+1	10	20	30	40	50	60	70	80	90	100	110	120	130
$t_{0}$	x.	5	15	25	35	45	55	65	75	85	95	105	115	125
	-2.8	0	0	0	0	0	0	0	0	0	0	0	0	0
	-2.4	0	0	0	0	0	0	0	0	0	0	0	0	0
	-2.0	0	0	0	0	.1000	.1000	0	0	0	0	0	.1000	.1000
ά. .υ.	-1.6	.0322	.0322	.0322	.1290	.0968	.0322	0	.0322	0	.0323	.0323	.0968	0
ŕ	-1.2	.0328	.0164	.1311	.0984	.1639	.0492	.1639	.0656	.0328	.0328	.0328	.0164	.0164
tio	-0.8	.0652	.0652	.1630	.1522	.1739	.1304	.0978	.0326	.0217	0	.0109	.0217	0
eva	-0.4	.0385	.1442	.2404	.2115	.1250	.1058	.0288	.0385	.0096	.0192	.0096	0	0
e1	0.0	.0789	.2281	.2544	.2105	.1228	.0351	.0526	0	0	.0088	.0088	0	0
zeq	0.4	.1892	.3063	.2703	.1261	.0631	.0270	.0180	0	0	0	0	0	٥
Standardized elevation, $ extcolored{y}_t^{}$	0.8	. 2625	.3250	. 2500	.0875	.0500	.0250	0	0	0	0	0	0	0
nda	1.2	.4286	.2857	.1786	.0893	.0179	0	0	0	0	0	0	0	0
Sta	1.6	.6250	.2500	.1250	0	0	0	0	0	0	0	0	0	0
	2.0	.7273	.2727	0	0	0	0	0	0	0	0	0	0	0
	2.4	1.0000	0	0	0	0	0	0	0	0	0	0	0	0
	2.8	0	0	0	0	0	0	0	0	0	0	0	0	0
τ	·	130	140	150	200	250	300	400	500	1,000	2,000	3,000		
	α+1	140	150	200	250	300	400	500	1,000	2,000	3,000	4,000		$m_{i,i}$
t		135	145	175	225	275	350	450	750	1,500	2,500	3,500		
	-2.8	0	0	0	0	0	0	0	0	0	0	0		0
	-2.4	l .							U	v	U	U	- 1	
	ı	0	.3333	0	.3334	0	0	0	0	0	0	. 3333		3
	-2.0	.1000	.3333											3 10
$y_i$		i .		0	.3334	0	0	0	0	0	0	.3333		
n, y <sub>i</sub>	-2.0	.1000	0	0 0	.3334	0 .1000	0	0 .1000	0 .1000	0 .1000	0 .1000	.3333		10
tion, $y_i$	-2.0 -1.6	.1000	0.0322	0 0 .0322	.3334 0 .1613	0 .1000 .0322	0 0 .0322	0 .1000 .0645	0 .1000 .0645	0 .1000 0	0 .1000 0	.3333 0 0		10 31
evation, $\boldsymbol{y}_i$	-2.0 -1.6 -1.2	.1000 .0322 .0164	0 .0322 0	0 0 .0322 .0328	.3334 0 .1613 .0328	0 .1000 .0322 0	0 0 .0322 .0328	0 .1000 .0645 .0164	0 .1000 .0645 .0164	0 .1000 0 0	0 .1000 0 0	.3333 0 0 0		10 31 61
elevation, $y_i$	-2.0 -1.6 -1.2 -0.8	.1000 .0322 .0164 .0109	0 .0322 0 .0109	0 0 .0322 .0328 .0217	.3334 0 .1613 .0328	0 .1000 .0322 0 .0109	0 0 .0322 .0328 .0109	0 .1000 .0645 .0164	0 .1000 .0645 .0164 0	0 .1000 0 0	0 .1000 0 0	.3333 0 0 0 0		10 31 61 92
zed elevation, $y_i$	-2.0 -1.6 -1.2 -0.8 -0.4	.1000 .0322 .0164 .0109	0 .0322 0 .0109	0 0 .0322 .0328 .0217 0	.3334 0 .1613 .0328 0	0 .1000 .0322 0 .0109	0 0 .0322 .0328 .0109	0 .1000 .0645 .0164 0	0 .1000 .0645 .0164 0	0 .1000 0 0 0	0 .1000 0 0 0	.3333 0 0 0 0 0		10 31 61 92 104
rdızed elevation, $ extit{y}_i$	-2.0 -1.6 -1.2 -0.8 -0.4	.1000 .0322 .0164 .0109 0	0 .0322 0 .0109 0	0 0 .0322 .0328 .0217 0	.3334 0 .1613 .0328 0 0	0 .1000 .0322 0 .0109 .0096	0 0 .0322 .0328 .0109 0	0 .1000 .0645 .0164 0	0 .1000 .0645 .0164 0	0 .1000 0 0 0 0	0 .1000 0 0 0 0	.3333 0 0 0 0 0		10 31 61 92 104 114
ndardized elevation, $oldsymbol{y}_t$	-2.0 -1.6 -1.2 -0.8 -0.4 0.0	.1000 .0322 .0164 .0109 0	0 .0322 0 .0109 0 0	0 0 .0322 .0328 .0217 0 0	.3334 0 .1613 .0328 0 0	0 .1000 .0322 0 .0109 .0096 0	0 0 .0322 .0328 .0109 0 0	0 .1000 .0645 .0164 0 0	0 .1000 .0645 .0164 0 0	0 .1000 0 0 0 0 0	0 .1000 0 0 0 0	.3333 0 0 0 0 0 0 0		10 31 61 92 104 114
Standardized elevation, $oldsymbol{y}_{oldsymbol{t}}$	-2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4	.1000 .0322 .0164 .0109 0 0	0 .0322 0 .0109 0 0	0 0 .0322 .0328 .0217 0 0	.3334 0 .1613 .0328 0 0 0	0 .1000 .0322 0 .0109 .0096 0	0 0 .0322 .0328 .0109 0 0	0 .1000 .0645 .0164 0 0 0	0 .1000 .0645 .0164 0 0 0	0 .1000 0 0 0 0 0	0 .1000 0 0 0 0 0	.3333 0 0 0 0 0 0 0		10 31 61 92 104 114 111 80
Standardızed elevation, $ extit{ heta}_{t}^{i}$	-2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2	.1000 .0322 .0164 .0109 0 0	0 .0322 0 .0109 0 0 0	0 0 .0322 .0328 .0217 0 0 0	.3334 0 .1613 .0328 0 0 0	0 .1000 .0322 0 .0109 .0096 0 0	0 0 .0322 .0328 .0109 0 0 0	0 .1000 .0645 .0164 0 0 0	0 .1000 .0645 .0164 0 0 0	0 .1000 0 0 0 0 0 0	0 .1000 0 0 0 0 0 0	.3333 0 0 0 0 0 0 0 0		10 31 61 92 104 114 111 80 56
Standardızed elevation, $y_t^{}$	-2.0 -1.6 -1.2 -0.8 -0.4 0.0 0.4 0.8 1.2	.1000 .0322 .0164 .0109 0 0 0	0 .0322 0 .0109 0 0 0	0 0 .0322 .0328 .0217 0 0 0	.3334 0 .1613 .0328 0 0 0 0	0 .1000 .0322 0 .0109 .0096 0 0	0 0 .0322 .0328 .0109 0 0 0	0 .1000 .0645 .0164 0 0 0 0	0 .1000 .0645 .0164 0 0 0 0	0 .1000 0 0 0 0 0	0 .1000 0 0 0 0 0 0	.3333 0 0 0 0 0 0 0 0 0		10 31 61 92 104 114 111 80 56 32

Note:  $\tau_{\alpha}$ ,  $\tau_{\alpha+1}$ , and  $t_{\alpha}$  are in minutes.

 $m_{i,i}$  is the total number of bed forms contained in the  $y_x(t)$  record and which also contain both an up-crossing and a down-crossing at the elevation  $y_i$ .

Table	6.	Variation of conditional mean and variance of	f
		rest periods with elevation of deposition;	
		$\hat{E}[T Y_{D}=y]$ and $\widehat{Var}[T Y_{D}=y]$	

Standardized Elevation	Ê[I	$[Y_{D}=y_{i}].$	min	Var[2	$[Y_{D}=y_{i}],$	min <sup>2</sup>
y <sub>i</sub>	Run 4A	Run 16	Run 17	Run 4A	Run 16	Run 17
-2.8						
-2.4	1,400.8		1,443.7	4,179,378		4,671,616
-2.0	252.2	1,550.3	610.6	109,148	84,679	762,079
-1.6	120.5	714.8	198.0	19,552	204,840	74,793
-1.2	83.4	505.6	92.5	8,696	246,658	14,193
-0.8	58.2	230.9	53.5	3,529	39,416	2,548
-0.4	40.2	185.8	40.4	1,288	28,686	1,129
0.0	27.4	151.7	29.8	409	16,715	316
0.4	21.0	64.5	22.3	193	4,030	188
0.8	16.7	61.5	19.0	127	2,691	145
1.2	13.9	36.0	14.0	82	1,039	107
1.6	11.4	35.6	9.4	49	351	50
2.0	8.5	19.8	7.5	30	2	18
2.4	7.1	6.9	2.6	29		0.
2.8						
Ê[T]	53.8	180.5	60.1			
Var[T]				42,110	93,077	60,469
$\sum_{i}^{m}$	2,167	134	708	2,167	134	708

Table 7. Estimates of parameters and the results of goodness of fit test for the conditional rest periods (two-parameter gamma)

st indar bred		Run	4A			Ru	n 16			Run	17	
Elevation ${}^{y}i$	k. 1/	r <sub>2,y</sub> 2/	$^{m}i$ , $i$	Goodness of Fit Test <sup>3</sup> /	k <sub>2,y</sub> min-1	r <sub>2,y</sub>	m <sub>i,i</sub>	Goodness of Fit Test	k <sub>2,y</sub>	r <sub>2,y</sub>	$^m$ i,i	Goodness of Fit Test
-2.4	0.0003	0.470	11				0				3	
-2.0	.0023	.583	67				2		0.0008	0.489	10	
-1.6	.006	.742	137	$x^2 > x_c^2$	0.003	2.494	6		.003	.524	31	$x^2 < x_c^2$
-1.2	.010	.800	194	$x^2 > x_c^2$	. 002	1.036	8		.007	.604	61	$x^2 > x_c^2$
-0.8	.016	.959	253	$x^2 < x_c^2$	. 006	1.353	15		.021	1.125	92	$x^{2} > x_{c}^{2}$
-0.4	.031	1.255	314	$x^2 > x_c^2$	.006	1.204	17	$x^2 < x_c^2$	. 036	1.444	104	$x^2 > x_c^2$
0.0	.067	1.835	361	$x^2 < x_c^2$	.009	1.376	18	$x^2 < x_c^2$	.094	2.811	114	$x^2 < x_c^2$
0.4	.109	2.276	338	$x^2 < x_c^2$	.016	1.031	29	$x^2 < x_c^2$	.118	2.638	111	$x^2 < x_c^2$
0.8	. 131	2.193	248	$x^2 < x_a^2$	.023	1.407	18	$x^2 < x_c^2$	. 131	2.488	80	$x^2 < x_c^2$
1.2	.157	2.361	150	$x^2 < x_c^2$	.035	1.247	13		.131	1.827	56	$x^2 < x_c^2$
1.6	.235	2.680	72	$x^2 < x_c^2$	.101	3.611	5		.186	1.747	32	$x^2 < x_{\mathcal{O}}^2$
2.0	.283	2.408	19				2		.417	3.125	11	
2.4			3				1				3	

 $<sup>\</sup>frac{1}{k_{2,y}} = \frac{\hat{\mathbf{E}}[T|Y_D = y]}{\widehat{\mathbf{Var}}[T|Y_D = y]}$ 

 $r_{2,y} = \frac{\left(\hat{\mathbf{E}}[T|Y_D=y]\right)^2}{\sqrt{\hat{\mathbf{Ar}}[T|Y_D=y]}}$ 

 $x_c^2$  = critical chi-square value at a significant level of 0.05

#### SUPPLEMENTAL DATA TABLES

Table 8. Sample joint probability mass function of rest periods and elevation of deposition,  $p_{T,Y_n}(t_\alpha,y_t)$  (Run 4A) 140 0 30 40 70 80 90 100 110 120 130 10 20 50 60  $\tau_{\alpha}$ 10 20 30 40 50 60 70 80 90 100 110 120 130 140 150  $^{\tau}_{\alpha+1}$ 5 35 75 85 105 125 135 145 15 65  $t_{\alpha}$ -2.8 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 .0005 -2.4 0 0 0 0 0 0 0 0 0 0 0 0 .0004 -2.0 .0011 .0019 .0027 .0023 .0011 .0008 .0004 .0008 .0008 .0004 G .0008 .0004 .0004 -1.6 .0025 .0075 .0062 .0040 .0040 .0066 .0009 .0022 .0009 .0009 .0004 .0013 .0026 -1.2 .0093 .0065 .0089 .0014 .0019 .0033 .0131 .0107 .0061 .0037 .0023 .0033 .0009 .0005 .0028 elevation, -0.8 .0114 .0253 .0158 .0124 .0104 .0124 .0030 .0040 .0040 .0030 .0035 .0030 .0025 .0020 .0020 -0.4 .0155 .0338 0203 .0266 .0179 .0072 .0053 ,0034 .0019 .0014 .0024 .0010 .0005 .0097 .0024 0.0 .0253 .0440 .0365 .0222 .0129 .0080 .0058 .0018 .0013 .0009 0 .0004 .0013 0 0 Standardized .0341 .0567 .0318 .0161 .0111 .0023 .0028 .0009 0.8 .0377 0 .0473 .0203 .0092 .0024 .0029 0 0 0 0 0 0 0 0 1.2 .0286 .0248 .0131 .0033 .0005 0 0 0 0 0 0 0 0 0 0 1.6 .0127 .0093 .0021 .0007 0 0 0 0 0 0 0 2.0 .0064 .0018 .0004 0 0 0 0 0 0 0 0 Ω 0 0 ß 2.4 .0005 .0003 0 0 0 0 0 0 0 0 2.8 0 0 0 0 0 0 0 ٥ 0 0 0 0 0 0 0  $p_T(t_\alpha)$ .1857 . 2658 .0077 .0070 .0070 .0074 .1600 .1033 .0691 .0487 .0255 .0160 .0150 .0089  $\tau_{\alpha}$ 150 160 170 180 190 200 250 350 400 500 600 1,000 2,000 160 170 180 190 200 400 600 1,000 2,000 8,000 250 300 350 500 <sup>τ</sup>α+1  $t_{\alpha}$ 155 165 175 185 195 225 275 325 375 450 550 800 1,500 5,000 -2.8 0 0 0 0 0 0 0 0 0 0 0 0 0 0 -2.4 0 0 0 0 0 .0015 .0005 0 0 .0005 0 .0005 .0005 .0015 -2.0 0 0 .0011 0 .0019 .0008 .0011 0 .0004 .0015 .0004 .0008 .0011 .0023 -1.6 .0026 .0013 .0009 . 0004 .0013 .0013 .0013 .0026 .0013 .0018 .0013 .0009 0 0 -1.2 .0023 .0005 .0005 .0009 .0014 .0028 .0014 .0005 .0043 .0014 0 Standardized elevation, -0.8 .0020 .0025 .0010 .0010 .0005 .0025 .0010 0 0 0 0 0 0 O -0.4 .0005 .0005 .0005 0 .0005 .0005 0 0 .0005 0 0 0 0 0.0 0 0 0 0 0 0 0 0 0 0 0 0 0.4 0 0 0 0 0 0 0 0 0 0 0.8 0 0 0 0 0 0 0 0 0 1.2 0 0 0 0 0 0 0 0 0 0 0 0 0 0 1.6 0 0 0 0 0 2.0 0 0 0 0 0 0 0 0 0 0 0 0 0 2.4 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0  $p_T(t_{\alpha})$ .0074 .0048 .0032 .0035 .0032 .0105 .0090 .0044 .0026 .0049 .0024 .0037 .0017 .0015

Note:  $\tau_{\alpha}$ ,  $\tau_{\alpha+1}$ , and  $t_{\alpha}$  are in minutes.

# STOCHASTIC ANALYSIS OF PARTICLE MOVEMENT OVER A DUNE BED

Table 9. Sample joint probability mass function of rest periods and elevation of deposition,  $P_{T,Y_{-}}(t_{x},y_{x})$  (Run 16)

					$p_{T_{i}}$	$Y_D^{(t_{\alpha}, y_i)}$	) (Run 1	6)				
τ	α	0	20	40	60	80	100	120	140	160	180	200
	α+1	20	40	60	80	100	120	140	160	180	200	220
	α	110	30	50	70	90	110	130	150	170	190	210
	-2.8	0	0	0	0	0	0	0	0	0	0	0
	-2.4	0	0	0	0	0	0	0	0	0	0	0
	-2.0	0	0	0	0	0	0	0	0	0	0	0
·	-1.6	0	0	0	0	0	0	0	.0063	0	0	0
Standardized elevation, $y_i$	-1.2	.0151	0	0	0	0	0	0	.0075	0	0	0
ıtic	-0.8	.0138	.0069	.0069	0	0	.0069	.0069	.0069	0	0	.0069
eva	-0.4	.0216	.0072	0	,0072	.0072	.0072	.0072	.0072	.0072	0	.0072
[e]	0.0	.0148	.0074	.0148	.0074	.0074	.0074	.0148	.0148	.0074	0	0
zec	0.4	.0817	.0204	.0136	.0068	.0068	.0409	.0068	0	.0068	.0068	0
ırdi	0.8	.0381	.0229	.0153	.0153	.0076	.0076	.0229	0	.0076	0	0
ında	1.2	.0524	0	.0225	0	.0225	0	0	0	0	0	0
Sta	1.6	.0093	.0185	.0185	0	0	0	0	0	0	0	0
	2.0	.0082	.0082	0	0	0	0	0	0	0	0	0
	2.4	.0048	0	0	0	0	0	0	0	0	0	0
	2.8	0	0	0	0	0	0	0	0	0	0	0
p	$T^{(t_{\alpha})}$	. 2598	.0915	.0916	.0367	.0515	.0700	.0586	.0427	.0290	.0068	.0141
τ	α	220	240	260	300	400	500	600	700	800	1,300	
	α+1	240	260	300	400	500	600	700	800	1,300	1,800	
t	α	230	250	280	350	450	550	650	750	1,050	1,550	
	-2.8	0	0	0	0	0	0	0	0	0	0	
	-2.4	0	0	0	0	0	0	0	0	0	0	
	-2.0	o	.0	0	0	0	0	0	0	0	.0191	
.67	-1.6	0	0	0	.0063	0	0	.0063	.0063	.0063	.0063	
e e	-1.2	0	0	.0075	0	.0075	0	0	.0075	.0075	.0075	
tio	-0.8	.0069	0	.0069	.0138	.0069	.0069	.0069	0	0	0	
eva	-0.4	0	.0072	.0144	.0072	.0072	0	.0072	0	0	0	
el	0.0	0	.0148	.0148	0	0	.0074	0	0	0	0	
Standardized elevation, $\boldsymbol{y}_{i}$	0.4	.0068	0	0	0	0	0	0	0	0	0	
rdi	0.8	0	0	0	0	0	0	0	0	0	0	
nda	1.2	0	0	0	0	0	0	0	0	0	0	
Sta	1.6	0	0	0	0	0	0	0	0	0	0	
	2.0	0	0	0	0	0	0	0	0	0	0	
	2.4	0	0	0	0	0	0	0	0	0	0	
	2.8	0	0	0	0	0	0	0	0	0	0	
p	$T^{(t_{\alpha})}$	.0137	.0220	.0436	.0273	.0216	.0143	.0204	.0138	.0138	.0329	
		I										

Note:  $\tau_{\alpha}$ ,  $\tau_{\alpha+1}$ , and  $t_{\alpha}$  are in minutes.

## SUPPLEMENTAL DATA TABLES

Table 10. Sample joint probability mass function of rest periods and elevation of deposition,  $P_{T,Y}(t_\alpha,y_i)$  (Run 17)

						",Y <sub>D</sub> (ta,	i (ma	17)					
τ	α	0	10	20	30	40	50	60	70	80	90	100	110
	α+1	10	20	30	40	50	60	70	80	90	100	110	120
t	ά	5	15	25	35	45	55	65	75	85	95	105	115
	-2.8	0	0	0	0	0	0	0	0	0	0	0	0
	-2.4	0	0	0	0	0	0	0	0	0	0	0	0
	-2.0	0	0	0	0	.0017	.0017	0	0	0	0	0	.0017
·5,	-1.6	.0014	.0014	.0014	.0056	.0042	.0014	0	.0014	0	.0014	.0014	.0042
Ĕ,	-1.2	.0029	.0015	.0017	.0088	.0146	.0044	.0146	.0058	.0029	.0029	.0029	.0015
tio	-0.8	.0085	.0085	.0211	.0197	.0226	.0169	.0127	.0042	.0028	0	.0014	.0028
eva	-0.4	.0059	.0220	.0367	.0323	.0191	.0162	.0044	.0059	.0015	.0029	0	.0029
e]	0.0	.0121	.0351	.0392	.0324	.0189	.0054	.0081	0.	0	.0013	.0013	0
zed	0.4	.0294	.0476	.0420	.0196	.0098	.0042	.0028	0	0	0	0	0
rdı	0.8	.0309	.0382	.0294	.0103	.0059	.0029	0	0	0	0	0	0
Standardızed elevation, $\boldsymbol{y}_t$	1.2	.0312	.0208	.0130	.0065	.0013	0	0	0	0	0	0	0
Sta	1.6	.0256	.0102	.0051	0	0	0	0	0	0	0	0	0
	2.0	.0137	.0051	0	0	0	0	0	0	0	0	0	0
	2.4	.0035	0	0	0	0	0	0	0	0	0	0	0
	2.8	0	0	0	0	0	0	0	0	0	0	0	0
p	$T^{(t_{\alpha})}$	. 1651	. 1904	. 1996	. 1352	.0981	.0531	.0426	.0173	.0072	.0085	.0070	.0131
τ	α	120	130	140	150	200	250	300	400	500	1,000	2,000	3,000
	α+1	130	140	150	200	250	300	400	500	1,000	2,000	3,000	4,000
	α	125	135	145	175	225	275	350	450	750	1,500	2,500	3,500
	-2.8	0	0	0	0	0	0	0	0	0	0	0	0
	-2.4	0	0	.0018	0	.0018	0	0	0	0	0	0	.0018
	-2.0	.0017	.0017	0	0	0	.0017	0	.0017	.0017	.0017	.0017	0
3.	-1.6	0	.0014	.0014	.0014	.0070	.0014	.0014	.0028	.0028	0	0	0
ŗ,	-1.2	.0015	.0015	0	.0029	.0029	0	.0029	.0015	.0015	0	0	0
tio	-0.8	0	.0014	.0014	.0028	0	.0014	.0014	0	0	0	0	0
eva	-0.4	.0015	0	0	0	0	.0015	0	0	0	0	0	0
e]	0.0	0	.0	0	0	0	0	0	0	0	0	0	0
pez	0.4	0	0	0	0	0	0	0	0	0	0	0	0
rdi	0.8	0	0	0	0	0	0	0	0	0	0	0	0
Standardized elevation, $\boldsymbol{y}_i$	1.2	0	0	0	0	0	0	0	0	0	0	0	0
Sta	1.6	0	0	0	0	0	0	0	0	0	0	0	0
	2.0	0	0	0	0	0	0	0	0	0	0	0	0
	2.4	0	0	0	0	0	0	0	0	0	0	0	0
	2.8	0	0	0	0	0	0	0	0	0	0	0	0

Note:  $\tau_{\alpha}$ ,  $\tau_{\alpha+1}$ , and  $t_{\alpha}$  are in minutes.

	£	6.5 7.5	0	0	2	2	1	9	10	12		5	2	0	0	0	0	
	£¢.	4.5 5.5			0.5000					.0834	.1429							040
$Y_D^{=y}j = -2.8$	Class Mark, xg, ft	3.5		1	0.5000 0	- 2000	.3333	.3333	.2000	.2500	1	- 2000		1	:	:		.151
,	5	2.5		:		- 0.5000	6667	3 .3333	0003. 0	3 .3333	7 .5714	0 .4000	0	:	-		-	.358
		1.5						0.3333	.3000	. 3333	.2857	.4000	1.0000					.451
		0.5		-		-	-	-			-	-	-	-	-	-		00.00
		${}^{1}E^{=}y_{z}$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	8.0-	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	$p_{X Y_D}(x_{B} y_j)$
	1	ī, .j	0	0	0	0	0	-	2	2	-	0	0	0	0	0	0	
		7.5			:		1											
		6.5					;		-			-	-		-		;	
		5.5			:	-	:	1					-	1	:	-		
5.2	x8, ft	4.5			1			-	-			1			:		!	
$Y_D^{=y}j = -3.2$	Class Mark, $x_{\mathrm{g}},$ ft	3.5		!		-								:				
	J	2.5		!	-		-		1.0000	:	.5000	:	-	:	!			.348
		1.5		-				1.0000		:	.5000	1.0000		;	:			.652
		0.5		-	-	-	-		-	;	-	-	-	:	-	-	-	0.000
		.2 6 2 7 .	1	-2.8		-2.0	-1.6	-1.2	8.0-	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	$p_{X Y_{2}}(x_{\mathfrak{g}} y_{j})$

Table 13. Sample conditional probability mass function of step lengths,  $p_{X|X_{\rm B},Y_{\rm D}}(\kappa_{\rm B}|y_{\rm L},y_{\rm J})$  (Run 4A)

				D 3						
Y			CI	Class Mark, $x_{ m g}$ , ft	$x_{B}, \; ft$				ş	
.7 e 3	0.5	1.5	2.5	3.5	4.5	5.5	6.5	7.5	"t,j	
-3.2	-	i	•	:		-	1		0	
-2.8			:					:	0	
-2.4	-	:	0.5000	0.5000				:	2	
-2.0	-		.2000	.4000	0.4000				ß	
-1.6	-	0.1818	.2728	.3636	.1818			1	11	
-1.2		.2778	.2778	.3333	.1111			-	18	
8.0-		.2333	.4666	. 2668	.0333		:		30	
-0.4	-	.3421	.2632	.3421	.0526		:	-	38	
0.0	0.0294	.3529	.4706	.1471	-		-	-	35	
0.4	-	.3500	.5000	.1500		;		:	20	
8.0	-	.3077	.6154	-	.0769			-	13	
1.2	:	1.0000		-	:		į	-	2	
1.6		1.0000	:	:		-			-	
2.0	-	:	-		:		!	:	0	
2.4									•	
$p_{X Y_D}(x_B y_J)$	.004	.450	.364	.155	.027					$p_X _Y$

The class widths for step lengths are equal to 1 foot for all 8.

Note: The class widths for step lengths are equal to 1 foot for all 8.

YE"Y.			-	,					
18-81.			Ü	Class Mark, xg, ft	xB, ft				
-3.2	0.5	1.5	2.5	3.5	4.5	5.5	6.5	7.5	ī, j
		1							٥
-2.8	:			:	;	1.0000			-
-2.4	:		0.2500	0.2500	}	.2500	!	0.2500	4
-2.0		-	.1538	. 3077	0.3078	6920.	0.1538		13
-1.6	:	0.1905	.2857	.1429	.2380	.1429		:	21
-1.2		.1795	.3590	.2820	.1282	.0513		1	39
8.0-	:	.2154	.4461	.2154	.1027	.0154			99
-0.4	0.0361	.2771	.3735	.2169	.0603	.0120	.0241	1	83
0.0	.0135	.3649	.4189	.1622	:	.0405			74
0.4	:	.3958	. 3958	.1458	.0208	.0418		-	48
8.0	0690.	.3793	.4483	-	.1034	-		-	53
1.2		.5000	.3750	1	.1250			:	∞
1.6		1.0000	:		-	:	:		7
2.0		-		1	:				0
2.4					-			:	0
$p_{X Y_D}(x_B y_j)$	.018	.388	.372	.122	.071	.025	.004	000.	1

Note: The class widths for step lengths are equal to 1 foot for all 8.

16. --

Run 4A)			ī, j	2	∞	18	40	75	136	217	264	263	210	124	46	6	-	0	;
Sample conditional probability mass function of step lengths, $p_{X Y_{p_r},Y_n}(x_{\pmb{\beta}} y_i,y_j)$ (Run 4A)			7.5		;				0.0073	.0046	.0038	-	1						.001
$p_X _{Y_F,Y_D}^{(x)}$	a a		6.5			0.0555	-	:	.0073		!	9200.	.0048				1	:	.002
lengths,			5.5		0.2500	.1111	.1000	9990.	.0367	.0230	.0151	9200.	.0095	.0081	.0217	:	}	;	.016
on of step	2	xg, ft	4.5	1	0.2500	.2222	.2000	.1600	9560.	.0461	9090	.0380	.0333	.0403		;	:	:	.042
iass functi	$Y_D = y_j = -1.2$	Class Mark, xg,	3.5	0.5000	.1250	.3333	.2750	. 2932	.2307	.2812	.2083	.1255	.0904	.0564	.1302	.2222			.150
obability n	7	5	2.5	0.5000	.3750	.1111	.3500	.3600	.3675	.3550	.3258	.3764	.3284	.2821	.1519				. 282
itional pr			1.5	1		0.1666	.0750	.0931	.1323	.2443	.3144	,3384	.4474	.4916	,6293	.5555	1.0000		.420
omple condi			0.5				-	0.0267	.0220	.0461	.0720	.1064	.0857	.1209	.0651	.2222		-	.087
Table 16. Sc		ì	.1 e 9	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	$P_{\chi Y_D}(x_{\mathfrak{g}} y_j)$
Run 4A)				-	Ŋ	13	28	46	82	128	152	146	108	09	23	4		0	:
( ( , y, y, )			7.5			-			0.0122	8200.	9900.	-	-					:	.002
$p_X _{Y_E,Y_D}^{(x)}$			6.5	1		0.0769		.0216	.0122	-	9900.	.0068	.0093	-	-		-		.005
, lengths,			5.5		0.2000	.0768	.1072	.0652	.0366	.0391	.0197	.0205	.0093	.0167	.0435		-		.025
on of step	9.	$x_{B}$ , ft	4.5	1	0.4000	.3077	.3214	.1522	.1341	.0391	.0395	.0274	.0278	9990.	.0870		:	-	.054
nass functi	$Y_{D}^{=y}j^{=}-1.6$	Class Mark, $x_{eta}$ , ft	3.5	1.0000	. 2000	.3847	.3571	. 2609	.2804	.2890	.2303	.1233	.1204	.0333	.0435	.2500			.153
obability m	,	ט	2.5	1	0.2000	.1539	.2143	.4131	.3779	.4218	.3158	.3972	.3148	.3501	.0870	-		1	.300
itional pre			1.5	1		-	:	0.0870	.1341	.1953	. 3289	.3630	.4908	.4167	.7392	.5000			.403
Sample conditional probability mass function of step lengths, $p_{X Y_{E},Y_{D}}(x_{\mathbf{B}} y_{\mathbf{I}},y_{\mathbf{J}})$ (Run 4A)			0.5	-	-	-	-	-	0.0122	.0078	.0526	.0616	.0278	.1167	-	.2500	-	:	.058
Table 15. S		î A	-E 3:	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	$p_{X Y_D}(x_{g} y_j)$

Note: The class widths for step lengths are equal to 1 foot for all B.

Table 18. Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y_D}(x_{\mathbb{B}}|y_{\mathbb{E}},y_{\mathbb{F}})$  (Run 4A) Sample conditional probability mass function of step lengths,  $p_{X|Y_{E},Y_{D}}(x_{g}|y_{t},y_{j})$  (Run 4A) 17. Table

Y	C1:	2.5	1	0.3333	.2857	.4156	.3931	.4453	.4043	.3459	.3134	.2515	.1503	.1783	.1538			.264	,
		1.5	0.3333	.1667	.1428	6060.	.1241	.1396	.2861	.3835	.4312	.4676	.5654	.5581	. 3846	1,0000		. 439	
		0.5	-	1		0.0130	.0207	.0189	.0331	.0860	.1443	. 2082	.2157	. 2016	.3846	:		.154	
	ì	$^{1}E^{-3}i$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	$p_{X Y_D}(x_{\beta} y_j)$	
	1	$\vec{i}, \vec{j}$	2	O1	56	62	114	210	336	427	440	358	213	91	16	2	0		
		7.5						0.0048	.0030	.0023		-	-		-	:		.001	
		6.5			0.0385	!	:	.0048	-		.0045	.0028				:		.001	
		5.5		0.2222	.0770	.0322	.0351	.0288	.0149	.0047	.0045	.0028	.0047	.0110	:			800°	
~	ig, ft	4.5		0.2222	.2695	.2737	.1403	. 1009	9650.	.0491	.0318	.0195	.0281	:	:		;	.037	
$Y_D=y_j = -0.8$	Class Mark, xg, ft	3.5	0.5000	.2222	.2310	.2254	.3333	. 3216	.2771	.1895	6660.	.0725	.0516	.0550	.0625		:	.126	
<sup>7</sup> z	C18	2.5		0.1111	.2310	.3542	.3596	.3792	.3397	.3416	.3564	.3153	.2204	.1870	.0625	;		.275	
		1.5	0.5000	.2222	.1540	.1127	.1140	.1440	.2593	.3440	.3723	.4213	.5487	.6490	,6250	1.0000		.440	
		0.5				-	0.0175	.0240	.0477	.0679	.1293	.1646	. 1454	0660.	.2500	}	:	.112	
		$^{I}_{E}$ $^{z}y_{z}$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	$P_{X Y_D}(x_{B} y_j)$	

Note: The class widths for step lengths are equal to 1 foot for all  $\boldsymbol{\beta}.$ 

35 35 77 145 265 423 558 603 509 306 129 26 .0018 .0038 .0024 1 , ; .0038 .0033 .0020 0.0286 .0350 .0357 .0390 .0345 .0151 .0018 .0017 .0020 8200. -0.3333 .1667 .2000 .1688 .1103 .0591 .0484 .0199 : : Class Mark,  $x_{g}$ , ft 3.5° 4.5  $Y_D = y_j = -0.4$ 

The class widths for step lengths are equal to 1 foot for all 8.

.001

Sample conditional probability mass function of step lengths,  $p_{X|X_{p},Y_{p}}(x_{B}|y_{\ell},y_{j})$  (Run 4A)

20.

Note: The class widths for step lengths are equal to 1 foot for all 8.

Table 19. Sample conditional probability mass function of step lengths,  $p_X|_{Y_{S,D}}(x_{\mathbf{k}}|_{y_{i}},y_{j})$  (Run 4A) Tabla

		ï.,j	2	13	41	06	163	285	420	548	999	701	455	223	42	4	•	
		7.5						0.0035					:	:	:			000.
ે ચ		6.5			0.0244				.0024	.0018	.0030			:	;	:	1	.001
		5.5		1	0.0488	.0222	.0061	.0105	.0048	.0055	1	.0014	}	}	}	}	-	.002
	β, ft	4.5		0.1538	.1463	.1444	.0920	.0807	.0381	.0182	.0120	.0071	9900.	.0045	:	:	-	.020
$Y_D = y_j = 0.4$	Class Mark, xg, ft	3.5	1.0000	.3846	.3171	.3111	.3190	.2737	.2071	.1332	.0646	.0314	.0264	.0179		-		.085
$I_{L}$	CIE	2.5		0.3846	.4146	.4222	.4601	.4456	.4310	.3887	. 2973	.1498	.1099	.0852	.1190	1	-	.237
		1.5		0.0769	.0488	6880.	.1104	.1789	.3024	.4197	.4850	.4864	.4835	.4350	.3571	.5000	:	.413
		0.5			!	0.0111	.0123	0020	.0143	.0328	.1381	.3138	.3736	.4574	.5238	.5000	-	.242
	>	$I_{E}^{-}y_{z}$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	8.0-	-0.4	0.0	0.4	0.8	1.2.	1.6	2.0	2.4	$P_{X Y_D}(x_B y_j)$
	E	1,1	7	14	39	89	167	596	457	902	725	929	403	187	39	4	0	
		7.5	-	!	-	-	!	0.0034	.0022	;	-	-			:	:		000.
<b>1</b>			1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		0.0256		!	0.0034	.0022	.0017	.0028							.000000
1		7.5				•	•			•	•		•		•	·		
à	zβ, ft	6.5 7.5	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		0.0256			!	!	.0017	.0028				!			.001
$^{2}y_{j}^{2}=0.0$	ass Mark, $x_{f B}$ , ft	5.5 6.5 7.5	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	0.0714	.0769 0.0256	.0337	.0180	6910.	9900.	.0033 .0017	.0014 .0028	.0032	.0025					.004 .001
$Y_D = y_j = 0.0$	Class Mark, $x_{\mathrm{B}}$ , ft	4.5 5.5 6.5 7.5		0.1429 0.0714	.2051 .0769 0.0256	.1573 .0337	.1138 .0180	.0676 .0169	.0416 .0066	.0281 .0033 .0017	.0152 .0014 .0028	.0096 .0032	.0074 .0025	.0053				.021 .004 .001
$Y_D = y_j = 0.0$	Class Mark, $x_{eta}$ , ft	3.5 4.5 5.5 6.5 7.5	1.0000	.2857 0.1429 0.0714	.2308 .2051 .0769 0.0256	.3033 .1573 .0337	.3174 .1138 .0180	.2939 .0676 .0169	.4026 .2144 .0416 .0066	.1256 .0281 .0033 .0017	.0648 .0152 .0014 .0028	.0399 .0096 .0032	.0298 .0074 .0025	.0321 .0053	. 0256			.092 .021 .004 .001
$Y_{D}=y_{j}=0.0$	Class Mark, $x_{oldsymbol{eta}}$ , ft	2.5 3.5 4.5 5.5 6.5 7.5	1.0000	0.4286 .2857 0.1429 0.0714	.3846 .2308 .2051 .0769 0.0256	3933 .3033 .1573 .0337	.4251 .3174 .1138 .0180	.4426 .2939 .0676 .0169	.4026 .2144 .0416 .0066	. 3179 . 1256 . 0281 . 0033 . 0017	. 2938 .0648 .0152 .0014 .0028	.2093 .0399 .0096 .0032	.1315 .0298 .0074 .0025	.1176 .0321 .0053	.0769 .0256			.238 .092 .021 .004 .001

The class widths for step lengths are equal to 1 foot for all  $\beta.$ 

Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y}(x_B|y_E,y_j)$  (Run 4A)

Table 21.

																	$p_X _{X_L}$
E	i, j	2	11	36	80	137	225	307	391	468	512	470	232	45	4	0	
	7.5			;		-	-	:		-			:			!	
	6.5			0.0278			1	-					:		-		000
	5.5	:	-	0.0278	.0250	.0073	.0178	.0065		.0021		:			:		.002
ß, ft	4.5		0.1818	.1389	.1000	.0511	.0622	.0261	.0205	.0171	.0039		;		;		.015
Class Mark, $x_{eta}$ , ft	3.5	1.0000	0.3636	. 3889	.3500	.2847	. 2444	.2020	.1253	.0684	.0293	.0149	.0062		:		080
Cla	2.5		0.3636	.3889	.4375	.5255	.4444	.4658	.4246	.3162	.1660	9920.	.0310	.0444	-	:	. 243
	1.5	*****	0.0909	.0278	.0875	.1314	.2267	.2932	.3862	.4936	.5742	.4149	.2755	.3555	.2500	!	.403
	0.5		-		-	!	0.0044	. 0065	.0435	.1026	. 2266	.4936	.4056	0009.	.7500	;	. 257
) = A	$^{\dagger}E^{-g}i$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	$p_{X Y_D}(x_{\mathcal{B}} y_{j})$

Note: The class widths for step lengths are equal to 1 foot for all 8.

Note: The class widths for step lengths are equal to 1 foot for all  $\beta$ .

.002

.013

.086

.250

.407

.242

Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y_D}(\pi_{\textbf{k}}|y_{z},y_{z})$  (Run 4A) 24 51 88 139 139 217 227 253 276 284 .0114 .0040 .0196 0.0417 .0144 .0588 0.1250 .0228 .0576 .0387 .0184 .0158 .0036 4.5 Class Mark,  $x_{\beta}$ , ft  $Y_{D}=y_{j}=1.2$ 0.5000 .3750 .2446 .1521 .0042 .5000 . 2955 .2210 .0751 .0362 .0141 3.5 0.5000 .4583 .5098 .4332 .3399 .2246 .0915 .0336 .0435 .5000 .4676 .4420 .5227 .2086 .3825 .2500 0.0980 .1364 .2928 .5257 .4930 .3067 .6051 .6555 .0138 0.0114 .0072 .0055 .0395 .1304 .4014 -0.5 Table 22. -2.0 -3.2 -1.6 0.4

Sample conditional probability mass function of step lengths,  $p_X|_{Y_E,Y_D}(x_b|_{\mathcal{Y}_z,\mathcal{Y}_y^*})$  (Run 16)

Fable 26.

Class Mark,  $x_{eta}$ , ft  $Y_D = y_j = -3.2$ 

The class widths for step lengths are equal to 1 foot for all 8.

ςŽ
24.
Table 24
(Run 4A)
$ Y_E,Y_D(x_B y_i,y_j)$
lengths, $p_{X\mid Y_{E},Y}$
unction of step
probability mass f
Sample conditional
Table 23.

		#;,j	0	0	-	2	4	4	4	4	4	4	4	4	4	4	0	
		7.5		1			-		:				:			:	;	
3 2		6.5		;	!	-	:	;	;	:		;	;	;	;		-	
		5.5		;	1.0000	.5000						;	-	;	;		;	.003
	ß, ft	4.5		;	1		0.2500	.2500	.2500	.2500	.2500	-			. ;	:	:	.149
$Y_D=y_j=2.0$	Class Mark, xg, ft	3.5		;	;	0.5000	.2500			;		. 2500	. 2500	.2500	;	:	:	.160
$^{1}D^{-1}y_{j} = 2.0$	Cla	2.5					0.5000	.7500	.7500	.2500	.2500		-		.2500	-	:	.137
		1.5	-	-		:	!	-		0.5000	.5000	.7500	. 7500	.2500		.2500	:	.374
		0.5		1 1			:	1		:	-	•		0.5000	.7500	.7500	:	771.
		1 E 3.	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	$P_{X Y_D}(x_{B} y_j)$
			0	4	13	20	33	44	48	53	26	63	29	89	46	4	0	1
		7.5		-								-		:	:			
		6.5		:		-				:	1		-	:	;			
		5.5			0.0769	.0500	.0303	.0227		1			-			:		.003
	,8, ft	4.5	1		0.1538	;	.0303	.0455	.0625	.0566	.0357		-	-				.022
$Y_D^{=y}j = 1.6$	Class Mark, $x_{\mathrm{g}}$ , ft	3.5		0.2500	.2308	.4000	.3030	.2500	.2083	.1509	.0357	.0317	.0299	.0147		;		.083
$\chi_L$	C12	2.5		0.7500	.4615	.4500	.4848	.4773	.4583	.4151	.4643	.3333	.1045	.0147	.0435	:		.284
		1.5		-	0.0769	.1000	.1515	.2045	.2708	.3585	.4286	.5556	.5821	.3971	.1304	.2500		.395
		0.5	1	:		:	-	:	-	0.0189	.0357	.0794	.2836	.5735	.8261	.7500		.213
	V = 17	1 8 3 1 1 2 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	$P_{X Y_D}(x_{\mathrm{g}} y_{j})$

Sample conditional probability mass function of step lengths,  $p_{X|Y_{E},Y_{D}}(x_{g}|y_{z},y_{j})$  (Run 16) 25.

The class widths for step lengths are equal to 1 foot for all 8.

		0.5		1 1 1 1 1		****			;			-			1			-	-	,	-	1		
		$^{I}_{E}^{=}y_{i}$	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	4.0-	0.0		0.4	8.0	1.2	1.6	2.0	2.4	8 6	3.2	3.6	$P_{X Y_D}(x_{\mathcal{B}} y_j)$	
		٦,٠٠٠	0	0	0	0	0	0	0	1	,	1		-		-	0	0	0	0	. 0	0		
		7.5									1	:		-			1	-				:		
		6.5	;	1	1				1	;	1	:		1 1 1 1 1 1	1		-	:				:		
		5.5	;		1 1 1 1 1			1	1	,		-		1	,			:			;			
s c	zg, ft	4.5	:		,	1 1 1	:		:			-		1 1 1	1	:		:	:		;	-		
$Y_D = y_j = -3.6$	Class Mark, $x_{\mathrm{g}}$ , ft	3.5		:	:	1	:	1	1	1.0000	1.0000					!		:		:		:	. 292	
I	C1	2.5		-		-	:	;				1.0000	,	1.0000	1.0000	-	-	-			:	;	.583	
		1.5	1	!	!	-	-	:		-	:	i		;	1 0	1.0000	1	-	;			:	.125	
		0.5	-				:	:			!	-		-	:	:	-	:	:					
	V = 1/	2 G H	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	•		٥.٠	7.1	1.6	2.0	2.4	2.8	3.2	3.6	$P_X _{Y_D}(x_{\mathcal{B}} _{\mathcal{Y}_{\mathcal{J}}})$	

The class widths for step lengths are equal to 1 foot for all B.

0.5000 .5000 0.5000 .5000 1.0000 0.5000 .5000 .292 1.0000

The class widths for step lengths are equal to 1 foot Note:

Sample conditional probability mass function of step lengths,  $p_{X|Y_{E},Y}(x_{B}|y_{L},y_{J})$  (Run 16)

Table 28. Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y_D}(\pi_{\beta}|y_{\hat{z}},y_{\hat{z}})$  (Run 16) Table 27.

	3	ř., j	0 0 1 4	10 118 21 30 31	28 22 7 3	0000	;	
		7.5		0.1000			.004	
		6.5		0.1111			.012 for all 8.	
		5.5	0.2500	.2000 .1111 .0952 .1667			.080 to 1 foot	
	rg, ft	4.5	1.0000	.3000 .4444 .2381 .1667	.0714		.142 are equal	
6- 7	Class Mark, $x_{eta}$ , ft	3.5	0.2500	.4000 .3333 .4762 .3000	.2857		.256	
,	CI	2.5		0.1429	.4643 .4091 .5714 .6667		.332 ths for ste	
		1.5		0.0333	.1786 .3182 .2857 .3333 1.0000		.154 class widt	
		0.5			0.0455		.020 .154 .332 .256 .142 .080 .012 Note: The class widths for step lengths are equal to 1 foot for all	
	À	- B-87.	-3.6 -3.2 -2.8 -2.4	-1.6 -0.8 -0.4 0.0	4.00.8 4.00.1.1.2 6.00.1.00.00.00.00.00.00.00.00.00.00.00.0	4.8.5.6	$P_X _{Y_D}(x_{g} _{\mathcal{Y}_{\mathcal{J}}})$	
	E	i, , j	0 0 1 8 4 4 3 3 1	9 6 8 7 6	6 / K 0 0	0000		
		7.5		29		1111	.007	
				0.1667				
		6.5		0.1429			ei.	
		5.5 6.5			, , , , ,		.044 .025 to 1 foot for all 8.	
	$x_{eta},$ ft			0.1429			.044 .025 to 1 foot for all 8.	
D - J	ass Mark, $x_{eta}$ , ft	5.5		7 0.1429 0.2222 0.1210 1.1111			.044 .025 to 1 foot for all 8.	
p _ 3	Class Mark, $x_{B}$ , ft	4.5 5.5	1.0000 1.0000 1.0000	.5000 0.1429 .2887 0.1429 .2500 1250 0.2222			.044 .025 to 1 foot for all 8.	
C_O	Class Mark, $x_{eta}$ , ft	3.5 4.5 5.5	1.0000 1.0000 0.2500 .7500	. 3333 . \$0000	.2222 .1111		.044 .025 to 1 foot for all 8.	
C _ D	Class Mark, $x_{eta}$ , ft	2.5 3.5 4.5 5.5	1,0000		.5556 .2222 .1111		.260 .323 .154 .044 .025	

Sample conditional probability mass function of step lengths,  $P_X|_{Y_B,Y_D}(x_{\parallel}|y_{\perp},y_{\downarrow})$  (Run 16) Table 29.

Note: The class widths for step lengths are equal to 1 foot for all 8.

Note: The class widths for step lengths are equal to 1 foot for all 8.

Table 30. Sample conditional probability mass function of step lengths,  $p_{X|X_{E},Y_{D}}(x_{B}|y_{L},y_{J})$  (Run 16)

, m			เรื่อ	Class Mark, $x_{eta}$ , ft	$x_{B}$ , ft				
*E-91	0.5	1.5	2.5	3.5	4.5	5.5	6.5	7.5	"i, j
-3.6		-	-	1.0000	:		1		-
-3.2	1		-	1.0000	1	1			7
-2.8			1 1 1 1	.2500	0.5000	1 1 1	0.2500	-	4
-2.4				. 3000	.4000	0.2500	-	0.0500	50
-2.0		-	0.0227	.2045	.4091	. 2955	.0682	-	44
-1.6	:		.0349	.3256	.3256	.2326	8690.	.0116	98
-1.2			.1016	.3281	.3281	.1563	.0859		128
-0.8	0.0065		.1806	.4258	.2129	.1484	.0258	-	155
-0.4	-	0.0829	. 2818	.3481	.2099	.0663	.0110	1	181
0.0	.0109	.1739	.3233	.2989	.1539	.0217	.0054		184
4.0	.0124	.1988	.4410	.2857	6250	.0062	1		191
8.0	.0242	.3387	.4839	.1290	.0242	.0081	,	:	124
1.2	.0308	.4615	.4154	.0769	.0154	1	1	-	65
1.6	.0385	.5769	. 3846			1	-	-	56
2.0	* 1 1 1 1 1	.6000	.4000		:	:	!	:	ß
2.4	!	1.0000				1		:	7
2.8				1 1 1 1 1		, ,			o
3.2			1						0
3.6					-	:	-	1	0
$P_{X Y_D}(x_B y_j)$	.012	.200	.316	.258	.139	090.	.014	.001	

Sample cond:	
Table 32.	
e 31. Sample conditional probability mass function of step lengths, $p_{K Y_{E},Y_{D}}(x_{B} y_{L},y_{j})$ (Run 16)	
Table 31.	

Run 16)			"t,j	1 7 7 74	152 229 293 355 376	334 254 151 60 22	80 H H O	:	
y,y,) (			7.5		0.0034			.001	
$ x _{Y_{E},Y_{E}}^{(x)}$	4		6.5	0.1429	.0658 .0437 .0102 .0028			.011	for all ß.
lengths, p			5.5	0.1852	.1776 .1179 .0785 .0423	.0090		.039	to 1 foot
m of step		'8, ft	4.5	0.5714	.3421 .2969 .2389 .1803	.0359		.125	are equal
ss functio	$Y_D=y_j=-0.8$	Class Mark, xg, ft	3.5	1.0000 1.0000 .2857 .3704	.3553 .4148 .3993 .3099	.2006 .1102 .0662 .0167		. 237	p lengths
ability ma	$\gamma_L$	C1a	2.5	0.0370	.0592 .1179 .2287 .3352	.4162 .4016 .3046 .2166		. 299	The class widths for step lengths are equal to 1 foot for all
ional prob			1.5		0.0087 .0341 .1155	.3024 .4173 .5298 .6167	.8750	. 250	class widt
Sample conditional probability mass function of step lengths, $p_{X X_{\mathbb{R}},Y_{\mathbb{R}}}(\pi_{\mathbb{R}} y_{\mathfrak{f}},y_{\mathfrak{f}})$ (Run 16)			0.5		0.0068	.0359 .0551 .0927 .1500	.1250	.038	Note: The
Table 32. Sam		; >	, E_3;	-3.6 -3.2 -2.8 -2.4 -2.0	-1.6 -0.8 -0.4	0.4 0.8 1.2 1.6 2.0	4.8.5.8. 4.8.0.0.0	$p_{X Y_D}(x_B y_j)$	Z
8un 16)		3	1,,j	1 1 5 25 67	128 190 239 281 294	263 200 122 50 50	0 11 12		
14:43) (			7.5	0.0800	.0078			.002	
$ x Y_{E},Y_{D} ^{(x)}$	,		6.5	0.2000	.0547 .0526 .0126 .0107	.0038		.013	foot for all B.
lengths, p			5.5	0.2000	.2031 .1474 .1046 .0498	.0076		.046	
m of step		β, ft	4.5	0.6000	.3125 .3053 .2259 .1922	.0418		.130	are equal
use functio	$Y_D = y_j = -1.2$	Class Mark, $x_{eta}$ , ft	3.5	1.0000 1.0000 .2000 .2800	.3828 .3737 .4184 .3310	.2129 .1000 .0738 .0200		.242	p lengths
oability ma	I I	Cle	2.5	0.0400	.0391 .1105 .2134 .2989	.4677 .4450 .3033 .2800	.2000	.316	ths for ste
tional prot			1.5		0.0105 .0126 .1103	.2395 .4050 .5492 .5800	.8000 1.0000 1.0000	.227	class widt
Sample conditional probability mass function of step lengths, $p_{X Y_p,Y_p}(\mathbf{x}_{\beta} y_j,y_j)$ (Run 16)			0.5		0.0084	.0266		.024	Note: The class widths for step lengths are equal to l
Table 31. Som		i	· E 3 ·	-3.6 -3.2 -2.8 -2.4	-1.6 -0.8 -0.4	0.4 0.8 1.2 2.0	2.4 3.2 3.6	$P_{X Y_D}(x_B y_j)$	_

Table 34. Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y_D}(x_{\beta}|y_{\epsilon},y_{\epsilon})$  (Run 16) Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y_D}(x_{\parallel}y_{\perp},y_{\parallel})$  (Run 16) 33. Table

		1.5		} }	;	;	0.00	.0	90.	91.	.24	.38	.49	.50	.56	99.	. 70	1.00	;	i	.28
		0.5			-	-			0.0027	.0182	.0871	.0837	.1136	.2270	.2877	.2222	3000	****	1.0000	:	.084
		$r_{E^{m}y}$ .	-3.6	-5.2	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	2.8	3.2	3.6	$P_{X Y_D}(x_{\beta} y_j)$
		"t,j	- ·	າ∞	28	81	174	569	341	412	442	396	297	176	6	56	6	~	7	0	
		7.5			-	0.0123	.0115	.0074	.0029				1		!	-				:	.002
		6.5			0.0714	.0617	.0517	.0372	.0147	.0049	.0023		:		:	-		-		!	.010
		5.5		0.1250	.1786	.1358	.1667	.1004	.0674	.0437	.0158	9200.	.0034			:		!	:	;	.037
	'8, ft	4.5	2222 0	. 5000	.3214	.4568	.3161	. 2639	.2346	.1553	.0633	.0253	.0067		:	:	:	:	:		.112
D=9; = -0.4	Class Mark, xg, ft	3.5	1.0000	.3750	. 3929	. 2963	.3908	.4498	.3578	. 2864	.2715	.1843	.0774	.0511		:	;	-	:		.228
7,	Cla	2.5		-	0.0357	.0370	.0632	.1338	.2786	.3617	.3710	.4015	. 3603	.2784	.1571	.1154	;	:	;		. 298
		1.5	: :	:	-		:	0.0074	.0411	.1335	.2240	.3283	.4714	.5511	.5857	. 7692	.8889	1.0000	1		.257
		0.5			:	-	:		0.0029	.0146	.0520	.0530	.0808	.1193	.2571	.1154	.1111		1.0000		950.
	i	Est	-3.6	-2.8	-2.4	-2.0	-1.6	-1.2	8.0	-0.4	0.0	0.4	8.0	1.2	1.0	2.0	2.4	2.8	3.2	3.6	$P_{X Y_{D}}(x_{g} y_{j})$

0.0055

.0331 .0106 .0082 .0023

.1602 .1131 .0438 .0273

.2928 .2191 .2192 .1390

.1215 .1837 .3205 .3736 .2859

.0047

.0163

.1488

.3651 .3218 .2324 .1507

0.0357

0.1250

0.3750

0.1250

Class Mark,  $\pi_{\beta}$ , ft  $Y_D = y_j = 0.0$ 

Note: The class widths for step lengths are equal to 1 foot for all  $\beta.$ 

181 365 365 4439 482 482 317 185 73 27 27 27 Note: The class widths for step lengths are equal to 1 foot for all  $\beta$ . .005 .031 760. .292 00055 0106 0603 11663 2407 2814 4984 5081 5616 6667 7000 00000

Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y_D}(\mathbf{x}_{B}|y_{L},y_{J})$  (Run 16) Table 36. Sample conditional probability mass function of step lengths,  $p_{X|Y_{E},Y_{D}}(x_{B}|y_{L},y_{J})$  (Run 16) Table 35.

		2,3	1 3 8 26 67	146 223 272 319 350	364 331 190 73 27	10 1 0		
		7.5						
		6.5	0.0385	.0137			.002	for all 8.
		5.5	0.1250	.1027 .0448 .0331 .0157			.018	The class widths for step lengths are equal to 1 foot for all 8.
	g, ft	4.5	0.2500	.2466 .2377 .1912 .1003	.0137		.082	are equal
$Y_D = y_j = 0.8$	Class Mark, $x_{eta}$ , ft	3.5	1.0000 .3333 .3750 .3462 .3731	.4726 .4305 .3713 .2947	.0989		.195	ep lengths
$I_L$	Cla	2.5	0.6667 .2500 .1923	.1575 .2646 .3566 .4608	.3352 .2115 .1737 .1096	.0741	. 298	ths for st
		1.5		0.0068 .0179 .0441 .1285	.4478 .4471 .5105 .4932	.6000	.290	class wid
		0.5		0.0200	.1044 .3021 .3105 .3973	.4000	1115	Note: The
		$^{1}E^{=}y_{i}$	-3.6 -3.6 -2.8 -2.4	-1.6 -1.2 -0.8 -0.4	0.8 0.8 1.6 2.0	3.228	$p_{X Y_D}(x_{g} y_{j})$	
		#; , ;	1 3 8 29 81	169 266 337 401 444	448 325 189 73 27	10 1 0		
		7.5		0.0030			000.	
9		6.5	0.0355	.0018			.003	foot for all 8.
		5.5	0.1250	.1775 .0752 .0326 .0150	.0022		.025	-
	$x_{\beta}$ , ft	4.5	0.3750	.2249 .2331 .2166 .1147	.0112		.087	are equal
$Y_D = y_j = 0.4$	Class Mark, xβ, ft	3.5	1.0000 .6667 .2500 .3793	.4497 .4737 .3442 .2993	.1161		.203	ep lengths
4	15	2.5	0.3333 .2500 .1034	.1302 .1992 .3501 .4190	.3326 .2646 .1958 .1233		. 292	ths for st
		1.5		0.0059 .0150 .0504 .1421	.5108 .5108 .4868 .5479	. 7000	.287	Note: The class widths for step lengths are equal to
		0.5		0.0075	.1272 .1723 .2857 .3288	.3000	.103	Note: The
		1 R. A.	6.5.2.2.2.2.2.3.6.2.2.3.2.2.3.2.2.3.3.2.2.3.3.2.2.3.3.2.2.3.3.2.2.3.3.2.3.3.2.3	1.1.6 0.0 0.0	0.0 1.2 2.0 2.0	2.2.2.2. 4.8.5.0.	$p_{X Y_D}(x_{B} y_{j})$	

Ţa	
Sample conditional probability mass function of step lengths, $p_{\gamma \mid \gamma = \gamma}$	$Y_D = y_j = 1.2$
Table 37.	
Ţ	

	*		a là di	7-10
		.6.1	1 1 5 113 31 1142 1142 1145 1175 1175 1197 110 110 110	
		7.5		
		6.5	0.0769	
		5.5	0.0968	
	zg, ft	4.5	0.2000 .3846 .3846 .3973 .3125 .2055 .0098 .0098	
,D_8j = 1.5	Class Mark, xg, ft	3.5	1.0000 6000 3846 3226 4107 4459 4107 5714 2284 1317 0580	
1	CI	2.5	1.0000 1.0000 1.536 1.290 1.290 1.216 1.220 1.232 1.233 1.2415 1.354 1.354 1.354 1.354	
		1.5	0.0205 0.0205 0.023 0.05	
		9.5	0.0195 1.0000 1.0000 1.0000	
	٥	1E-91.	13.6 1.5.2 1.5.2 1.0.6 1.0	CX   X D C B   2 J

The class widths for step lengths are equal to 1 foot for all  $\boldsymbol{\beta}$ .

Sample conditional probability mass function of step lengths,  $p_{X|Y_{E},Y_{D}}(x_{B}|y_{L},y_{J})$  (Run 16) 21 337 552 64 77 77 74 74 74 74 10 10 0.1429 .3333 .3243 .2115 .0938 Class Mark, xg, ft  $Y_D = y_j = 1.6$ 0.1429 0.0625 .7000 0.1084 .2874 .6081 able 38.

The class widths for step lengths are equal to 1 foot for all 8. Note:

.081

.325

.080

Sample conditional probability mass function of step lengths,  $p_{X|Y_E,Y_D}(x_{B}|y_{L},y_{J})$  (Run 16)

Table 42.

Note: The class widths for step lengths are equal to 1 foot for all 8.

Sample condit	
Table 40.	
9. Sample conditional probability mass function of step lengths, $P_X _{Y_E,Y_D}(\pi_{\beta} _{Y_E,Y_J})$ (Run 16)	
Table 39.	

(Run 16)			m, 5, 5	00000	3 0 11 14	16 16 16 16	10 1 0	
(+ K. + K)			7.5					
$p_{X Y}$ $Y$ (2	D - E - D		6.5					
lengths,			5.5		0.3333			.015
on of step		xg, ft	4.5		0.3333 .6666 .3000 .0909			.107
ass functi	$Y_D = y_j = 2.4$	Class Mark, xg, ft	3.5		0.3333 .1667 .4000 .3636	.1875		. 204
bability m	Y	ij	2.5		0.1667 .3000 .5455	.5625 .6250 .2500 .1875		2 .237 .405 .204 .107 .015
tional prol			1.5		0.2857	.2500 .3125 .6250 .6250	.2000	.237
Sample conditional probability mass function of step lengths, $p_{Y Y-Y}\left(x_{k} y_{2},y_{j}\right)$ (Run 16)			0.5			0.1250 11875 3750	.8000 1.0000 1.0000	.032
Table 40. Sam		Þ	18 3;	-3.6 -3.8 -2.8 -2.4	-1.6 -1.2 -0.8 -0.4	0.8 1.2 1.6 2.0	4.8.2.8. 4.8.2.8.	$p_{X Y_D}(x_{B} y_j)$
Run 16)		3	£	7 7 7 7 7	5 9 119 21 27	29 30 32 32 27	10 1 4	:
18: 4; 18;) (			7.5					
$p_{X Y_{x},Y_{n}}(x)$	3		6.5					016
lengths,			5.5		0.2000			
on of step		$x_{\mathrm{g}}$ , ft	4.5		0.2000 .4444 .2632 .1905			7 .239 .367 .206 .095 The class widths for stem lengths are equal to
iass functi	$Y_D = y_j = 2.0$	Class Mark,	3.5	1.0000	.4000 .1111 .4210 .3333	.1379		. 206
ability n	ζ	5	2.5	0.5000	.2000 .4444 .2632 .4286 .5185	.4483 .4000 .2188 .1563		.367
5			1.5		0.0476	.3793 .4588 .4688 .4375	.3000	.239
itional prob			7					1
Sample conditional probability mass function of step lengths, $p_{X Y_p,Y_n}(\mathbf{x}_{\beta} y_i,y_j)$ (Run 16)			0.5			0.0345 .0667 .3125 .4062 .5556	1.0000	.077

Table 41. Sample conditional probability mass function of step lengths,  $P_{X|Y_E,Y_D}(\pi_{\beta}|y_{\ell},y_{j})$  (Run 16)

	xg, ft	4.5					
000	Class Mark, xg, ft	3.5					
_	10	2.5		1.0000			.193
		1.5		1.0000	1.0000		.735
		0.5			1,0000	1.0000	.072
	<b>.</b>	$^{1}E^{-}\theta_{i}$	-3.6 -3.2 -2.8 -2.4	-1.6 -1.2 -0.8 -0.4	0 0 . 8 . 1 . 2 . 2 . 1 . 2 . 2 . 2 . 2 . 2 . 2 . 2 . 2 .	4.2.2.8. 4.2.2.3.	$p_{X Y_D}(x_{\beta} y_j)$
		i j	00000	2222	00000	2 1 0	
		7.5					
		6.5					i
		5.5					
	$x_{eta}$ , ft	4.5		1.0000			.046
6 0	Class Mark, $x_{\mathrm{g}}$ , ft	3.5		.5000 .5000 .5000 .5000			.224
	D C	2.5		00.5000	.5000		.372
		1.5		0.5000	.5000 .5000 .5000 1.0000	.5000	.347
		0.5			0.5000	.5000 1.0000 1.0000	.011
	Y=u.	1. B	-3.6 -3.2 -2.8 -2.4	-1.6 -1.2 -0.8 -0.4	0.8 0.8 1.6 0.0 0.0	4.2.2.2. 4.8.2.2. 6.0.0	$P_{X Y_D}(x_{B} y_j)$

Note: The class widths for step lengths are equal to 1 foot for all g.

Note: The class widths for step lengths are equal to 1 foot for all 8.

Table Sample conditional probability mass function of step lengths,  $p_{X|Y_{g},Y_{n}}(x_{\beta}|y_{\ell},y_{j})$  (Run 17) Table 43.

) (Run 17)			m j	00000	87779	C 4 I I I O	00		
Sample conditional probability mass function of step lengths, $p_{X Y_{n},Y_{n}}(x_{g} y_{\xi},y_{j})$ (Run 17)	<b>a</b>		5.5						all ß.
engths, $p_{X Y}$			4.5		0.3333			. 050	o l foot for
on of step L	4	, xg, ft	3.5		0.3333 .2857 .2857 .4286			.231	The class widths for step lengths are equal to 1 foot for all 8.
mass functi	$Y_D = y_j = -2.4$	Class Mark, $x_{\mathrm{\beta}}$ , ft	2.5		0.3333 .4286 .1428 .1428	.5714		.333	step lengths
probability			1.5		0.1429 .4286 .4286	.4286 .2500 1.0000 1.0000		.356	widths for
conditional			0.5			0.2500		.030	: The class
Table 44. Sample		>	$^{t}E^{=}y_{t}$ :	-3.6 -3.2 -2.8 -2.4 -2.0	-1.5 -0.8 -0.4 -0.0	0.8 0.8 1.2 2.0	2.4	$p_{X Y_D}(x_{\mathcal{B}} y_j)$	Note:
j) (Run 17)			m;,j	00000	00001	N H O O O	00	1	
y, (x   y, y	ž		5.5						all 8.
engths, $p_{X\mid Y}$			4.5					:	o 1 foot for all 8.
on of step l	8	(, π <sub>β</sub> , ft	3.5		1.0000			.368	are equal t
y mass functi	$Y_D = y_j = -2.8$	Class Mark, $x_{eta}$ , ft	2.5			0.5000		.179	step lengths
.42.			1.5			0.5000		.452	widths for
l probabil									-
Table 43. Sample conditional probability mass function of step lengths, $p_{X Y_{B},Y_{B}}(x_{B} y_{\xi},y_{\xi})$ (Run 17)			0.5					:	Note: The class widths for step lengths are equal to

Note: The class widths for step lengths are equal to 1 foot for all  $\beta$ .

Table 46. Sample conditional probability mass function of step lengths,  $P_X|_{Y_B,Y_D}(x_{\beta}|_{\mathcal{Y}_L,\mathcal{Y}_J})$  (Run 17)

			$^{1}D^{-}y_{j} = ^{-1.0}$	0			
>			Class Mark, $x_{eta}$ , ft	$,x_{\beta},$ ft			E
1 E B 2.	0.5	1.5	2.5	3.5	4.5	5.5	ī, j
-3.6	:	-	:	i	:		0
-3.2	:						0
-2.8				-			0
-2.4	:		:	1.0000			
-2.0	:		0.1111	.5556	0.3333		6
-1.6	0.0286		.1429	.4571	.3143	0.0571	35
-1.2	:	0.0395	. 2632	.4342	. 2632		92
-0.8	:	.1651	.3211	.3853	.1284		109
-0.4	:	.1797	. 4531	.3281	.0391		128
0.0	.0465	.2481	.4961	.2016	.0078	:	129
0.4	.0286	.3714	.5429	.0571	!	1	105
8.0	.0147	.5882	.3676	.0294	:		89
1.2	:	. 8214	.1786		:		28
1.6	.1111	.7778	.1111		:	:	6
2.0			:	-	!	:	0
2.4		:	;	;	:	:	0
2.8			!		!	:	0
$P_{X Y_D}(x_{\mathfrak{g}} y_j)$	.016	.317	.372	.223	690.	.003	:
	_						

Note: The class widths for step lengths are equal to 1 foot for all  $\beta$ .

Note: The class widths for step lengths are equal to 1 foot for all 8.

, (Run 17			ī.,j	0	0	н (	∞ ç	7,	113	270	406	480	493	414	267	116	30	2	0	0	:
$\mathcal{L}_{i,Y}(x_{B} y_{i,y})$	<b>a</b>		5.5					,	-	0.0111	.0049	.0021	.0020	;				-	-	:	. 002
ingths, PX Y			4.5		-	1.0000	2069	6007.	.2124	.0778	.0296	.0167	.0081	.0024	.0037			:			.033
m of etep L	8	, x8, ft	3.5		:		0.5000	001+	.4690	.4222	, 3202	.1875	6990.	.0290	.0075	.0172	!	:	:		,156
mass function	$Y_D = y_j = -0.8$	Class Mark, xg, ft	2.5		:		7793	56.65	.2655	. 3963	.4236	.4750	.4381	.3188	.1760	,0948	9990.	:	-,,,,,		. 329
probability			1.5			:			0.0442	.0852	.2094	.2792	. 4037	.5411	.6891	.6810	. 6667	1.0000		:	.402
Sample conditional probability mass function of etep lengths, $p_{X Y_R,Y_R}(x_B y_t,y_j)$ (Run 17)			0.5	1		1			0.0088	.0074	.0123	.0396	. 0811	.1087	.1236	. 2069	. 2667				.078
Table 48. Sample		٨	$F^{BJ}$	-3.6	-3.2	-2.8	-2.4	ò	-1.6	-1.2	8.0-	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	2.8	$p_{X Y_D}^{(x_{\beta} y_j)}$
17)	ļ	!		l																	
.) (Run			ř.,j	0	0	0 •	4 ,	3	84	180	259	298	301	250	159	29	19	2	0	0	
$(x_{\beta} y_i,y_j)$ (Run		1	5.5	0	_		0.2500		.0119 84		259	298	301	250			19	2	0	0	
engths, $p_{X Y_E,Y_D}(x_{\beta} y_t,y_j)$ (Run		E						reto.								_	_				
on of step lengths, $p_{X Y_E,Y_D}(x_{eta Y_U,Y_J})$ (Run	2		5.5	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		1000	0.2500	1840.	6110.	.1389	6250.	.0201	.0033		1 7 2 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1				:	1	.003
mass function of step lengths, $p_{X Y_E,Y_D}(x_{\beta} y_i,y_j)$ (Run	$Y_D = y_j = -1.2$	Class Mark, $x_{eta}$ , ft	4,5 5.5	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		100000000000000000000000000000000000000	1204 0.2500	1010.	.2976 .0119	.4278 .1389	.3745 .0579	.2350 .0201	.1030 .0033	. 0360	.0189			2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2		1 1 5 6 6 1	.042 .003
probability mass function of step lengths, $p_{X Y_E,Y_D}(x_{eta y_{oldsymbol{i}},y_{oldsymbol{j}}})$ (Run	$Y_D^* = y_j = -1.2$		3.5 4.5 5.5	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		110000000000000000000000000000000000000	0.2500 0.2500 0.2500	10th: 100t. 00/t. 0/to:	.3810 .2976 .0119	.3611 .4278 .1389	.3629 .3745 .0579	.4631 .2350 .0201	.4884 .1030 .0033	.4000 .0360	.2138 .0189	.1492		2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2			.173 .042 .003
Sample conditional probability mass function of step lengths, $p_{X Y_E,Y_D}(x_{eta Y_i,Y_j})$ (Run 17)	$Y_D^{=}y_j = -1.2$		2.5 3.5 4.5 5.5			100100000000000000000000000000000000000	1.2500 0.2500 0.2500 0.2500 0.2500 1204 0.2500 0.25	10th: 1001: 00/t: 0/to:	0.0238 .2738 .3810 .2976 .0119	.0611 .3611 .4278 .1389	.1969 .3629 .3745 .0579	.2584 ,4631 ,2350 ,0201	.4884 .1030 .0033	.5080 .4000 .0360	.7233 .2138 .0189	.7612 .1492	.2105	1,0000		2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	.352 .173 .042 .003

Note: The class widths for step lengths are equal to 1 foot for all 8.

Note: The class widths for step lengths are equal to 1 foot for all 8.

Table 50. Sample conditional probability mass function of step lengths,  $p_{X|Y_{E},Y_{D}}(x_{\beta}|y_{\xi},y_{\xi})$  (Run 17)

Sample conditional probability mass function of step lengths,  $P_{X|Y_E,Y_E}(x_8|y_t,y_j)$  (Run 17)

Table 49.

 $Y_D = y_j = -0.4$ 

 $Y_{E}=y_{i}$ 

Class Mark. m. ft				-					a land				
						, u= 4			ciass mark, xg, rt	, π <sub>β</sub> , ιτ			£
2.5 3.5 4.5	5	4.5		5.5	٤٠,٦	F at	0.5	1.5	2.5	3.5	4.5	5.5	, t <sub>2</sub>
111111111111111111111111111111111111111		-			0	-3.6				,			
	•				0	-3.2		1	:				0
1.0000					2	-2.8			0.5000	0.5000			
0.3077 .6154 0.0769		0.0769			13	-2.4	:	:	.3077	.6154	0.0769		13
.4500		.1750		0.0250	40	-2.0	1	0.0455	.4318	.3636	.1591	-	4
.3358 .4552 .1493		.1493			134	-1.6	1	.0725	.4058	.4130	.1087		138
.3714	-	.0794		. 0032	315	-1.2	-	.1683	.4222	.3492	.0571	0.0032	315
.2710	-	.0287		.0041	487	-0.8	0.0120	.2371	.5020	.2271	.0199	.0020	205
.1443	-	.0116		.0017	603	-0.4	.0563	.3772	.4554	.1033	8200.	-	635
.0649		.0063			632	0.0	.1491	. 4574	.3509	.0398	.0028	;	704
.2697 .0275 .0018		.0018			545	0.4	.2210	.5645	.1984	.0161	:		62
	•				357	8.0	.2725	. 6448	.0730	. 0097	:	-	41
	•				159	1.2	. 3968	. 5556	.0423	. 0053		:	18
	•				45	1.6	.5273	.4364	.0364		-		Š
1 1 2 5 6 6 7 6 7 6 7 6 7		!		:	4	2.0	0009.	.4000		:		:	S
	•	-			•	2.4	1.0000		:	;			
		!		-	•	2.8		:	-	:	:	:	0
.310 .141 .024		. 024		.001	1	$p_{X Y_n}(x_{\beta} y_j)$	.164	.406	.297	.115	.017	.001	
			- 1			•							

Note: The class widths for step lengths are equal to 1 foot for all 8.

,113

 $P_{X|Y_D}(x_{\mathfrak{g}}|y_j)$ 

Sample conditional probability mass function of step lengths,  $p_{X|X_p,Y_n}(x_{\beta}|y_{i},y_{j})$  (Run 17) Table 52. Sample conditional probability mass function of step lengths,  $p_{X|X_{p},Y_{n}}(x_{B}|y_{L},y_{J})$  (Run 17)

		1	"i,j	0	0	-	œ	28	6	202	324	410	469	402	25	174	200	62	Ŋ	-	0	***************************************		
a d d a d a d a d a d a d a d a d a d a			5.5		!	-	:			!	-	:	-			1		:		:				а11 в.
310			4.5		1		0.1250	.1071	.0412	.0446	.0123	.0049	!			!	1		-	!	:		.011	l foot for
		, xg, ft	3.5		1	:	0.3750	.3929	.5155	.3416	.1759	.0732	.0299	0062	2000	. 0023	:		:	:			.102	The class widths for step lengths are equal to 1 foot for all B.
:	$r_D=y_j=0.8$	Class Mark, $x_{\mathrm{g}}$ , ft	2.5			1.0000	.3750	. 3929	.3814	.4802	.5741	.5195	.3454	1284	1 1 1 1 1	0100	0620.	.0161	!		!		. 304	step lengths
			1.5	1			0.1250	.1071	.0619	.1337	.2346	. 3683	.5565	6204	1754	1001	.3/00	.2258	.2000		!		.382	widths for s
			0.5						-		0.0031	.0341	.0682	2360	7077	/0/4.	nena.	.7581	. 8000	1.0000	:		.201	: The class
		Α ν.	. E=3.	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	4 0			7.1	1.6	2.0	2.4	2.8		$P_{X Y_D}(x_{B} y_j)$	Note:
		E	į, j	0	0	-	6	39	123	275	449	570	647	629	425	57	190	9	ın	7	•		1	
· · · · · · · · · · · · · · · · · · ·			5.5		1						0.0022	-	-			:			!	-			000.	all 6.
1 .			4.5		:	:	0.1111	11795	.0732	.0473	.0178	.0088	1					:	:	:	******		.015	to I foot for all 8.
		, xg, ft	3.5	:	-	:	0.5556	. 3077	,4636	.3600	.2004	.0825	.0294	0125	2700	. 004/	:	:		!			.107	
,	$Y_D^{=}y_j=0.4$	Class Mark, $x_{\mathrm{\beta}}$ , ft	2.5		-	1.0000	.2222	.4615	.3984	.4545	. 5367	. 4684	. 3338	1246	24.70	.0035	.0404	.0167	1	-			.292	step lengths
			1.5	1	,	*****	0.1111	.0513	,0650	.1382	.2405	.4088	.5008	2463	1000	59/5	. 4495	. 2667	.4000		:		.390	s widths for
			0.5		:			-	-	-	0.0022	.0316	.1360	2067	7000	. 3553	.5101	.7167	0009	1.0000	-		. 196	Note: The class widths for step lengths are equal
		5	1 E - 9 2	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	8.0	4.0-	0.0	•	,	8.0	1.2	1.6	2.0	2.4	2.8		$p_{X Y_D}(x_{\mathcal{B}} y_{\dot{\mathcal{J}}})$	Note

Sample conditional probability mass function of step lengths,  $p_{K}|_{Y_{E},Y_{D}}(x_{E}|y_{L},y_{J})$  (Run 17)

53.

-	, s	2,03	0	0	0	4	17	45	66	155	201	235	245	247	200	62		-	0	-	,
		5.5		;		-	!	-		-	1 1 1 1	!	-	;		:					
		4.5		:		0.2500	.1176	.0444	. 0303	.0065	.0050		-	1	:	:			-	010.	
	Class Mark, $x_{\mathrm{g}}$ , ft	3.5	-	-		0.2500	.2941	.4222	.2828	.1419	. 0945	.0340	.0082	.0040		:	;		-	. 092	
	Class Mar	2.5		-		0.5000	.5882	.4889	.5455	,6258	.5274	.3447	.1429	.0526	.0200	.0161				. 327	
		1.5		1		-	-	0.0444	.1414	.2258	.3682	.5660	. 7143	.5223	.2300	.1452	.2000	-	-	. 384	
		0.5	-		-	-	-	;		-	0.0050	.0553	.1347	. 4211	.7500	. 8387	. 8000	1.0000	-	.187	
	A	.7 B 9.7.	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	2.8	$p_{X Y_{D}}(x_{B} y_{j})$	۵

Note: The class widths for step lengths are equal to 1 foot for all 8.

Table 54. Sample conditional probability mass function of step lengths,  $p_X|_{Y_p,Y_n}(x_{\beta}|_{\mathcal{Y}_p,\mathcal{Y}_j})$  (Run 17)

, (Run 1			"i.j	0	0	0	2	2			n 4	×o	œ	10	;	7	12	12	12	ß	-	4 (	9	:	
y=18   y, x	, ,		5.5		:	:	:	:				:	:						-					-	all 8.
engths, $p_{X Y}$			4.5		:	: : :	-	:				:	:::::	:		:			:						o l foot for
on of step 1		, xg, ft	3.5		:	:	0.5000	.3333	1111		0007	1720		:			:		:				:	.058	are equal t
mass functi	$Y_D=y_j=2.0$	Class Mark, xg, ft	2.5				0.5000	. 6667	. 6667		0000	05/5.	.3750	. 4000		2797	.0833		*	1 1 1 1	,			.304	step lengths
probability			1.5	,,,,,	:					0000	0.200	. 2000	.6250	0009	6	2019	. 5000	. 5000	.1667	,				.497	widths for
Sample conditional probability mass function of step lengths, $p_{X Y_{p_i},Y_n}(x_{\beta} y_{z^i}y_j)$ (Run 17)			0.5								:	:					0.4167	2000	.8333	1.0000	0000			.141	Note: The class widths for step lengths are equal to 1 foot for all
Table 55. Sample		Þ	$^{1}E^{=}y_{2}$	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6		7:1-	. O-	-0.4	0.0	•	4.0	8.0	1.2	1.6	2.0	2.4		9.7	$p_{X Y_D}(x_{\beta} y_j)$	Note
, (Run 17)			9																				1		
		3	2,3	0	0	0	7 1	`	15	37	7.7	0.9	5 6	98	84	ď	3 6	Ç (	79	ı,		٥		:	
E, Y (x   y, y,		3	5.5	0		0	_		15	37				08	84				79			0		:	. all β.
engths, $p_X _{Y_E,Y_D}^{(x,y)}(x_{eta} _{y_{m{i}}},y_{m{j}})$		2		0	_	1 1 1 1 1 1 1 1	_	:					1								1	0			o 1 foot for all 8.
on of step lengths, $p_{X Y_{\mathbf{E}},Y_{D}}(x_{\mathbf{B}} y_{t},y_{o})$			5.5	•			:	.1429		.0220		0146	C+10.							:				:	-
mass function of step lengths, $p_{X Y_{E},Y}(x_{eta} y_{\mathfrak{t}},y_{\mathfrak{d}})$	$Y_{E}=y_{j}=1.6$	Class Mark, $x_{\beta}$ , ft	4.5 5.5				0.5000	.2857 .1429	.5333	2162 .0270		3810	C\$10.	.0750	0110				******			111111111111111111111111111111111111111			-
probability mass function of step lengths, $p_{X X_{E},Y}(x_{eta} y_{L},y_{c},y_{c})$	$Y_E = y_j = 1.6$		3.5 4.5 5.5				0.5000 0.5000	.2857 .1429	.5333	6486 2162 .0270	1001	3410 0311	CA10. CETT. CA46.	.3500 .0750	1667 0119		10000 TANO.	.0235	******		1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1			.107 .010	-
Sample conditional probability mass function of step lengths, $p_X _{Y_E^{1,Y}}(x_b y_L,y_j)$ (Run 17)	$Y_E=y_j=1.6$		2.5 3.5 4.5 5.5				0.5000 0.5000	0.5714 .2857 .1429		0.1081 6486 2162 .0270	1001	1910 Carr 2000 Carr	.4203 .4493 .1159 .0145	.3500 .0750	7147 2717	110. COLL	1980. 2886.	. 3412	.0161	.2000	# # # # # # # # # # # # # # # # # # #			.323 ,107 .010	Note: The class widths for step lengths are equal to 1 foot for all 8.

Sample conditional probability mass function of step lengths,  $p_{X\mid Y_{\Sigma},Y_{D}}(x_{\beta\mid Y_{\Sigma},y_{J}})$  (Run 17) Table 56.

		Class Mark	, x <sub>β</sub> , ft			
0.5	1.5	2.5	3.5	4.5	5.5	ı, j.
-	:	:	1			o
	;	-		:	:	0
	!					0
	:			:		0
:		-	:	-		0
	:		;	į	;	c
	;	1	*****	-	:	o C
-		;		1		
		-	:		-	
-	:	!	* • • • • • • • • • • • • • • • • • • •	:	:	0
;	1.0000	;				
-	1.0000					٠.
1.0000						٠.
1.0000	;			;		٠
1.0000	;	-		-		٠
1.0000	;		;	į		-
	-	:				- 0
.333	.667					
7, 7, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0,	ni deke Con			,		
i. Ine ciass	Widths tor	step lengtns	are equal t	o l foot for	all β.	
	0.5 0.5 0.5 0.000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000	0.5 1.5	0.5 1.5 2.5  0.5 1.0000  1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.00000 1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.00000 1.00000 1.0000 1.00000 1.00000 1.000000 1.000000 1.00000000	0.5 1.5 2.5 3.5  0.5 1.5 2.5 3.5  1.00000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.000000 1.000000 1.000000 1.000000 1.00000000	0.5 1.5 2.5 3.5 4.5  0.5 1.5 2.5 3.5 4.5  1.0000	1.5 2.5 3.5 4.5  1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000

# STOCHASTIC ANALYSIS OF PARTICLE MOVEMENT OVER A DUNE BED

Table 57. Conditional means and variances of step lengths;  $\hat{E}[X|Y_E=y_i, Y_D=y_j]$  and  $\widehat{\text{Var}}[X|X_E=y_i, Y_D=y_j]$  (Run 4A)

$Y_E = y_i$	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4
-3.6									4.189	4.057	4.000	3.934	3.857			
-3.2						3.974	3.038 1.246	2.842 1.417	3.150 1.235	3.647 .101	3.549 .141	3.440 .191	3.283 .302			
-2.8					5.792	4.172 1.570		3.620 1.918		3.239 1.021	3.207 .839	3.195 .941	2.965 .360	2.768 .395		
-2.4			3.636 .373	3.215 .132	4.730 3.957	4.024 1.500			3.409 1.684	3.459 1.357	3.357 1.156	3.309 .966	3.184 .743	3.199 .985	5.346	
-2.0			3.298 .605	3.843 .978	4.214 1.409		3.407 1.133			3.102 .872	3.036 .783	2.935 .665	2.909 .557		4.308 1.503	
-1.6			2.848 .507	2.944 1.114	3.417 1.665	3.239 1.175	3.173 1.176	3.076 .991	2.989 .974	2.956 .779	2.898 .714	2.804 .599	2.773 .555	2.820 .559	3.485 .987	
-1.2		2.861	2.704 .766	2.848 1.009		3.136 1.439		2.976 1.210	2.854 1.046	2.808 .872	2.785 .806	2.717 .726	2.710 .682	2.733 .669	3.175 1.055	
-0.8		2.611 .054	2.478 .392	2.588 .620		2.785 1.137	2.653 1.145	2.583 1.063	2.505 .887	2.498 .759	2.507 .683	2.466 .623	2.503 .603	2.562 .669	2.879 1.155	
-0.4		2.082	2.434 .836	2.506 .796	2.589 1.251		2.455 1.197		2.180 .925	2.170 .772	2.197 .656	2.205 .602	2.275 .563		2.604 1.225	
0.0		1.892	2.557 .609	2.300 .584	2.400 1.011	2.274 1.069	2.221 1.084	2.083 .950	1.929 .854	1.797 .813	1.843 .666	1.916 .581	2.006 .542		2.350 1.296	
0.4			2.312 .943	2.289 .401	2.365 .985	2.159 .986	2.054 .910	1.903 .771	1.728 .723	1.595 .652	1.437 .637	1.521 .495	1.661		2.087 1.192	
0.8			1.692 .010	2.133	2.257 1.076	2.108 1.091	1.933	1.775 .690	1.604 .634	1.439 .550	1.306 .485	1.127 .416	1.244 .367	1.394	1.829 1.106	
1.2				1.731 .020	2.279 .755	2.188 1.127	1.903 .889	1.756 .626	1.581 .608	1.373 .501	1.203 .429	1.036	.870 .301	.968 .297	1.508 1.035	
1.6				1.430	1.518 .241	1.779 1.693	1.687 .942	1.558 .539	1.503 .646	1.288 .424	1.142	.932 .273	.812 .275	.595 .270	1.081 .829	
2.0							1.670	1.377	1.375 .084	1.086 .120	.954 .125	.869 .128	.785 .132	.687 .147	.460 .203	
2.4																
$\hat{\mathbf{E}}[x Y_D = y_j]$		1.194	2.172	2.397	2.577	2.486	2.384	2.257	2.119	2.025	1.972	1.938	1.983	2.056	2.458	
$Var[X Y_D = y_j]$		1.396	1.232	. 884	1.376	1.415	1.277	1.141	1.084	1.035	1.008	.935	.892	.915	1.671	

Note: Upper values are the means, in feet, and lower values are the variances, in feet squared.

SUPPLEMENTAL DATA TABLES

Table 58. Conditional means and variances of step lengths;  $\hat{E}[X|Y_E=y_i, Y_D=y_j]$  and  $\hat{\text{Var}}[X|Y_E=y_i, Y_D=y_j]$  (Run 16)

V -::																		
$Y_E = y_i$	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	2.8	3.2
-3.6					3.655	3.592	3.531	3.469	3.407	3.338	3.263	3.186	3.111					
-3.2					3.442	3.379	3.318	3.496 .115	3.594 .164	3.427 .091	3.102 .026	3.042 .051	2.897					
-2.8			4.600	4.516			4.447 1.179	4.293 .930	4.142 .832	4.027 .868	3.833 1.047	3.696 .983	3.830 .573					
-2.4			4.334		4.548 1.059	4.612 .979	4.621 1.077	4.480 1.053		4.217 .869	4.063 .892	3.905 .913	3.909 .843	3.098	3.051			
-2.0			4.185 .067	4.477 .545	4.623 .809	4.626 .760	4.514 .699	4.437 .752	4.399 .904	4.223 .746	4.083 .728	3.931 .716	4.012 .656	3.591 .1 <b>9</b> 9				
-1.6				4.646 1.396			4.369 .931		4.261 1.067	4.075 .953	3.981 .829	3.834 .790	3.987 .733		3.816 1.163	4.430 .858	4.379	
-1.2			4.390 1.397	4.513 1.076			4.113 .960	4.039 .918	3.942 .978	3.783 .880	3.694 .748	3.562 .705	3.685 .678	3.689 .665	3.568 1.074	3.952 .842	3.243 .904	2.391
-0.8	3.410	3.256	3.919 1.062	4.020 .837			3.726 1.131				3.356 .833	3.287 .783	3.474 .747	3.481 .609	3.519 .840	3.580 .507	2.933	2.112
-0.4	3.197		3.823 1.307								2.927 .869	2.917 .762	3.072 .656	3.116 .524	3.108 .603	3.129 .483		1.913
0.0	2.917	3.802 2.159					2.851 1.028				2.436 .851	2.457 .714	2.638 .585	2.674 .462	2.582 .584	2.657 .502	2.459 .611	1.726
0.4	2.534	3.273 1.598	2.893 .860	2.727 .6 <b>9</b> 6	2.769 .737	2.692 .745	2.559 .745	2.437 .760	2.316 .756	2.155 .733	1.963 .766	1.960 .675	2.174 .516	2.262 .411	2.296 .498	2.421 .424	2.219 .433	1.574
0.8	2.266		2.413 .739	2.317 .576	2.285 .599	2.300 .591	2.207 .629	2.104 .617	1.987 .582	1.846 .558	1.718 .551	1.519 .612	1.684 .465	1.824 .367	1.938 .459	2.116 .358	1.999	1.409
1.2	1.728	1.574	1.431 .213	2.067 .601	2.127 .428	2.058 .481	1.966 .527	1.844 .510	1.713 .468	1.598	1.498 .427	1.388 .410	1.175 .416	1.358	1.544	1.788 .312	1.769 .308	1.197
1.6				2.125 .041	1.808	1.780 .311	1.687 .373	1.628 .371	1.481	1.385	1.297	1.215 .340	1.114	.946 .383	1.195 .390	1.438 .252	1.512	.998
2.0				1.590	1.371	1.811 .160	1.718 .297	1.676 .229	.1507 .242	1.419	1.340 .232	1.269 .234	1.198 .232	1.109 .237	.960 .276	1.054	1.257	.814
2.4						1.734 .020	1.729 .029	1.414	1.269 .067	1.168	1.080 .084	1.014	.949 .083	.879 .083	.802 .084	.672 · .096	.923 .017	.652
2.8							1.604	1.291	1.195	1.096	1.000	.936	.885	.832	.780	.727	.675	.496
3.2							1.284	.971	.875	.777	.681	.671	.565	.512	. 460	.407	. 355	.176
$\hat{\mathbb{E}}[x y_D = y_j]$	2.073	2.512	3.008	2.924	3.097	3.049	2.927	2.825	2.716	2.567	2.473	2.402	2.517	2.500	2.516	2.550	2.308	1.495
$\widehat{\text{Var}[X Y_D = y_j]}$	1.463	3.476	2.290	1.635	1.530	1.512	1.499	1.509	1.547	1.477	1.432	1.381	1.392	1.234	1.252	1.409	1.164	. 368

Note: Upper values are the means, in feet, and lower values are the variances, in feet squared.

## STOCHASTIC ANALYSIS OF PARTICLE MOVEMENT OVER A DUNE BED

Table 59. Conditional means and variances of step lengths;  $\hat{\mathbb{E}}[X|Y_E=y_i,Y_D=y_j]$  and  $\hat{\text{Var}}[X|Y_E=y_i,Y_D=y_j]$  (Run 17)

$Y_E = y_i$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	2.8
-3.2																
-2.8							4.657	3.553	3.157 .113	2.792	2.663					
-2.4				3.984	3.909	4.026 1.057	3.773 .427	3.354 .407	3.188 .414	3.191 .504	3.048 .608	3.206 .491	3.580 .365	3.308 .340		
-2.0				4.118 .395	3.699 .335	3.441 .725	3.319 .575	3.285 .596	3.104 .546	3.051 .531	3.000 .542	3.030 .456	3.223 .360	2.868 .245		
-1.6			3.655 .456	3.693 .320	3.686 .824	3.471 .751	3.320 .623	3.210 .565	3.076 .546	3.056 .475	2.983 .421	2.932	2.979 .208	2.633 .247		
-1.2			3.066 1.202	3.369 .700	3.408 .632	3.117 .611	3.008 .629	2.922 .589	2.818 .552	2.816 .477	2.769 .453	2.704 .411	2.707 .416	2.515 .706		
-0.8			2.656 1.053	3.013 .935	3.017 .733	2.790 .638	2.662 .681	2.574 .610		2.476 .452	2.465 .408	2.441 .342	2.411	2.259 .391		
-0.4			2.516 .833	2.772 .664	2.709 .613	2.461 .583	2.354 .616	2.232 .608	2.141 .529	2.129 .445	2.151	2.187 .354	2.219 .407	1.998 .287		
0.0		3.396	2.878 .223	2.523 .563	2.379 .522	2.153 .496	2.041 .516	1.934 .522	1.792 .502	1.765 .435	1.814 .351	1.877 .312	1.942 .358	1.798 .220		
0.4		2.110 .965	2.131 .344	2.123 .457	2.091 .370	1.896 .360	1.763 .398	1.652 .405	1.513 .382	1.377 .403	1.421 .318	1.521 .260	1.625 .288	1.527 .181	1.624	
0.8		1.108	1.636	1.888	1.846 .237	1.696 .245	1.573 .289	1.464 .279	1.332 .257	1.219 .256	1.069 .272	1.135	1.291 .247	1.396 .182	1.234	
1.2			1.817	1.584 .090	1.621 .220	1.526 .253	1.400 .245	1.280 .248	1.146 .214	1.048	.941 .203	.786 .212	.891 .202	.984 .147	.967	
1.6			1.226	1.285 .106	1.520 .248	1.540 .254	1.289 .155	1.190 .183	1.038	.945 ,168	.841 .161	.725 .157	.545 .161	.639 .133	.753	
2.0						1.557 .060	1.366	1.088	.948 .150	.866 .148	.793 .147	.714	.629 .123	.494 .102	.546	
2.4									.941	.796	.682	.577	.494	.411	.328	
2.8																
$\hat{\mathbf{E}}[X Y_D = y_j]$		.989	2.311	2.473	2.440	2.271	2.153	2.042	1.925	1.866	1.841	1.847	1.899	1.777	.506	
$\hat{\text{Var}}[X Y_D = y_j]$		1.815	1.104	1.069	1.010	.845	.861	. 849	. 811	.818	.797	.762	.765	.606	.424	

Note: Upper values are the means, in feet, and lower values are the variances, in feet squared.

# SUPPLEMENTAL DATA TABLES

Table 60. Estimates of parameters describing two-parameter gamma distribution for conditional step lengths (Run 4A)

$Y_E = y_i$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	2.8
-3.2						2.438 7.407	2.006 5.700	2.551 8.034	36.109 131.689							
-2.8					2.657 11.086	1.780 6.804	1.887 6.832	2.214 7.494	3.172 10.275	3.822 12.258	3.395 10.848	8.236 24.420	7.008 19.397			
-2.4		9.748 35.444	24.356 78.305	1.195 5.654	2.683 10.795	1.950 7.029	1.966 6.834	2.024 6.901	2.549 8.817	2.904 9.749	3.425 11.335	4.285 13.644	3.248 10.389			
-2.0		5.451 17.978	3.929 15.120	2.991 12.603	3.825 14.422	3.007 10.245	3.036 9.901	2.931 9.182	3.557 11.035	3.877 11.772	4.414 12.954	5.222 15.193	4.243 12.730			
-1.6		5.617 15.998	2.643 7.780	2.052 7.012	2.756 8.929	2.698 8.561	3.104 9.548	3.069 9.173	3.795 11.217	4.059 11.762	4.681 13.126	4.996 13.855	5.045 14.226			
-1.2		3.530 9.545	2.822 8.039	2.683 8.096	2.179 6.834	2.204 6.750	2.460 7.319	2.728 7.787	3.220 9.042	3.405 9.623	3.742 10.168	3.974 10.768	4.085 11.328	3.009 9.555		
-0.8	48.352 126.247	6.321 15.664	4.174 10.803	2.916 8.011	2.449 6.822	2.317 6.147	2.430 6.276	2.824 7.074	3.291 8.221	3.671 9.202	3.958 9.761	4.151 10.390	3.830 9.811			
-0.4	21.245 44.232	2.911 7.086	3.148 7.889	2.070 5.358	2.050 5.163	2.051 5.035	2.334 5.451	2.357 5.138	2.811 6.100	3.349 7.358	3.663 8.076	4.041 9.193	3.614 8.529			
0.0		4.199 10.736	3.938 9.058	2.374 5.697	2.127 4.837	2.049 4.551	2.193 4.567	2.259 4.357	2.210 3.972	2.767 5.100	3.298 6.319	3.701 7.424	3.854 8.020			
0.4		2.452 5.668	5.078 13.066	2.401 5.678	2.190 4.727	2.257 4.636	2.468 4.697	2.390 4.130	2.446 3.902	2.256 3.242	3.073 4.674	3.818 6.342	4.135 7.336			
0.8			2.391 5.101	2.098 4.734	1.932 4.073	2.309 4.464	2.572 4.566	2.530 4.058	2.616 3.765	2.693 3.517	2.709 3.053	3.390 4.217	3.905 5.443			
1.2				3.018 6.879	1.941 4.248	2.171 4.190	2.805 4.926	2.600 4.111	2.741 3.763	2.804 3.373	3.310 3.429	2.890 2.515	3.259 3.155			
1.6				6.299 9.562	1.051 1.869	1.791 3.021	2.891 4.503	2.327 3.497	3.038 3.913	2.877 3.285	3.414 3.182	2.953 2.398	2.204 1.311			
2.0								16.369 22.507	9.050 9.828	7.632 7.281	6.789 5.900	5.947 4.668	4.673 3.211	2.266 1.042		
2.4																
2.8																

Note: Upper values are k<sub>1,y,y</sub>, and lower values are r<sub>1,y,y</sub>, .

# STOCHASTIC ANALYSIS OF PARTICLE MOVEMENT OVER A DUNE BED

Table 61. Estimates of parameters describing two-parameter gamma distribution for conditional step lengths (Run 16)

	Γ																	
$Y_E = y_i$	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	2.8	3.2
-3.6																		
-3.2																		
-2.8					5.681 21.885	4.261 19.393		4.616 19.817		4.639 18.683	3.661 14.032	3.760 13.897						
-2.4			135.437 586.986		4.295 19.532		4.291 19.827			4.853 20.464								
-2.0			62.463 216.406	8.215 36,777	5.714 26.418	6.087 28.158	6.458 29.150	5.900 26.179	4.866 21.406	5.661 23.906	5.608 22.900	5.490 21.582						
-1.6			2.394	3.328 15,462	4.132	4.384	4.693	4.555	3.993	4.276	4.802	4.853	5.439	6.211	3.281	5.163		
-1.2			3.142	4.194 18.928	4.138	3.956	4.284	4.400	4.031	4.299	4.938	5.052	5.435	5.547	3.322	4.694	3.587	
-0.8			3.690	4.803 19.308	4.034	3.601	3.294	3.302	3.284	3.359	4.029	4.198	4.651	5.716	4.189	7.061	3.559	
-0.4		1.207	2.925	2.939	3.256	3.221	2.938	2.831	2.722	2.772	3.368	3.828	4.683	5.946	5.154	6.478	3.681 9.931	
0.0		5.300	3.532	2.817	3.070	2.849	9.679	9.026 2.675	8.382 2.492	8.140 2.370	2.862	3.441	4.509	5.788	4.421	5.293	4.024	
0.4		6.695 2.048	11.954 3.364	8.665 3.918	9.617 3.757	8.460 3.613	7.907 3.435	7. 305 3. 206	6.523 3.063	5.797 2.940	6.973 2.563	2.904	4.213	5.504	4.610	5.710	9.896 5.125	
0.8		6.704 1.671	9.732 3.265	4.022	3.815	9.727 3.892	8.790 3.509	7.814 3.410	7.095 3.414	6.336 3.308	5.030 3.118	5.691 2.482	9.159 3.622	12.449 4.970		13.824 5.911	5.932	
		4.752	7.879 6.718	9.320	8.716 4.970	8.951 4.279	7.744 3.730	7.175 3.616	6.784 3.660	6.107 3.682	5.357 3.508	3.770 3.385	6.099	9.065	8.183 3.566	12.507	11.858	
1.2	l		9.614	7.109	10.570	8.805 5.723	7.334	6.667 4.388	6.270 4.330	5.884	5.255 3.837	4.699 3.574	3.319	5.284		10.247 5.706		
1.6					25.145	10.188	7.630	7.144	6.413	5.726	4.977	4.342	3.650 5.164	2.336	3.662 3.478	8.206	9.043	
2.0						20.498	9.938	12.266	9.384	8.679	7.740	6.882	6.186	5.189	3.339	6.070	11.367	
2.4								19.639 27.769						9.309	9.548 7.657			
2.8																		
3.2																		

Note: Upper values are  $k_{1,y,y}$ , and lower values are  $r_{1,y,y}$ .

## SUPPLEMENTAL DATA TABLES

Table 62. Estimates of parameters describing two-parameter gamma distribution for conditional step lengths (Run 17)

$Y_E = y_i$	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0.0	0.4	0.8	1.2	1.6	2.0	2.4	2.8
-3.2																
-2.8								10.359 36.804								
-2.4						3.809 15.335	8.836 33.338	8.241 27.640	7.700 24.594	6.331 20.203	5.013 15.280	6.530 20.934	9.808 35.113			
-2.0				10.425 42.931		4.746 16.332	5.772 19.158	5.512 18.106	5.685 17.646	5.746 17.530	5.535 16.605	6.645 20.134	8.953 28.855	11.706 33.573		
-1.6	l .			11.541 42.620	4.473 16.488	4.622 16.042	5.329 17.692	5.681 18.237	5.634 17.329	6.434 19.661	7.086 21.136		14.322 42.666			
-1.2			2.551 7.820	4.813 16.214	5.392 18.377	5.101 15.901	4.782 14.385	4.961 14.496	5.105 14.386	5.904 16.624	6.113 16.926	6.579 17.790	6.507 17.615			
-0.8			2.522 6.699	3.222 9.709	4.116 12.418	4.373 12.201	3.909 10.406	4.220 10.861	4.715 11.783	5.478 13.563	6.042 14.893	7.137 17.422	7.306 17.615	5.777 13.051		
-0.4			3.020 7.599	4.175 11.572	4.419 11.972	4.221 10.388	3.821 8.996	3.671 8.194	4.047 8.665	4.784 10.186	5.616 12.080	6.178 13.511	5.452 12.098	6.962 13.909		
0.0			12.906 37.143	4.481 11.306	4.557 10.842	4.341 9.346	3.955 8.073	3.705 7.165	3.570 6.397	4.057 7.161	5.168 9.375	6.016 11.292	5.425 10.535			
0.4		2.186 4.614	6.195 13.201	4.646 9.862	5.651 11.817	5.267 9.986	4.430 7.809	4.079 6.739	3.961 5.993	3.417 4.705	4.468 6.350	5.850 8.898	5.642 9.169	8.436 12.882		
0.8			3.727 6.097	7.933 14.977	7.789 14.379	6.922 11.740	5.443 8.562	5.247 7.682	5.183 6.904	4.762 5.805	3.930 4.201	4.710 5.345	5.227 6.748			
1.2				17.600 27.878	7.368 11.944	6.032 9.204	5.714 8.000	5.161 6.606	5.355 6.137	4.990 5.230	4.635 4.362	3.708 2.914	4.411 4.327			
1.6				12.123 15.578	6.129 9.316	6.063 9.337	8.316 10.719	6.503 7.738	5.966 6.192	5.625 5.316	5.224 4.393	4.618 3.348	3.385 1.845			
2.0								10.074 10.960	6.320 5.991	5.851 5.067	5.394 4.278	5.137 3.668	5.114 3.216	4.843 2.392		
2.4																
2.8																

Note: Upper values are  $k_{1,y,y}$ , and lower values are  $r_{1,y,y}$ , .

Table 63. Results of goodness of fit test for conditional step lengths (Run 4A)

		$Y_E = y_i =$	-0.8			Y <sub>E</sub> =y <sub>i</sub> =	0.0			Y <sub>E</sub> =y <sub>i</sub> =	0.8	-
$Y_{D}=y_{j}$	k <sub>1,y,y</sub> , <sup>1</sup> / ft <sup>-1</sup>	r <sub>1,y,y</sub> , <sup>2/</sup>	$m_{i,j}$	Goodness of Fit Test 3/	k <sub>1,y,y</sub> , ft <sup>-1</sup>	r <sub>1,y,y</sub> ,	<sup>m</sup> i,j	Goodness of Fit Test	k <sub>1,y,y</sub> , ft <sup>-1</sup>	r <sub>1,y,y</sub> ,	m <sub>i,j</sub>	Goodness of Fit Test
-1.6	2.449	6.822	128	$x^2 < x_c^2$	2.127	4.837	146	$x^2 < x_c^2$	1.932	4.073	60	$x^2 < x_c^2$
-1.2	2.317	5.147	217	$x^2 < x_c^2$	2.049	4.551	263	$x^2 < x_c^2$	2.309	4.464	124	$x^2 < x_c^2$
-0.8	2.430	6.276	336	$x^2 < x_c^2$	2.193	4.567	440	$x^2 < x_c^2$	2.257	4.566	213	$x^2 < x_c^2$
-0.4	2.824	7.074	423	$x^2 < x_c^2$	2.259	4.357	603	$x^2 < x_c^2$	2.530	4.048	306	$x^2 < x_c^2$
0.0	3.291	8.221	457	$x^2 < x_c^2$	2.210	3.972	725	$x^2 < x_{\mathcal{C}}^2$	2.616	3.765	403	$x^2 < x_c^2$
0.4	3.671	9.202	420	$x^2 < x_c^2$	2.767	5.100	666	$x^2 < x_c^2$	2.693	3.517	455	$x^2 < x_c^2$
0.8	3.958	9.761	307	$x^2 < x_C^2$	3.298	6.319	468	$x^2 < x_c^2$	2.709	3.053	470	$x^2 < x_c^2$
1.2	4.151	10.390	181	$x^2 < x_c^2$	3.701	7.424	253	$x^2 < x_c^2$	3.390	4.217	284	$x^2 < x_c^2$
1.6	3.830	9.811	48	$x^2 < x_c^2$	3.854	8.020	56	$x^2 < x_{\mathcal{O}}^2$	3.905	5.443	67	$x^2 < x_c^2$
	,y,y' = \frac{\hat{E}[}{\frac{Var}{Var}}	$X Y_E=y, Y_D=$ $\{X Y_E=y, Y_D=y\}$	=y'] _=y']	2/ r	$y' = \frac{\left(\hat{E}\right[}{\widehat{Var}}$	$X \mid Y_E = y$ , $Y$ $[X \mid Y_E = y$ ,	<sub>D</sub> =y']) Y <sub>D</sub> =y']	2 <u>3</u> /	$x_c^2 = \text{cri}$ at 0.0	tical chi a signifi 5	-squar cant 1	e value evel of

Table 64. Results of goodness of fit test for conditional step lengths (Run 16)

		$Y_E = y_i =$	-0.8			Y <sub>E</sub> =y <sub>i</sub> =	0.0			Y <sub>E</sub> =y <sub>i</sub> =	0.8	
$Y_{D}^{=y}j$	N	r <sub>1,y,y</sub> , <sup>2/</sup>	$m_{i,j}$	Goodness of Fit Test	k <sub>1,y,y</sub> , ft <sup>-1</sup>	r <sub>1,y,y</sub> ,	$m_{i,j}$	Goodness of Fit Test	k <sub>1,y,y</sub> , ft <sup>-1</sup>	r <sub>1,y,y</sub> ,	$^{m}_{i,j}$	Goodness of Fit Test
-1.6	3.601	13.964	155	$x^2 < x_c^2$	2.849	8.460	184	$x^2 < x_c^2$	3.892	8.951	124	$x^2 < x_c^2$
-1.2	3.294	12.275	239	$x^2 < x_c^2$	2.773	7.907	294	$x^2 < x_c^2$	3.509	7.744	200	$x^2 < x_c^2$
-0.8	3.302	11.992	293	$x^{2} < x_{c}^{2}$	2.675	7.305	376	$x^2 < x_c^2$	3.410	7.175	254	$x^2 < x_c^2$
-0.4	3.284	11.667	341	$x^2 < x_c^2$	2.492	6.523	442	$x^{2} < x_{o}^{2}$	3.414	6.784	297	$x^2 < x_c^2$
0.0	3.359	11.416	365	$x^2 < x_c^2$	2.370	5.797	482	$x^{2} < x_{o}^{2}$	3.308	6.107	317	$x^2 < x_c^2$
0.4	4.029	13.521	337	$x^2 < x_c^2$	2.363	6.973	444	$x^{2} < x_{c}^{2}$	3.118	5.357	325	$x^{2} < x_{c}^{2}$
0.8	4.198	13.799	272	$x^2 < x_c^2$	3.441	8.455	350	$x^2 < x_c^2$	2.482	3.770	331	$x^2 < x_c^2$
1.2	4.651	16.156	146	$x^2 < x_c^2$	4.509	11.896	197	$x^2 < x_c^2$	3.622	6.099	207	$x^2 < x_{\mathcal{O}}^2$
1.6	5.716	19.897	52	$x^{2} < x_{c}^{2}$	5.788	15.477	77	$x^2 < x_c^2$	4.970	9.065	83	$x^2 < x_o^2$
1/ k <sub>1</sub>		$\frac{ X Y_E=y, Y_D}{ X Y_E=y, Y_D}$		2/ r <sub>1,y</sub>	$y' = \frac{\left(\widehat{E}\right)}{\widehat{Var}}$	$\frac{ X Y_E=y, Y}{ X Y_E=y,}$	' <sub>D</sub> =y']) Y <sub>D</sub> =y']	2 <u>3/</u>		tical chi a signifi 5		

Table 65. Results of goodness of fit test for conditional step lengths (Rum 17)

		Y <sub>E</sub> =y <sub>i</sub> =	-0.8			Y <sub>E</sub> =y <sub>i</sub> =	0.0			$Y_{E}=y_{i}$	0.8	
$Y_D = y_j$	k <sub>1,y,y</sub> , 1/ ft <sup>-1</sup>	r <sub>1,y,y</sub> ,2/	<sup>m</sup> i,j	Goodness of Fit Test 3/	k <sub>1,y,y</sub> , ft <sup>-1</sup>	r <sub>1,y,y</sub> ,	<sup>m</sup> i,j	Goodness of Fit Test	k <sub>1,y,y</sub> , ft <sup>-1</sup>	r <sub>1,y,y</sub> ,	<sup>т</sup> і,j	Goodness of Fit Test
-1.6	4.116	12.418	109	$x^2 < x_c^2$	4.557	10.842	129	$x^2 > x_c^2$	7.789	14.379	68	$x^2 < x_c^2$
-1.2	4.373	12.201	259	$x^2 > x_c^2$	4.341	9.346	301	$x^2 < x_c^2$	6.922	11.740	159	$x^2 < x_c^2$
-0.8	3.909	10.406	406	$x^{2} > x_{c}^{2}$	3.955	8.073	493	$x^{2} < x_{c}^{2}$	5.443	8.562	267	$x^2 < x_c^2$
-0.4	4.220	10.861	487	$x^{2} < x_{c}^{2}$	3.705	7.165	632	$x^2 < x_c^2$	5.247	7.682	<b>3</b> 5 7	$x^2 < x_o^2$
0.0	4.715	11.783	502	$x^{2} < x_{c}^{2}$	3.570	6.397	704	$x^2 > x_c^2$	5.183	6.904	411	$x^{2} < x_{c}^{2}$
0.4	5.478	13.563	449	$x^2 < x_c^2$	4.057	7.161	647	$x^2 > x_c^2$	4.762	5.805	425	$x^2 < x_c^2$
0.8	6.042	14.893	324	$x^{2} < x_{c}^{2}$	5.168	9.375	469	$x^2 < x_c^2$	3.930	4.201	427	$x^2 < x_c^2$
1.2	7.137	17.422	155	$x^2 < x_c^2$	6.016	11.292	235	$x^{2} < x_{c}^{2}$	4.710	5.345	247	$x^2 < x_c^2$
1.6	7,306	17.615	53	$x^2 < x_c^2$	5.425	10.535	80	$x^2 < x_C^2$	5.227	6.748	85	$x^2 < x_c^2$

 $<sup>\</sup>frac{1}{k_{1,y,y'}} = \frac{\hat{\mathbf{E}}[\mathbf{x}|\mathbf{x}_{\underline{E}}=y, \mathbf{y}_{\underline{D}}=y']}{\hat{\mathbf{Var}}[\mathbf{x}|\mathbf{y}_{\underline{E}}=y, \mathbf{y}_{\underline{D}}=y']} \qquad \frac{2}{r_{1,y,y'}} = \frac{\left(\hat{\mathbf{E}}[\mathbf{x}|\mathbf{x}_{\underline{E}}=y, \mathbf{y}_{\underline{D}}=y']\right)^{2}}{\hat{\mathbf{var}}[\mathbf{x}|\mathbf{y}_{\underline{E}}=y, \mathbf{y}_{\underline{D}}=y']} \qquad \frac{3}{x_{\mathcal{D}}^{2}} = \text{critical chi-square value at a significant level of 0.05}$ 

ζ m;,j

Sample conditional probability mass function of step lengths,  $p_X|_{Y_D}(x_{\rm B}|y_{\rm J})$  (Run 16)

Table 67.

0.000 0.000

		6.5	0.000	. 025	.012	.016		.014	.013	.011	.010	. 005		.003	.002	. 002	000.	000.		000.	000.	000.	000.
		5.5	0.000	2,6	.080	. 059		.060	.046	.039	.037	.031		.025	.018	.017	.010	910.		.015	000.	000.	000.
	xg, ft	4.5	0.000	.154	.142	.145		.139	.130	.125	.112	.097		.087	.082	660.	.081	360.		.107	.046	000.	000.
	Class Mark, $x_{\mathrm{f g}}$ , ft	3.5	0.292	.323	. 256	. 268		.258	.242	.237	. 228	. 206		. 203	.195	.218	. 259	. 206		.204	.224	000.	000.
	5	2.5	0.583	.260	. 332	. 309		.316	.316	. 299	. 298	.292		.292	.298	. 305	.325	.367		.405	.372	.193	000.
		1.5	0.125	.153	.154	161.		.200	. 227	.250	.257	.284		.287	. 290	.263	. 245	.239		.237	.347	.735	000.
		0.5	0.000	.034	.020	600.		.012	.024	.038	.056	.084		.103	.115	960.	080.	.077		.032	.011	.072	000.
	V = 1	£ 0.	-3.6	-3.2	-2.4	-2.0		-1.6	-1.2	-0.8	-0.4	0.0		4.0	8.0	1.2	1.6	2.0		2.4	2.8	3.2	3.6
(6,12, 44)	(At IIIA)	1	į, i, j	49	174	387	796	1,413		2,306	3,094	3,653	3,653	2,920		1,811	519	43	0		<b>20,818</b>		
( :	$(x_1, y_2, y_3)$ (with the $(x_1, y_2, y_3)$ )		7.5	0.000	000	000	.002	.00		.001	.001	000	000	000		000	000	000.	-		∑ 2 m	1.67 1.63	
n catego	KA COURTON		6.5	0.000	000	.004	.005	.002		.001	.001	.001	.001	000.		000	000.	000.					
1 actor 20	dans for		5.5	0.000	000	.025	.025	.016		800.	.005	.004	.002	.002		.002	.003	.003	:				
man franchion of atom	onnoin s	Mark, $x_{eta}$ , ft	4.5	0.040	.027	.071	.054	.042		.037	.031	.021	.020	.015		.013	.022	.149	:				
		Class Mark	3.5	0.151	.155	.122	.153	.150		.126	. 105	. 092	.085	.080		980.	.083	.160					
Sommia anditional probabilities	and amo		2.5	0.358	.364	.372	.300	. 282		.275	. 264	. 238	.237	.243		. 250	. 284	.137	-				
mia annut	de contra		1.5	0.451	.450	.388	.403	.420		.440	.439	.435	.413	.403		.407	. 395	.374	-				
99			0.5	0.000	.004	.018	. 058	.087		.112	.154	. 209	. 242	.257		. 242	.213	.177	:				
Tohle	100	<u>, , , , , , , , , , , , , , , , , , , </u>	, p_a,	-2.8	-2.4	-2.0	-1.6	-1.2		-0.8	-0.4	0.0	0.4	8.0		1.2	1.6	2.0	2.4				

Note: The class widths for step lengths are equal to 1 foot for all 8.

Variation of conditional mean and variance of grep lengths with elevation of deposition;  $\mathbb{E}[X|Y_D^*y]$  and  $\mathrm{Var}[X|Y_D^*y]$ 

.69 Table

 $\sum_{i} \sum_{j} m_{i,j} = 19,607$ 

Note: The class widths for step lengths are equal to 1 foot for all 8.

Table 68. Sample conditional probability mass function of step lengths,  $p_{K|Y_n}(x_{\boldsymbol{\beta}}|y_j)$  (Run 17)

		Class Mark, $x_{\mathrm{g}}$ , ft	zβ, ft			£
0.5	1.5	2.5	3.5	4.5	5.5	
:			:	1	1	0
8	0.452	0.179	0,368	0.000	000.0	*
30	.356	. 333	.231	.050	000	43
22	.316	.349	.229	.080	.004	223
910.	.317	. 372	.223	690.	.003	697
41	.389	.352	.173	.042	.003	1.646
78	. 402	. 329	.156	.033	.002	2,696
13	.411	.310	.141	.024	.001	3,336
64	.406	. 297	.115	.017	.001	3,638
.196	. 390	. 292	.107	.015	000.	3,440
. 201	.382	. 304	.102	.011	000	2.717
87	. 384	. 327	.092	.010	000	1,516
27	. 399	.323	.107	.010	000	585
41	.497	.304	.058	.000	000	92
33	.667	000.	000.	.000	000.	9
i	;	:	:	1	;	0
					Ş Ş mz.	$\sum_{i} \sum_{j} m_{i,j} = 20,639$

Note: The class widths for step lengths are equal to 1 foot for all  $\beta$ .

$Var[X Y_D=y_j]$ , ft <sup>2</sup>	Run 4A Run 16 Run 17	1.463	96 3.476	32 2.290 1.815	.884 1.635 1.104	76 1.530 1.069	15 1.512 1.010	77 1.499 .845	41 1.509 .861	84 1.547 .849	35 1.477 .811	08 1.432 .818	.935 1.381 .797	.892 1.392 .762	.915 1.234 .765	71 1.252 .606	1.409 .424	1.164	368	
	Run	-	1.396	1.232	·.	1.376	1.415	1.277	1.141	1.084	1.035	1.008	6.	<u>«</u>	o:	1.671	;	;	!	-
ft	Run 17	-	:	0.989	2.311	2.473	2.440	2.271	2.153	2.042	1.925	1.866	1.841	1.847	1.899	1.777	.506	!	:	:
$\hat{\mathbf{E}}[\mathbf{x} \mathbf{X}_D=\mathbf{y}_j]$ , ft	Run 16	2.073	2.512	3.008	2.924	3.097	3.049	2.927	2.825	2.716	2.567	2.473	2.402	2.517	2.500	2.516	2.550	2.308	1.495	;
Ēŗ	Run 4A	}	1.194	2.172	2.397	2.577	2.486	2.384	2.257	2.119	2.025	1.972	1.938	1.983	2.056	2.458	!	:	-	-
Standardized Elevation	$\theta_j$	-3.6	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	8.0-	-0.4	0.0	0.4	8.0	1.2	1.6	2.0	2.4	2.8	3.2	3.6

Sample joint probability mass function of step lengths and elevation of deposition,  $p_{X,Y_D}(x_B,\nu_J)$  (Run 16)

Table 71.

Table 70. Sample joint probability mass function of step lengths and elevation of deposition,  $p_{X_1,Y_2}(\kappa_B,\nu_J)$  (Run 4A)

		-														
Y = 17				Class Mar	ıss Mark, x <sub>B</sub> , ft								Class Mark, xg, ft	, xg, ft		
f a a_	0.5	1.5	2.5	3.5	4.5	5.5	6.5	7.5	$^{1}D^{-}b_{j}$	0.5	1.5	2.5	3.5	4.5	5.5	ف
-3.2			-						-3.7	0000	0 0008	0.0019	0.0013	0 0005	0.005	:
-2.8		0.0000	0.000	0.0000	0.000	:			-2.8	.0003	.0012	.0020	.0025	.0012	.0003	0.0
-2.4	0.000	.0025	.0020	6000.	.0002			:	-2.4	.0002	.0017	.0036	.0027	.0015	6000	٠.
-2.0	.0005	6600.	.0095	.0031	.0018	9000.0	0.0001	-	-2.0	.0002	.0037	6500.	.0051	.0028	.0011	•
-1.6	.0035	.0244	.0182	.0093	.0033	.0015	.0003	0.0001	-1.6	.0005	.0075	.0119	. 0097	.0052	.0023	٠.
-1.2	6200.	.0381	.0255	.0136	.0038	.0014	.0002	.0001	-1.2	.0014	.0137	.0190	.0146	.0078	.0028	٠.
8.0	.0140	.0551	.0345	.0158	.0046	.0010	.0001	.0001	-0.8	. 0039	.0258	.0309	.0245	.0129	.0040	•
4.0-	.0234	9990.	.0400	.0159	. 0047	8000.	. 0002	.0002	-0.4	8900.	.0314	.0364	.0279	.0137	.0045	٠.
0.0	.0335	8690.	.0382	.0148	.0034	9000.	.0002		0.0	.0112	.0378	.0389	.0274	.0129	.0041	٠.
4.0	.0377	. 0644	.0369	.0133	.0031	. 0003	.0002		4.0	.0203	.0566	.0576	.0401	.0172	.0049	٠
8.0	.0308	.0483	.0291	9600'	.0016	.0002			8.0	.0158	.0398	.0409	.0268	.0113	.0025	٠.
1.2	.0170	.0286	.0176	0900.	6000.	1000.			1.2	.0093	.0256	.0297	.0212	9600.	.0017	•
1.6	.0053	8600.	.0070	.0021	.0005	1000	:	;	1.6	.0037	.0114	.0151	.0120	.0038	.0005	i
2.0	.0015	.0032	.0012	.0014	.0013	0000.		-	2.0	.0013	.0039	0900.	.0035	.0016	.0003	i
2.4	-					!		:	2.4	.0002	.001	6100.	.0010	.0005	.0001	į
$p_X(x_\beta)$	.175	. 421	.260	.106	620.	.007	.001	.001	2.8	-				-		i
	Note:	Note: The close widths for seen 1	lake for the			× 10 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	"		$p_{\chi}(x_{\mathrm{B}})$	.075	. 262	.302	.221	.103	.030	9

ite: The class widths for step lengths are equal to 1 foot for all  $\boldsymbol{\beta}.$ 

Note: The class widths for step lengths are equal to 1 foot for all B.

Table 72. Sample joint probability mass function of step lengths and elevation of deposition,  $p_{X_1,Y_D}(\kappa_b,\nu_g)$  (Run 17)

	6.5		:	!	:	1	:		:			:	:	:	;	:		:	
	5.5		:	:	0.0001	.0001	.0003	.0003	.0002	.0002	:	į			:	-		.001	
t	4.5			0.0003	.0014	.0030	.0037	.0043	.0037	.0026	.0023	.0013	.0007	.0004	-			.024	
Class Mark, xg, ft	3.5	******	0.0006	.0012	.0040	.0097	.0154	.0202	.0215	.0176	.0165	.0120	.0067	.0044	.001	-	*****	.131	
Clas	2.5		0.0003	.0018	.0061	.0162	.0314	.0425	.0473	. 0456	.0450	.0357	. 02 38	.0132	.0057			.315	
	1.5		0,0007	6100.	.0055	.0138	.0347	.0520	.0627	.0623	.0601	.0448	.0280	.0163	.0093	. 0023		.394	
	0.5		0.000	. 0002	.0004	.0007	.0037	.010	.0173	.0252	.0302	.0236	.0136	9900.	.0026	.0012		.135	
	, D_B j	-3.2	-2.8	-2.4	-2.0	-1.6	-1.2	-0.8	-0.4	0:0	4.0	8.0	1.2	1.6	2.0	2.4	2.8	$p_X(x_g)$	

Note: The class widths for step lengths are equal to 1 foot for all 8.

		•	
· ·			